Juvenile salmon growth, movement, and survival in Butte Creek and the Sutter
Bypass- A look at the past and present tagging studies


Jeremy Notch, Flora Cordoleani, Alex McHuron, Clint Garman, Tracy McReynolds



## Study Objectives

## Part 1:

-What is the growth and residence time of spring-run juveniles rearing in Butte Creek and the Sutter Bypass? (CDFW CWT Study, 1996-2004)

## Part 2:

-What are the survival and movement rates of smolts out-migrating from the Sutter Bypass? (NOAA Acoustic Tagging Study, 2015-2017)



Mill, Deer, Butte Creek Spring-Run Escapement 1992-2018


- ~750,000 spring-run juveniles (30-40mm) CWT tagged between 1996-2004 near spawning grounds in late January - early February
- 769 recaptured $\sim 70$ miles downstream in the Sutter Bypass
- Unique ID on CWT allows for analysis of group movement and growth rates


1996 - 2004 CWT Recaptures in Sutter Bypass


1996-2004 CWT Recaptures in Sutter Bypass


Walker


Runners



## Butte Basin

- Largest contiguous wetland habitat in the Sacramento Valley
- Butte Sink managed by USFWS as a wildlife refuge
- Mostly comprised of private land with 32 conservation easements



JSATS - Juvenile Salmon Acoustic Telemetry System


- Minimum fish size $>80 \mathrm{~mm}, 6.0 \mathrm{~g}$
- Unique ID for each tag, pings every 5 seconds for 30 days
- Fish released at 9pm to allow for recovery



## Genetics By Year



Av. length $=105 \mathrm{~mm}$

2016


Av. length $=110 \mathrm{~mm}$

2017


Av. length $=93 \mathrm{~mm}$

## Survival estimation



- Covariates: fish length, fish condition factor, temperature at release, flow at release, travel rate

| Model | \# Parameters | Delta AICc |
| :---: | :---: | :---: |
| Reach + Year | 18 | 0 |
| Reach + Travel Rate <br> (Sutter, Sac, Delta, Bay) | 20 | 0.7 |
| Reach + Flow at Release | 17 | 8.34 |
| Reach + Temp at Release | 17 | 8.67 |



## Movement Rates vs. Survival




## Conclusions

- Juveniles tend to walk (72 days, $83 \%$ ) vs run (11 days, 17\%) through Butte Creek and the Sutter Bypass
- Growth rates averaged $44 \mathrm{~mm}(0.6 \mathrm{~mm} /$ day $)$ for walkers
- Low smolt survival rates through the Sutter Bypass and Delta in recent years
- Smolt survival appears to be correlated with movement speed: faster movement speeds lead to higher survival rates



## Thank you



Flora Cordoleani


- Clint Garman
- Tracy McReynolds

- Paul Ward



## Questions ?



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## NOAA Technical Memorandum NMFS



SEPTEMBER 2017

# LIFE HISTORY AND CURRENT MONITORING INVENTORY OF SAN FRANCISCO ESTUARY STURGEON 

Joseph Heublein, Russ Bellmer, Robert D. Chase, Phaedra Doukakis, Marty Gingras, Douglas Hampton, Joshua A. Israel, Zachary J. Jackson, Rachel C. Johnson, Olaf P. Langness, Sean Luis, Ethan Mora, Mary L. Moser, Larissa Rohrbach, Alicia M. Seesholtz, Ted Sommer, and Jeffrey S. Stuart

## NOAA Technical Memorandum NMFS

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## Life History and Current Monitoring Inventory of San Francisco Estuary Sturgeon

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## Preface

This document describes life history and current monitoring of the two endemic sturgeon species of the San Francisco Estuary watershed: the southern Distinct Population Segment (sDPS) of North American green sturgeon green sturgeon and the Sacramento-San Joaquin River white sturgeon. It serves as background information used in the development of the conceptual models in Heublein et al. (2017). This document as well as Heublein et al. (2017) are fundamental in identifying existing and expanded monitoring needs necessary to track the status, trend, and viability of sDPS green sturgeon identified in National Marine Fisheries Service recovery planning efforts. Finally, this synthesized information is relevant in the development of future quantitative life cycle models for both sturgeon species. A comparative life history and a monitoring inventory are summarized below to relate areas where management actions and monitoring opportunities overlap for the two species and identify factors that may be unique to each species. More detailed life history descriptions and monitoring inventories follow; these are organized by species life stage.

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## Comparative Life History and Monitoring of SFE Sturgeon

The spawning distribution of southern Distinct Population Segment (sDPS) green sturgeon and Sacramento-San Joaquin River or San Francisco Estuary (SFE) white sturgeon do not typically overlap (geographic regions of the SFE watershed are illustrated in Figure 1 of Heublein et al. [2017]). The majority of spawning for both species occurs in the mainstem Sacramento River; however, the downstream extent of sDPS green sturgeon spawning is approximately 80 kilometers (km) upstream of the typical spawning range of white sturgeon (Schaffter 1997; Poytress et al. 2015). Most sturgeon spawning in the SFE watershed occurs in spring, although white sturgeon can spawn earlier in winter, and green sturgeon spawning extends into early summer with periodic late summer and fall spawning (CDFG 2002; W. Poytress, USFWS, 2017a, unpublished data, see "Notes"). Spawning by both species typically occurs in areas that are deep and turbulent, but white sturgeon eggs occur in lower gradient reaches where channelized river habitat and fine substrate are common. There is no long-term monitoring of sturgeon spawning events (i.e., sampling for sturgeon eggs) on the Sacramento River. Recently, however, sturgeon egg sampling has occurred with some consistency in secondary spawning rivers (Feather and San Joaquin rivers).

With many of California's native fish subject to elevated and unsuitable water temperatures green sturgeon spawning in the upper Sacramento River is somewhat unique. Managed reservoir releases from Shasta Lake typically generate relatively low temperatures in the upper reaches of green sturgeon spawning and incubation habitat on the Sacramento River, and maintain suitable temperatures even in extreme drought conditions. In contrast, water temperatures that are suitable for normal egg development do not persist through the entire spawning and incubation period for white sturgeon in the San Joaquin River. Conditions on the Feather River in late spring and summer are also not appropriate for green sturgeon spawning and incubation in most years. This limitation to spawning habitat may become a more prominent issue if water temperatures continue to increase. Notably, spawning habitat in the upper reaches of the Sacramento River may provide sDPS green sturgeon resilience to climate change under a future scenario where most SFE sturgeon spawning habitats exceed optimal temperature ranges.

Coupling river conditions during sturgeon spawning and incubation with habitat requirements for egg and larval survival is a clear management consideration for both species. In laboratory studies, embryos of northern Distinct Population Segment (nDPS) green sturgeon and SFE white sturgeon have similar temperature optima and tolerances (see Fertilization to Hatch sections in this document for specific temperature ranges by species; Wang et al. 1985; Van Eenennaam et al. 2005). In this regard, management targets for spawning and incubation temperature may be similar for both species. In addition, river flows (which also influence water temperature) are heavily manipulated throughout the spawning ranges of both species. The direct effect of river flow magnitude and timing on spawning and incubation for both species remains uncertain. There is an outstanding management need for a complete understanding of the effects of managed and unmanaged flow and temperature on SFE sturgeon spawning and incubation.

General patterns of larval dispersal and habitat occupancy are different for green and white sturgeon. Although there is temporal overlap in the larval distributions of the two species, green sturgeon larvae in the Sacramento River have been collected from early spring through summer and white sturgeon larvae are typically collected from late winter through spring. In the Sacramento River, green sturgeon larvae are consistently collected near spawning reaches (Poytress et al. 2009, 2010, 2011, 2012, 2013). White sturgeon larvae, however, disperse from hatching areas and are collected throughout the freshwater estuary (CDFW, 2016a, unpublished data, see "Notes"). In the San Joaquin River, white sturgeon larvae appear to disperse less broadly than their Sacramento River counterparts (Z. Jackson, USFWS, 2017, unpublished data, see "Notes").

There is no intentional long-term monitoring of larval sturgeon in the SFE but larvae of both sturgeon species are captured incidentally in studies targeting other fish. Larval green sturgeon are captured annually in salmonid monitoring at Red Bluff Diversion Dam (RBDD) with rotary screw traps, and larval white sturgeon are captured episodically in the California Department of Fish and Wildlife (CDFW) " 20 mm Survey," which uses a rigid-opening trawl net constructed of 1,600-micrometer mesh to monitor delta smelt in the SFE. Data from these surveys represent an incomplete picture of larval sturgeon distribution, but some general patterns emerge. White sturgeon larvae may move into the freshwater estuary in spring and, as a result, would have access in some years to freshwater estuarine habitat and seasonally inundated wetlands. In contrast, sDPS green sturgeon larvae remain in riverine habitats until juvenile metamorphosis. This difference in larval distributions between species poses unique considerations for management of environmental conditions in their distinct habitats. Larval green sturgeon habitat in the upper Sacramento River is relatively stable due to flow and temperature management through operation of upstream dams and reservoirs. However, larval white sturgeon potentially utilize lower river and estuarine areas where conditions are more variable.

As described above, green sturgeon likely rear in spawning habitat on the Sacramento River for a few months or more before migrating to the estuary as juveniles (W. Poytress, USFWS, 2017a, unpublished data, see "Notes"), whereas white sturgeon may enter the estuary as larvae or early stage juveniles. Juveniles of both sturgeon species have been found throughout the Delta. There is likely significant overlap in distribution at the juvenile life stage, but green sturgeon appear to utilize brackish and seawater portions of the SFE more readily than white sturgeon (Thomas and Klimley 2015; M. Holm 2016, personal communication, see "Notes"). Catch of small juvenile sturgeon of both species in the estuary is episodic with year-class successes likely occurring during wet years with elevated flow (CDFW, unpublished data, 2015, 2016d, see "Notes"; Fish 2010). In laboratory studies using nDPS green sturgeon and SFE white sturgeon, exogenous larvae and juveniles reared with ample food show rapid growth rates in water temperatures at or above 19 degrees Celsius ( ${ }^{\circ} \mathrm{C}$; Cech et al. 1984; Mayfield and Cech 2004; Allen et al. 2006).

Juvenile green and white sturgeon are rarely caught in estuarine monitoring surveys; in many years, juvenile sDPS green sturgeon are absent from all monitoring. Data from juvenile white sturgeon sampling in the estuary is used to generate a coarse recruitment index and forecast of adult cohort abundance (Fish 2010). Although increasing juvenile collection of both species may be challenging, these potential data are critical to understanding environmental factors influencing juvenile sturgeon recruitment and establishing necessary early life stage abundance measures for sDPS green sturgeon.

Adult white sturgeon are found throughout the year in brackish areas of the SFE, and individuals tagged in the SFE are rarely detected or recaptured in marine areas or non-natal estuaries (Kohlhorst et al. 1991). The SFE appears to be exclusively occupied by sDPS green sturgeon with subadult and adult life stages more common in summer and fall compared to other seasons (Israel et al. 2009; National Marine Fisheries Service [NMFS] 2015). A mixture of nDPS and sDPS subadult and adult green sturgeon also occupy estuaries and nearshore marine areas in summer and early fall in Oregon and Washington (Lindley et al. 2011). Green sturgeon in the SFE may comprise a small proportion of the subadult and adult sDPS based on population analysis of non-natal estuarine aggregations of green sturgeon (NMFS 2015; Schreier et al. 2016). As a result, sDPS green sturgeon have been studied extensively through sampling of these mixed distinct population segment aggregations in non-natal estuaries. Additionally, green sturgeon are encountered with some regularity as bycatch in ocean commercial fisheries, allowing opportunities for commercial fishery-dependent monitoring.

White sturgeon are the more common and abundant sturgeon species in the SFE and support a large sport fishery. Thus, directed sturgeon sampling studies and fishery-dependent monitoring in the SFE watershed are focused predominantly on white sturgeon. While disproportionate catch may suggest habitat segregation between species in the SFE, directed sampling of adult green sturgeon in the SFE has almost always involved catch of both species (M. Holm 2016, personal communication, see "Notes"). In this regard, green and white sturgeon most likely occupy similar areas within the SFE. As described above, juvenile SFE sturgeon are rarely encountered, and larval SFE sturgeon are only collected incidentally in studies directed at other species and are typically too small for non-lethal tissue analyses. Ongoing monitoring of adult sturgeon in the SFE and associated studies (e.g., population modeling, tagging, analysis of genetics and contaminants) presents an opportunity to increase understanding of both species.

Spawning sDPS green sturgeon enter the SFE from the ocean from late winter to spring and ascend the Sacramento River with minimal staging and feeding in the estuary (Heublein et al. 2009). Adult white sturgeon spend the majority of their lives in the SFE and can initiate upstream migration in the late fall through spring (Kohlhorst and Cech 2001). Thus, consumption of contaminated food in the SFE may have a greater effect on white sturgeon vitellogenesis and eggs compared to sDPS green sturgeon. Both green and white sturgeon are susceptible to stranding at the Sutter and Yolo Bypass Weirs during the same hydrologic conditions and timing. The observed spawning habitat partitioning on the Sacramento River most likely indicates different preferences for spawning substrate or other habitat attributes between
species (e.g., channel gradient and velocity). Both species show a variety of post-spawn migration and holding behaviors, but it is more common for post-spawn green sturgeon to hold for a few months or more in freshwater habitats and for white sturgeon to return to the SFE immediately after spawning (Heublein et al. 2009; Klimley et al. 2015).

The sDPS green sturgeon spawning run on the Sacramento River has been estimated annually using dual-frequency identification sonar (DIDSON) survey data since 2010 (Mora 2016). There is no directed monitoring of the white sturgeon spawning run on the Sacramento River beyond angler catch reporting in spawning reaches. Recently, directed study of spawning sturgeon has occurred on the San Joaquin River (white sturgeon) and, to lesser extent, the Feather River (primarily green sturgeon). White and green sturgeon migrating to spawning habitat in the Sacramento and Feather rivers are periodically sampled at the Sutter and Yolo Bypass Weirs, and white sturgeon are collected in general fish community monitoring in the Yolo Bypass. Annual spawning measures (e.g., annual run-size, spawning distribution, verification of successful spawning) are fundamental metrics necessary for management of both sturgeon species. Monitoring of spawning sturgeon should be expanded such that these metrics are uniformly available throughout the spawning ranges of both species.

## Specific Life History Descriptions and Monitoring by Life Stage

Prior to developing the comparative life history and monitoring of SFE sturgeon (above), the following inventory of monitoring (summarized in Tables 1 and 2) and specific life history descriptions by life stage were established.

Table 1. Green sturgeon monitoring inventory.
Four general categories are used to organize the discussion of monitoring: life stage-surveys; collection or survey of fish; tissue analyses-analysis of fish tissue samples to estimate age, sex, reproductive condition, migratory histories, etc.; telemetry-monitoring of fish movement and behavior through implantation of active or passive tags and subsequent tag detection; and population modeling and synthesis-modeling or synthesis of survey, tissue analysis, and/or telemetry data to estimate demographics, population trends, movement patterns, etc. Study leads are provided in parentheses.

| Green Sturgeon | Eggs | Larvae | Juveniles | Subadults/Adults | Spawning Adults |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Life stage surveys | - Sacramento River rotary screw trap collections at RBDD (2016, 2017; USFWS) <br> - intermittent egg mat surveys on the Feather River (DWR) <br> - habitat mapping of past egg collection sites and putative spawning areas on the Sacramento River (USFWS; NMFS) | - Sacramento River rotary screw trap collections at RBDD (USFWS) and GCID (GCID) <br> - intermittent larval D-netting on Sacramento River (USFWS) and Feather River (DWR) | - benthic trawl surveys near RBDD (USFWS) <br> - Sturgeon Fishing Report Cards (location and length; CDFW) <br> - intermittent collection in Bay Study (Delta; CDFW) intermittent salvage at the federal and state Delta pumping facilities (DWR; BOR) | - Sturgeon Study (trammel netting in Bay), Sturgeon Fishing Report Cards, boat logs, creel surveys (CDFW) <br> - bycatch in California halibut trawl fishery (NMFS; CDFW) <br> - intermittent sampling in San Pablo Bay, Columbia River, Grays Harbor, and Willapa Bay (UCD; WDFW; NMFS) | - DIDSON surveys in spawning habitat NMFS; CDFW; DWR) <br> - hook and line sampling in Feather River (DWR) <br> - Sturgeon Fishing Report Cards, creel surveys (CDFW) <br> - Yolo Bypass and Knights Landing fyke traps, stranding at Fremont and Tisdale weirs (DWR; CDFW) |
| Tissue analyses |  | - analysis of larval development (spawn timing; UCD) | - genetic analysis of benthic trawl samples (NMFS) | - fin ray analysis from Sturgeon Study (age; USFWS) <br> - genetic analysis of California halibut fishery bycatch (NMFS) | - reproductive stage of sturgeon stranded at Fremont and Tisdale weirs (CDFW; UCD) |
| Telemetry |  |  | - lower Sacramento River monitoring (primarily gill-netting) for acoustic tag implantation (CDFW; UCD) <br> - acoustic tagging of juveniles captured in benthic trawls near $\qquad$ | - satellite tagging of bycatch in California halibut trawl fishery (NMFS; CDFW); <br> - detection of previously tagged fish in existing acoustic arrays (UCD; NMFS) | - detection of previously tagged adults in acoustic array (NMFS; WDFW; UCD; DWR) <br> - acoustic tagging of sturgeon stranded a Fremont and Tisdale weirs (CDFW; UCD) <br> - acoustic tagging of hook and line capture of adults in Feather River (DWR) |
| Population modeling and synthesis |  | - comparison of RBDD larval timing and abundance and environmental conditions (USFWS) | - hind-cast of recruitment success through estimated age of collected and reported juveniles (CDFW) |  | - modeling of run size using DIDSON and acoustic detection data, expansion of run size estimates to adult abundance estimates with spawning periodicity (NMFS) |

## Notes:

Study leads are provided in parentheses in table (RBDD: Red Bluff Diversion Dam; USFWS: U.S. Fish and Wildlife Service; NMFS: National Marine Fisheries Service; DWR: California Department of Water Resources; GCID: Glenn Colusa Irrigation District; UCD: University of California, Davis; CDFW: California Department of Fish and Wildlife; BOR: Bureau of Reclamation; WDFW: Washington Department of Fish and Wildlife)

Table 2. White sturgeon monitoring inventory.

| White Sturgeon | Eggs | Larvae | Juveniles | Adults | Spawning Adults |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Life stage surveys | - egg mat surveys, habitat mapping (San Joaquin River; USFWS) | - larval D-netting (San Joaquin River; USFWS) <br> - 20mm Survey (CDFW) <br> - salvage at the federal and state Delta pumping facilities (BOR; DWR) | - Bay Study (Delta; CDFW) <br> - salvage at the federal and state Delta pumping facilities (BOR; DWR) | - Sturgeon Study (trammel netting in Bay), Sturgeon Fishing Report Cards, boat logs, creel surveys (CDFW) <br> - salvage at the federal and state Delta pumping facilities (BOR; DWR) | - Sturgeon Study (post-spawn), Sturgeon Fishing Report Cards, creel surveys (CDFW) <br> - Yolo Bypass and Knights Landing fyke traps (DWR;) <br> - stranding at Fremont and Tisdale weirs (DWR; CDFW) <br> - gill and trammel netting in San Joaquin River spawning areas (USFWS) |
| Tissue analyses | - development (spawn timing; UCD) <br> - preliminary genetic analysis of eggs (UCD; Cramer Fish Sciences) | - development analysis of larvae (spawn timing; UCD; BOR) | - juvenile age at length (CDFW) | - fin ray analysis from Sturgeon Study (age and growth), tissue analysis from fishing derbies (age and contaminants), preliminary fin microchemistry analysis (USFWS) | - fin ray, reproductive condition, and genetic analysis of sturgeon captured in San Joaquin River spawning habitat (USFWS; UCD) |
| Telemetry |  |  | - lower Sacramento River monitoring (primarily gill-netting) for acoustic tag implantation (CDFW; UCD) | - detection of previously tagged fish in existing array (DWR; UCD) | - tagging of adults from fykes, bypass stranding, and San Joaquin River spawning habitat and detection of current and previously tagged adults in acoustic array (DWR; USFWS; UCD) |
| Population modeling and synthesis |  |  | - ongoing analysis of YCI and hind-cast relative brood year abundance (CDFW) | - ongoing estimates of harvest and adult abundance with Sturgeon Study and Sturgeon Fishing Report Card data (CDFW) | - analyses of recapture rates of externally tagged sturgeon in spawning habitat (CDFW) |

## Fertilization to Hatch

Green sturgeon eggs. Southern DPS green sturgeon spawn during the spring and summer in the middle and upper mainstem Sacramento River. Spawning has been documented on the Feather River in only two years, 2011 and 2017 (Seesholtz et al. 2015; M. Manuel, PSMFC, 2017, personal communication, see "Notes"). Late summer and fall spawning on the Sacramento River has also been documented in four years by larval collections at RBDD (2016; W. Poytress, USFWS, 2017a, unpublished data, see "Notes") and Glenn-Colusa Irrigation District (GCID) oxbow (1997, 1999, and 2000; CDFG 2002).

A time range for fertilization can be estimated using information on egg and larval development and growth from laboratory studies, observed length or development stage at capture, and river water temperatures (Dettlaff et al. 1993; Poytress et al. 2015). Coarse spatial distribution of fertilized eggs or spawning can be generated by combining estimated time ranges of fertilization with corresponding adult detections in spawning habitat. Poytress et al. (2009, 2010, 2011, 2012, 2013) sampled eggs directly in the upper and middle Sacramento River between 2008 and 2012 (summarized in Poytress et al. 2015). The authors confirmed that spawning occurred at seven locations by egg matt sampling along the Sacramento River between the GCID oxbow (river kilometer [rkm] 332.5 from the confluence of the Sacramento and San Joaquin rivers) and Inks Creek (rkm 426). Eggs were collected at depths from 0.6 to 11.3 meters (m), with water temperatures ranging from 9.8 to $17.1^{\circ} \mathrm{C}$. Based on the observed egg development stages and temperatures, spawning had occurred from April through July (Poytress et al. 2015). Notably, green sturgeon eggs were collected in rotary screw traps at RBDD (rkm 391) for the first time in 2016 (W. Poytress, USFWS, unpublished data, see "Notes").

Sturgeon early life history sampling has been conducted in the Feather River since 2011 with egg mats and intermittent D-net sampling (Seesholtz et al. 2015; A. Seesholtz, DWR, 2017, unpublished data, see "Notes"). Sturgeon have never been collected in rotary screw traps in the Feather River. Spawning on the Feather River has been confirmed in only two locations: below Fish Barrier Dam, which is a barrier to anadromous fish passage, and 13 km downstream at the outlet of the Thermalito Afterbay (Thermalito Outlet). Seesholtz et al. (2015) estimated that spawning at the Thermalito Outlet occurred from June 12 through 19, 2011, when water temperatures were 16 to $17^{\circ} \mathrm{C}$, whereas spawning occurred in substantially cooler temperatures at FishFish Barrier Dam (approximately $11^{\circ} \mathrm{C}$ ) in late May or early June 2017 (M. Manuel, PSMFC, 2017, personal communication, see "Notes").

Direct observation or DIDSON detection of sturgeon suggests that sDPS green sturgeon may spawn in other areas of the Feather River watershed. Anglers reported catch of numerous green sturgeon about 5-km downstream from Fish Barrier Dam in 2006 (A. Seesholtz, DWR, 2017, unpublished data, see "Notes"). Green sturgeon aggregations have been detected with DIDSON surveys much lower in the Feather River near Shanghai Bend ( $69-\mathrm{km}$ downstream of Fish Barrier Dam) in habitats that appear suitable for spawning (e.g., suitable depth, water velocity, substrate; A. Seesholtz, DWR, 2017, unpublished data, see "Notes"). Adult green sturgeon have also been observed in spring and summer at Daguerre Point Dam on the Yuba

River (a large tributary to the Feather River) in 2011, 2016, and 2017 (Bergman et al. 2011; M. Beccio, CDFW, 2017, personal communication, see "Notes").

There are no long-term surveys that sample green sturgeon eggs, and locations of egg collection from short-term surveys are unlikely to represent the entire egg distribution. Adult detections in potential spawning habitat during spring and early summer can be used to infer a broader egg distribution. Adults have been detected in spring and summer in approximately 21 areas with suitable spawning habitat in the Sacramento River from above the GCID to the confluence of the Sacramento River and Cow Creek (rkm 451; Heublein et al. 2009; Klimley et al. 2015). Based on these observed adult distributions, contemporary egg distribution on the Sacramento River can range from the GCID oxbow in the middle Sacramento River to Cow Creek in the upper Sacramento River basin.

Nearly all laboratory-based data on green sturgeon physiology (e.g., growth rates, metabolism, swimming performance) involve the nDPS. With the overlap in estuarine and marine habitat between the population segments, and with similarities of adjacent historical natal freshwater habitats (e.g., watersheds of the Sacramento and upper Klamath rivers), it is assumed that physiology of the two distinct population segments is similar. However, the nDPS is an "untested surrogate" for the laboratory-based physiology of the sDPS because laboratory-based data for the sDPS do not exist for comparison. In areas or life stages where field-based information on habitat optima or preference is unavailable for sDPS, laboratory-based nDPS data should be cautiously applied because it is generated in a controlled setting and involves a latitudinally distinct population segment.

The following early life stage physiology information was collected in laboratory study of nDPS green sturgeon. Eggs sampled from a single nDPS Klamath River female hatched after 144 to 192 hours of incubation at $15.7 \pm 0.02^{\circ} \mathrm{C}$ and were an average of $4.44 \pm 0.15$ millimeter (mm) in diameter (Deng et al. 2002). Van Eenennaam et al. (2005) exposed the fertilized eggs of one nDPS female to six temperature regimes ( $11,14,17,20,23$, and $\left.26^{\circ} \mathrm{C}\right)$. Egg survival to hatch was highest in the $14^{\circ} \mathrm{C}$ and $17^{\circ} \mathrm{C}$ treatments ( $39 \%$ and $36 \%$, respectively), with total mortality at temperatures of $23^{\circ} \mathrm{C}$ and above (Van Eenennaam et al. 2005). Van Eenennaam et al. (2005) also found a decreased hatching rate in the $11^{\circ} \mathrm{C}$ treatment compared to $14^{\circ} \mathrm{C}$, but the lower temperature limit for embryo survival was not determined. Elevated water temperature can cause deformities in embryos. The proportion of hatched embryos with deformities was relatively high at $17^{\circ} \mathrm{C}$ and $20^{\circ} \mathrm{C}(10.3 \%$ and $51.6 \%$, respectively $)$ and low at $11^{\circ} \mathrm{C}$ and $14^{\circ} \mathrm{C}$ treatments ( $3.7 \%$ and $1.2 \%$, respectively; Van Eenennaam et al. 2005). Based on this information, Van Eenennaam et al. (2005) concluded temperatures less than $17.5^{\circ} \mathrm{C}$ are optimal for normal development of embryos (results summarized in Table 3).

Table 3. Green sturgeon temperature tolerances by life stage.
Laboratory studies involving nDPS green sturgeon from Klamath River broodstock (a, b, c, d, dd, f) were used to rate water temperatures for the eggs, larvae, and juveniles. Water temperatures recorded during sDPS green sturgeon egg and larvae collection on the upper Sacramento and Feather rivers (e, g, and h) were used to establish "acceptable temperature" for spawning adults and larvae. Categorization of temperature tolerance is not directly comparable at upper and lower levels in this table because "impaired fitness" may be related to both indirect sources of mortality (e.g., reduced growth rate) and direct sources of mortality (e.g., increased rate of deformities). $\mathrm{a}=$ Mayfield and Cech 2004; $\mathrm{b}=$ Van Eenennaam et al. 2005; $\mathrm{c}=$ Werner et al. 2007; d = Allen et al. 2006; $\mathrm{e}=$ Poytress et al. 2012; $\mathrm{f}=$ Linares-Casenave et al. 2013; $\mathrm{g}=$ Poytress et al. 2015; $\mathrm{h}=$ Seesholtz et al. 2015; and dd = Allen et al. 2006b.

| temperature ${ }^{\circ} \mathrm{C}$ | 8 | 9 | 10 | 11 | 12 | 13 | 14 | 15 | 16 | 17 | 18 | 19 | 20 | 21 | 22 | 23 | 24 | 25 | 26 | 27 | 28 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| temperature ${ }^{\circ} \mathrm{F}$ | 46.4 | 48.2 | 50.0 | 51.8 | 53.6 | 55.4 | 57.2 | 59.0 | 60.8 | 62.6 | 64.4 | 66.2 | 68.0 | 69.8 | 71.6 | 73.4 | 75.2 | 77.0 | 78.8 | 80.6 | 82.4 |
| egg |  |  |  | b | b | b | b | b | b | b | b | b | b | b | b, f | b,f | b,f | b,f | b,f | b | b |
| larvae |  |  |  |  |  |  | e | e | e | c | $f$ | dd, f | dd,f | dd, f | dd,f | dd,f | dd,f | dd, c,f | f | f |  |
| juvenile |  |  |  | a | a | a | a | a | a | a | a | a | a | a | a | a | a,d | a | a | a | a |
| spawning adult |  |  | g | g | g | g | g | g | g , h | g, h |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | optimal temperature |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | acceptable temperature |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | impaired fitness; avoid prolonged exposure; increasing chance of lethal effects |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | likely lethal |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | lethal |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
|  | unknown effect upon survival and fitness |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

The following general discussion of water temperature in sDPS green sturgeon egg habitat involves coarse gauge data reported by the DWR California Data Exchange Center (CDEC; Sacramento River gauges at Keswick Dam - rkm 488, Bend Bridge - rkm 415, RBDD rkm 391, and Wilkins Slough - rkm 190, see hyperlinks for historical and current CDEC gauge data, DWR, 2017a, in "Notes") and more-specific temperature information from unpublished agency monitoring and modeling efforts. In late spring and summer, flow and temperature on the Sacramento River are manipulated in part for agricultural diversion (GCID - rkm 332.5, and RBDD are the largest diversion facilities on the Sacramento River) and maintenance of a $13.3^{\circ} \mathrm{C}$ ( $56^{\circ}$ Fahrenheit) water temperature "compliance point" for Sacramento River winter-run Chinook salmon (SRWC; Oncorhynchus tshawytscha) spawning and incubation. The compliance point typically ranges from rkm 415 to rkm 444 , and water temperatures upstream of rkm 444 are progressively cooler up to Keswick Dam, where Shasta Lake reservoir release temperatures typically range from 10 to $12^{\circ} \mathrm{C}$. As a result, optimal temperatures for sturgeon egg incubation (below $17^{\circ} \mathrm{C}$ ) extend downstream to RBDD in spring and summer of most years, but typically remain below laboratory-based optima for egg incubation upstream of existing spawning and egg incubation areas (closer to Keswick Dam). In downstream areas closer to GCID, water temperatures are typically suboptimal for incubation in the late spring and summer (above $17.5^{\circ} \mathrm{C}$ ). During periods of extremely low reservoir storage and outflow, incubation temperatures may be suboptimal in the late spring and summer near RBDD and approach potentially lethal temperatures (above $20^{\circ} \mathrm{C}$ ) in areas closer to GCID.

Fish Barrier Dam on the Feather River creates a small impoundment below Oroville Reservoir called the Thermalito Diversion Pool. Water from the Thermalito Diversion Pool is either released to the river at Fish Barrier Dam or routed through the shallow Thermalito Afterbay, thereby bypassing a short tailwater and warming for crop irrigation. Warmer water is then diverted directly from the Thermalito Afterbay or released back into the Feather River at the Thermalito Outlet. Water management downstream of the Thermalito Outlet may have a variable effect on green sturgeon spawning habitat depending on the time of the year and operations. Peak winter and spring flows on the Feather River are typically captured in upstream reservoirs. Summer base flows can then be augmented by releases from the Thermalito Outlet, and flow augmentation may also decrease summer water temperatures. Water released from the Thermalito Outlet is warmer than the potential cold water pool released into the Feather River at Fish Barrier Dam. Hence, suitable conditions for normal egg incubation (water temperatures below $17.5^{\circ} \mathrm{C}$ ) typically only persist into May downstream of the Thermalito Outlet.

White sturgeon eggs. The SFE white sturgeon spawn primarily in the mainstem Sacramento River and to a lesser extent in the San Joaquin River in late winter and spring. Based on collection of gravid adults, spawning also periodically occurs in other rivers (e.g., the Feather and Bear rivers) in the Sacramento River basin. There is no long-term monitoring of white sturgeon eggs in the Sacramento River, but egg distribution is likely similar to the putative spawning reach, which typically ranges from just upstream of Colusa (rkm 252) to near Verona (rkm 129; Kohlhorst 1976; Schaffter 1997). Adult white sturgeon have been reported by anglers
near the confluence of the Sacramento River and Deer Creek (rkm 354) and have been observed and collected during summer and early fall near the GCID oxbow (rkm 332.5; Bergman 2011; M. Manuel, PSMFC, 2016, personal communication, see "Notes"). Juvenile white sturgeon have been collected in the rotary screw traps at GCID and reported by anglers in the area (M. Manuel, PSMFC, 2016, personal communication, see "Notes"). Given those observations, spawning white sturgeon and eggs presumably occur upstream from the GCID oxbow in some years. Egg distribution upstream of GCID requires verification by egg or larval collections because juveniles collected in the GCID trap may have moved there from downstream.

Sturgeon egg sampling is underway in the San Joaquin and Feather rivers in years when adult sturgeon are detected in spawning habitats. However, no white sturgeon eggs have been identified in samples from the Feather River. As described above, sturgeon egg distribution can also be inferred from ongoing acoustic detections of tagged adults in spawning habitat, reports of adults in spawning habitat gathered from the Sturgeon Fishing Report Cards, and from incidental bycatch of larvae and early stage juveniles in other monitoring or research studies.

In the Sacramento River, white sturgeon typically spawn between February and June. Tagged adult white sturgeon move from the SFE to spawning reaches of the Sacramento River between mid-February and early June (Miller 1972; Schaffter 1997; E. Miller UCD Biotelemetry Laboratory, 2017, unpublished data, see "Notes"). Kohlhorst (1976) used D-nets to collect white sturgeon eggs and larvae as early as mid-February, with the majority (93\%) collected in March and April. Schaffter (1997) collected eggs in the Sacramento River on artificial substrates from March through May. Angler reporting of tagged adults in spawning habitat coincident with early season flow events also indicates migration and possibly spawning occur in December and January in some years (CDFW, 2016b, unpublished data, see "Notes"). In the San Joaquin River basin, fertilized eggs were collected downstream of Grayson at rkm 138 (measured from the Sacramento River confluence) in late April 2011 and downstream of Vernalis between rkm 115 and 140 from late March through mid-May in 2012 (Jackson et al. 2016). In March and April 2016, fertilized eggs were collected between rkm 115 and 140, though collection of a single larvae approximately 1 day post-hatch (dph) at rkm 101 indicates that spawning occurs further downstream than previously known (Z. Jackson, USFWS, 2017, unpublished data, see "Notes").

Benthic substrate in white sturgeon spawning habitat is variable. Substrate in the Sacramento River ranges from fine sand to coarse sand near Verona and Wilkins Slough (rkm 189.5), fine gravel to medium gravel farther upstream near Colusa, and gravel and cobble in spawning sites just upstream from Colusa where the river gradient is higher. Schaffter (1997) found eggs in water depths of 1.5 to 4.6 m , with water velocities greater than 1.0 meter per second. Eggs were also collected in San Joaquin River spawning areas at depths of 1.6 to 10.5 m (Jackson et al. 2016). The river bottom in egg-collection areas of the San Joaquin River was dominated by silt and sand substrates, though patches of hardpan clay and fine gravels were present near several sites (Jackson et al. 2016). Fine substrate and lack of interstitial space in spawning habitat can decrease survival of white sturgeon eggs (Hildebrand et al. 2016). Thus,
substrate in SFE white sturgeon spawning habitats, especially on the San Joaquin River, may limit recruitment to the larval or juvenile life stage.

Wang et al. (1985) found that development and survival of white sturgeon embryos were temperature-dependent. Optimal survival to hatching was observed when water temperatures during incubation in the hatchery were between $14^{\circ} \mathrm{C}(88.6 \% \pm 2.2 \%$ survival $)$ and $17^{\circ} \mathrm{C}(83.6 \%$ $\pm 1.9 \%$ survival). Embryo mortality increased as water temperatures increased to $20^{\circ} \mathrm{C}(49.1 \% \pm$ $3.2 \%$ survival), and water temperatures of more than $20^{\circ} \mathrm{C}$ were lethal to developing embryos. Water temperatures in spawning habitat on the Sacramento River typically remain below this level, but median daily water temperatures in excess of $20^{\circ} \mathrm{C}$ were recorded in Sacramento River incubation habitat during drought conditions in April 2015 (CDEC Sacramento River gauges at Wilkins Slough, DWR, 2017a, see "Notes"). Spring water temperatures regularly reach or exceed suitable levels for egg incubation in the San Joaquin River, but viable white sturgeon embryos have been collected in water temperatures above $20^{\circ} \mathrm{C}$ (Jackson et al. 2016). The time necessary for white sturgeon eggs to hatch is also temperature-dependent (Wang et al. 1985; Deng et al. 2002). Under an optimal incubation water temperature of $15.7 \pm 0.2^{\circ} \mathrm{C}$, Deng et al. (2002) found that development time to hatching ranged from 152 to 200 hours and averaged 176 hours. Hatch time is an important consideration for managers if there are potentially variable habitat conditions (e.g., inundation levels, temperature) during egg development.

## Hatch through Metamorphosis

Larval green sturgeon. Metamorphosis of nDPS green sturgeon (to juveniles) occurs at approximately 45 dph and an approximate length of 75 mm (Deng et al. 2002). Growth begins to increase as larvae transition from endogenous to exogenous feeding, which occurs at approximately 15 dph (Gisbert et al. 2001). Laboratory studies have shown that nDPS larvae possess limited swimming ability and generally seek refuge at 0 to 18 dph , suggesting that complex habitat with interstitial space for refuge (e.g., large cobble substrate) is critical at this life stage (Kynard et al. 2005). Under laboratory conditions, nDPS individuals 18 to 45 dph demonstrated initiation of diel downstream migration, favoring nighttime migratory movements (Kynard et al. 2005). Based on egg mat and rotary screw trap sampling at RBDD, Poytress et al. (2011) found similar results in field study of sDPS larvae compared to the laboratory investigations of nDPS larvae described above; larval dispersal was initiated at 18 dph , it peaked at 23 to 24 dph , and complete dispersion from hatching areas occurred by 35 dph .

Historically, larval sturgeon collected in the Sacramento and San Joaquin river systems were not routinely identified to species. However, green sturgeon larvae were identified in RBDD rotary screw trap samples when the study was initiated in 1995, which confirmed green sturgeon spawning at or above this location (Poytress et al. 2009). Larvae have been consistently collected in rotary screw traps in spring and summer at RBDD and GCID as part of a long-term monitoring program in the Sacramento River focused on juvenile salmonids. Data from rotary screw traps at RBDD have shown a similar temporal distribution data for larvae in the Upper Sacramento River compared to D-net sampling data (W. Poytress, USFWS, 2015a, unpublished data, see "Notes"). Larval catch numbers at GCID rotary screw traps are typically lower than

RBDD rotary screw traps (CDFG 2002), and appear to be less representative of temporal spawning downstream of RBDD.

Brown (2007) discovered larvae in the upper Sacramento River at Bend Bridge (rkm 415), and spawning has been confirmed as far downstream as the GCID facility (Poytress et al. 2015), suggesting that larvae occupy over 100 km of the Sacramento River. Due to the limited number of larval collections, the upstream egg distribution described above (Sacramento River to Cow Creek and Feather River to Fish Barrier Dam) is also applied to larvae. Larval sturgeon distribution may range well downstream from spawning habitat. Larval white sturgeon and unidentified larval sturgeon have been collected more than 100-km downstream from known white or sDPS green sturgeon spawning areas during the 20 mm Survey and in salvage at the Tracy Fish Collection Facility and John E. Skinner Fish Protection Facility near federal and state Delta pumping facilities (CDFW, 2016a, unpublished data, see "Notes"). Thus, larval distribution is estimated to extend at least $100-\mathrm{km}$ downstream from spawning habitats on the Sacramento and Feather rivers in high flow years. This estimated downstream distribution corresponds with the Colusa area on the Sacramento River (rkm 252) and the confluence of the Sacramento and Feather rivers near Verona (rkm 129) for larvae originating in the Sacramento River and Feather River, respectively.

Larval abundance and distribution may be influenced by spring and summer outflow. There appears to be a positive relationship between annual outflow and larval abundance in the RBDD rotary screw traps; RBDD rotary screw trap collections of larval green sturgeon were far greater in the three most recent wet years ${ }^{1}(2011,2016$, and 2017) since the monitoring began in 1995 (W. Poytress, USFWS, 2015a, unpublished data, see "Notes"). Moreover, green sturgeon eggs and larvae were only collected during wet years (2011 and 2017) on the Feather River (Seesholtz et al. 2015; M. Manuel, PSFMC, 2017, personal communication, see "Notes").

In laboratory-based studies with abundant food and oxygen, larval nDPS green sturgeon ( 35 dph ) growth was significantly greater at $24^{\circ} \mathrm{C}$ than $19^{\circ} \mathrm{C}$ or cycling 19 to $24^{\circ} \mathrm{C}$ (Allen et al. 2006b). Furthermore, larval nDPS green sturgeon reared at 11 and $13^{\circ} \mathrm{C}$ under normal food conditions had significantly reduced growth rates compared to larvae reared at 16 and $19^{\circ} \mathrm{C}$, but temperature had no significant effect on size or growth rate when larvae were food-limited (J. Poletto, 2016, unpublished data, see "Notes"). Green sturgeon eggs, larvae, and juveniles cooccur in the middle and upper Sacramento River due to the protracted spawning period of sDPS green sturgeon. Spawning of critically endangered SRWC also occurs closer to Keswick Dam during sDPS spawning. Laboratory-based optima for larval and juvenile green sturgeon growth exceed suitable temperatures for green sturgeon spawning and embryo development (Table 3), and are well above suitable temperatures for Chinook salmon spawning and incubation. Thus,

[^1]water temperature in the middle and upper Sacramento River may have different effects on specific green sturgeon life stages and SRWC. Potentially optimal water temperatures for larval growth near RBDD could correspond to low concurrent spawning production and egg survival. Furthermore, warmer water temperatures near RBDD could be associated with poor survival of early life stages of green sturgeon in lower reaches of Sacramento River spawning habitat and SRWC closer to Keswick Dam.

Larval white sturgeon. In a laboratory study, metamorphosis of larval white sturgeon to juveniles occurred at approximately 45 dph (Deng et al. 2002). Larval white sturgeon are not specifically monitored, except for pilot efforts in 2013, 2015, and 2016 on the lower San Joaquin River (Faukner and Jackson 2014; Jackson et al. 2016) and a more directed effort begun in 2017 (Z. Jackson, USFWS, 2017, unpublished data, see "Notes"). White sturgeon larvae are also periodically collected in various locations throughout the Delta in general larval fish monitoring (e.g., 20 mm Survey) in late winter and spring and in salvage at federal and state Delta pumping facilities.

Upstream larval distribution in the Sacramento River can be extrapolated from adult monitoring in spawning habitat (e.g., in-river sportfishing catch data, adult telemetry studies). Periodic aggregations of gravid white sturgeon or telemetry data in the Feather River, Yolo Bypass, and Bear River indicate spawning and larval distribution may also extend to those upstream areas. Based on this information, larval distribution ranges from downriver of spawning habitats (primarily in the Sacramento and San Joaquin rivers) to the approximate downstream extent of the Delta at Chipps Island.

Kynard and Parker (2005) report that Sacramento River white sturgeon larvae use benthic habitat following brief post-hatch dispersal and availability of interstitial benthic space may increase larval survival through multiple mechanisms (i.e., reduction in metabolic demand, stress, and predation; Hildebrand et al. 2016). Laboratory studies indicate exogenous larval white sturgeon are tolerant of salinities up to 16 parts per thousand ( ppt ) and have relatively high growth rates in $20^{\circ} \mathrm{C}$ water (Cech et al. 1984; McEnroe and Cech 1985). However, suitable temperatures for white sturgeon larvae outside of controlled laboratory conditions and for early stage larvae may be lower (less than $16^{\circ} \mathrm{C}$; Hildebrand et al. 2016). A positive relationship exists between relative larval abundance and wet years based on data from long-term collection of white sturgeon larvae in the 20 mm Survey (CDFW, 2016a, unpublished data, see "Notes"). Furthermore, larvae were only collected in San Joaquin River spawning habitats in 2016 (relatively cool conditions) and 2017 (during flooding conditions; Z. Jackson, USFWS, 2017, unpublished data, see "Notes").

## Complete Metamorphosis to Ocean Migration or 75 Centimeters Fork Length

Juvenile green sturgeon. It is unknown how long juvenile sDPS green sturgeon remain in upriver rearing habitats after metamorphosis. Juveniles captured in the Delta by Radtke (1966) ranged in size from 200 to 580 mm fork length (FL), suggesting that juveniles remain upriver for several months before entering the Delta. The lack of juveniles smaller than 200 mm FL in Delta capture records further supports extended upriver rearing of sDPS juveniles before entering the
estuary (CDFG 2002). The first successful study of juveniles in the Sacramento River occurred in 2015 in the Red Bluff area, where approximately 40 large age- 0 green sturgeon ( 60 to 320 mm total length [TL]) were captured using benthic trawls throughout the summer (W. Poytress, USFWS, 2015b, unpublished data, see "Notes"). Based on those studies, it is likely that juveniles rear near spawning habitat for a few months or more before migrating to the Delta.

Duration of juvenile estuary rearing prior to ocean entry and subadult transition is also unknown. Small juvenile green sturgeon (age-0 and presumably age-1) have been incidentally captured in the following general fish community monitoring efforts and commercial fisheries during all months of the year: the CDFW San Francisco Bay Study trawl survey (hereafter referred to as the Bay Study), which has occurred throughout the SFE since 1980; salvage at the federal and state Delta pumping facilities; the USFWS Delta Juvenile Fish Monitoring Program; the San Francisco Bay herring gillnet fishery (M. Holm, 2016, personal communication, see "Notes"); and the San Francisco Bay shrimp trawl fishery (juvenile Acipenseridae, not identified to species). Larger juveniles are periodically captured in the following sampling studies targeting sturgeon: CDFW Sturgeon Population Study (hereafter referred to as the Sturgeon Study) trammel netting in San Pablo Bay and Suisun Bay (hyperlink to annual CDFW study summary in "Notes"); Sturgeon Fishing Report Card capture records from throughout the SFE (hyperlink to annual CDFW study summary, 2017, in "Notes"); and the historical CDFW juvenile sturgeon set-line study in the SFE (Dubois et al. 2010).

Thomas and Klimley (2015) reported juvenile sDPS green sturgeon movement in the SFE in a study that involved capturing larval green sturgeon at RBDD, rearing the fish in the laboratory for several months, and releasing those fish as tagged juveniles to the Delta. The tagged juveniles showed a broad range of movement patterns, from remaining in the Delta for several months to entering the ocean within a few months of release (Thomas and Klimley 2015). Results from Thomas and Klimley (2015) suggest that some individuals in the sDPS may enter the ocean and transition to the subadult life stage in their first year.

Juvenile growth in the nDPS is rapid as they move downstream, reaching up to $300-\mathrm{mm}$ TL in the first year and more than $600-\mathrm{mm}$ TL in years 2 and 3 (Nakamoto et al. 1995). Juveniles in the sDPS may grow considerably faster, with some having reached approximately $300-\mathrm{mm}$ TL in 6 months or less (W. Poytress, USFWS, 2015b, 2017b, unpublished data, see "Notes"). Northern DPS juveniles tested under laboratory conditions have optimal bioenergetic performance (growth, metabolic rate, temperature preference, and swimming performance) at 15 to $19^{\circ} \mathrm{C}$ (Mayfield and Cech 2004) and are highly tolerant to changes in salinity during their first 6 months (Allen et al. 2011). Although green sturgeon have the ability to enter seawater well before 1 year of age (Allen et al. 2011), the typical length of fish encountered in the ocean (greater than 600 mm TL ) suggests ocean entry occurs at a later age. However, length distributions of green sturgeon captured in the ocean may be biased high because most of those records are from commercial fisheries targeting relatively large fish species.

A pilot effort is underway by the UCD Biotelemetry Laboratory and CDFW to implant acoustic transmitters in juvenile green and white sturgeon captured in the Sacramento River and

Delta. Results from this study are likely to improve understanding of juvenile sturgeon habitat use in the SFE. California fishing regulations since 2007 have required sturgeon anglers to complete a Sturgeon Fishing Report Card, and juvenile green sturgeon capture date and length have been reported by this program. A preliminary recruitment time series by CDFW-using lengths of juvenile sDPS green sturgeon and nDPS age-length data-suggests recruitment was highest in the two most recent wet years ${ }^{2}$ (M. Gingras, CDFW, 2016, unpublished data, see "Notes"). The recent implementation of the Sturgeon Fishing Report Card and refinement of agelength keys also provide managers with an early qualitative indicator of recruitment or cohort success using reported lengths of juvenile sDPS green sturgeon and records in the Sturgeon Study.

From scant collections of small sDPS green sturgeon in the Bay Study, nearly all recruitment appears to have occurred in wet years (M. Gingras, CDFW, 2016, unpublished data, see "Notes"). Furthermore, at federal and state Delta pumping facilities the highest juvenile green sturgeon collection on record and the highest estimated salvage since 1985 occurred in a wet year (2006; Gartz 2007). Based on lengths of salvaged green sturgeon reported in Gartz (2007), it is likely that the majority of juveniles salvaged in 2006 hatched in the same year. All of those findings are consistent with white sturgeon and the relationship between recruitment to age-0 and wet years (CDFG 1992; Fish 2010).

Juvenile green sturgeon entrainment in the presence of unscreened water diversions (Mussen et al. 2014) and diversions with fish protection devices (Poletto et al. 2014a, 2014b, 2015) have been studied extensively in the laboratory. These studies suggest juvenile green sturgeon are at high risk of entrainment in unscreened diversions and impingement on screened diversions. Furthermore, Vogel (2013) captured four green sturgeon and one white sturgeon during evaluation of fish entrainment at 12 unscreened diversions in the middle Sacramento River. Verhille et al. (2014) reported larval and juvenile green sturgeon swimming performance and flow velocity recommendations for diversions by life stage; however, fish screen design criteria specifically for larval sturgeon have not been developed.

Juvenile white sturgeon. The typical white sturgeon life history does not involve extensive marine migration, and an obvious limit to the juvenile life stage (e.g., first ocean entry in sDPS green sturgeon) is not known for the species. Chapman et al. (1996) estimated 75- and 95 -centimeter (cm) FL as minimum sizes of maturity for male and female white sturgeon, respectively. Based on this estimate, $75-\mathrm{cm}$ FL was selected as the upper size limit of the juvenile life stage.

Juvenile white sturgeon are believed to initiate a secondary dispersal (the primary dispersal occurring at the larval stage) in spring by actively swimming downstream during the night (Kynard and Parker 2005). Dispersal duration is unknown, but observed swimming intensity and duration in laboratory studies indicate dispersal likely lasts several days and over many kilometers (Kynard and Parker 2005).

[^2]Radtke (1966) indicated that juvenile white sturgeon are found in the Sacramento River and Delta, with the majority of juveniles captured in the Sacramento River. Numerous juvenile and larval white sturgeon have been sampled from the lower San Joaquin River, but those fish are presumed to have entered the system from the Sacramento River through the lower Mokelumne River, Georgiana Slough, or Three Mile Slough (Stevens and Miller 1970). However, spawning has been documented in the San Joaquin River, such that small juvenile white sturgeon collected in San Joaquin River spawning areas (upstream of the Delta) are almost certainly of San Joaquin River origin.

Small juvenile white sturgeon are collected sporadically in various tributaries of the SFE, and the following are some notable examples: one juvenile white sturgeon was captured in a rotary screw trap on the lower Mokelumne River over the past 15 years of monitoring (Workman 2003); one juvenile white sturgeon (possibly two) was detected in the Feather River about 2- to 3-km downstream of Thermalito Afterbay Outlet in May 2000 (Schaffter and Kohlhorst 2001); juvenile white sturgeon have been sampled in the Yolo Bypass during fisheries surveys (B. Schreier, DWR, 2015, personal communication, see "Notes"); and one juvenile white sturgeon was captured in a rotary screw trap located in the American River during February 2015 (C. McKibbin, CDFW, 2015, personal communication, see "Notes"). Juvenile white sturgeon are typically salvaged in low numbers at the federal and state Delta pumping facilities, but a relationship between salvage numbers at the state facility and year-class abundance is lacking (Gingras et al. 2013).

Managers have an early indication of cohort strength and can relate general environmental conditions associated with cohort success. The only index of age-0 production comes from general fish community sampling throughout the SFE in the Bay Study and the index is typically influenced heavily by catch in the lower Sacramento River in the Rio Vista area before May. Indices of recruitment, or year-class index (YCI) have been generated by otter trawl catch of age-0 and age-1 white sturgeon in the Bay Study since 1980. Recruitment trends in the YCI are supported by back-calculation of recruitment with direct or assigned estimates of age (e.g., age analysis of pectoral fin rays or use of age-length keys) from collections of white sturgeon by various monitoring studies. CDFW and USFWS continue to develop those recruitment-modeling techniques by refining estimates of age at length, gear selectivity, and mortality rates.

A positive relationship between high outflow and white sturgeon recruitment in the SFE is supported by juvenile surveys (CDFG 1992; Fish 2010) and hind-cast estimates of relative brood-year abundance from adult monitoring studies (Shirley 1987; M. Gingras, CDFW, 2015a, unpublished data, see "Notes"). Recent investigations show a Delta outflow-recruitment threshold at about 1,416 cubic meters per second $\left(\mathrm{m}^{3} / \mathrm{s} ; 50,000\right.$ cubic feet per second), such that the juvenile YCI is generally strong when flows are above that level (Fish 2010). Cohort abundance information since 1938 shows a boom-and-bust pattern that appears to be related to those high outflow periods (Shirley 1987; CDFG 1992; Fish 2010). Large white sturgeon cohorts have only been detected twice in the last 20 years-in 1998 and 2006-and both years were
classified as wet (CDFW, 2016b, unpublished data, see "Notes"). It should be noted, however, that the primary white sturgeon recruitment survey (Bay Study) was incomplete in 2016 and 2017; indexes have not been generated since 2015.

Small juvenile white sturgeon are likely preyed upon by a variety of native and invasive piscivores (Hildebrand et al. 2016). A mark-recapture study involving small juvenile hatchery white sturgeon (approximately 200 mm FL) was recently conducted to estimate the efficiency of louvers intended to prevent entrainment of fish at the Tracy Fish Collection Facility (Karp and Bridges 2015). Karp and Bridges (2015) found white sturgeon in the stomachs of striped bass collected at the facility during this study. The application of an underwater noise device increased recapture of white sturgeon and louver efficiency estimates from $32.2 \%$ to $74.0 \%$, and the increases were attributed to a disruption in predation (Karp and Bridges 2015).

Poor water quality and consumption of contaminated prey species in the SFE watershed may affect juvenile white sturgeon growth and survival. For example, juvenile white sturgeon grew quickly in laboratory studies with ample food at 20 and $25^{\circ} \mathrm{C}$ but growth was negatively affected by reductions in dissolved oxygen (DO) at all temperature treatments ( 15,20 , and $25^{\circ} \mathrm{C}$; Cech et al. 1984). In laboratory studies, juvenile white sturgeon were able to tolerate abrupt transfer from freshwater ( 0 -ppt salinity) to 15 -ppt-salinity water for up to 5 days but experienced high mortality rates in abrupt transfers from freshwater to $25-\mathrm{ppt}$ and 35 -ppt-salinity water (McEnroe and Cech 1985; Amiri et al. 2009). Methyl mercury and selenomethionine (two ubiquitous contaminants in the SFE) added to food of juvenile white sturgeon reduced growth and increased mortality (Lee et al. 2011; De Riu et al. 2014). Furthermore, juvenile white sturgeon mortality increased in salinities of 15 to 20 ppt when they were previously exposed to dietary selenium (Tashjian et al. 2007).

Size rather than age appears to provide an osmoregulatory advantage (Amiri et al. 2009), such that larger juvenile white sturgeon are more common in brackish estuarine areas. White sturgeon sampled in CDFW setline surveys in Suisun and San Pablo bays had an average length of $86-\mathrm{cm}$ TL (DuBois et al. 2010). Larger juvenile white sturgeon are also presumed to become more piscivorous, and are known to consume herring (Clupea harengus pallasi) and their eggs, American shad (Alosa sapidissima), starry flounder (Platichthys stellatus), and unidentified gobies (Radtke 1966; McKechnie and Fenner 1971).

## Ocean Migration and Maturity

Subadult and adult green sturgeon. The subadult life stage begins at the first entry to the Pacific Ocean and extends to sexual maturation. Adults are mature fish that are not engaged in spawning or spawning migratory behavior. Subadults and adults inhabit marine and estuarine waters along the west coast of North America from Baja California (Rosales-Casian and Almeda-Juaregui 2009) to the Bering Sea (Colway and Stevenson 2007). When not in rivers for spawning, adults and subadults migrate seasonally along the coast and congregate at specific sites in nearshore marine waters (Lindley et al. 2008; Lindley et al. 2011). Subadults and adults enter estuaries and bays during early spring to summer months (presumably for feeding; Dumbauld et al. 2008) and return to the ocean during summer and fall (Moser and Lindley 2007;

Lindley et al. 2011). Southern DPS green sturgeon concurrently occupy several estuaries along the west coast during summer and fall months, primarily San Francisco Bay, Willapa Bay, Grays Harbor, and the Columbia River Estuary. The lack of angler reports or acoustic detection of tagged subadult-sized fish (roughly $60-$ to $100-\mathrm{cm} \mathrm{TL}$ ) upstream of the Delta suggest subadults do not use freshwater riverine habitats.

Only a small number (typically, less than 100) of sDPS subadult and adult green sturgeon are collected annually in all monitoring studies in California, and consistent capture occurs only in the long-term (more than 40 years) Sturgeon Study. Data from the Sturgeon Study have been occasionally used to estimate sDPS abundance, but estimates have been imprecise because surveys and associated analyses are primarily designed to study white sturgeon, and because the mark and recapture rates of sDPS green sturgeon has typically been low. Catch of subadult and adult sturgeon is reported by recreational anglers in the SFE and tributaries through the Sturgeon Fishing Report Card program. Some reporting also occurs in the lower Columbia River and bays of Oregon and Washington. Creel surveys of California recreational fisheries and commercial passenger fishing vessels report sturgeon catch; however, prior to 2011, that catch was rarely identified to species. Thus, it is difficult to generate fishery-dependent sDPS green sturgeon harvest and abundance estimates in the SFE from those data beyond applying a rough assumption of the percentage of sDPS green sturgeon in the overall sturgeon population. A major challenge in estimating abundance and overall harvest of sDPS green sturgeon has been mixed aggregations of nDPS and sDPS green sturgeon in Oregon and Washington estuaries and historical harvest in commercial and recreational fisheries in those estuaries.

The Oregon Department of Fish and Wildlife (ODFW) and Washington Department of Fish and Wildlife (WDFW) conducted a large-scale (approximately 1,500 green sturgeon tagged) mark-recapture study to estimate the population size of subadult and adult green sturgeon (described in NMFS 2015). That study generated a wide range of population estimates for adult and subadult green sturgeon (approximately 4,000 to 65,000 fish) in Oregon and Washington estuaries; the authors of that study concluded that an estimate of 40,000 fish was reasonable (NMFS 2015). Genetic analysis by Schreier et al. (2016) revealed $60 \%$ of sturgeon tissues collected in this study were from sDPS green sturgeon. Using estimates of annual run-size and spawning periodicity, Mora (2016) estimated the number of sDPS adults to be 1,990 $(95 \%$ confidence interval [CI] is equivalent to 1,172 to 2,808 adults). Mora (2016) also applied a conceptual demographic structure to that adult population estimate and estimated a sDPS subadult population of $10,450(95 \% \mathrm{CI}$ is equivalent to 6,155 to 14,745$)$. The population dynamics inferred in these studies-along with relatively low collection numbers of green sturgeon in SFE surveys-suggest the majority of the sDPS subadult and adult population occupies non-natal estuaries during summer months.

Sport and commercial harvest of green sturgeon has been prohibited since 2006, but past harvest of green sturgeon in California, Washington, and Oregon may have had a significant effect on sDPS green sturgeon abundance (Figure 1). Green sturgeon commercial harvest in the Columbia River Estuary over the last 20 years of record (1983 to 2002) averaged approximately

1,850 individuals per year (ODFW and WDFW 2002). Based on the distinct population segment composition study of green sturgeon in the Columbia River Estuary described in Schreier et al. (2016), the average annual harvest of sDPS green sturgeon in the Columbia River Estuary during this period was more than 1,100 individuals. It is highly plausible that the harvest of subadult and adult sDPS green sturgeon in other commercial and sport fisheries equaled or exceeded harvest in the Columbia River Estuary over this period. The USFWS (1996) estimated the total harvest of adult green sturgeon in Oregon and Washington fisheries at 5,000 to 10,000 individuals per year. The abundance of subadult and adult sDPS green sturgeon prior to Endangered Species Act listing is unknown, but if abundance has remained relatively stable (e.g., 12,440 individuals; Mora 2016), past harvest involved a large portion of the population in some years.


Figure 1. Historical yield of green sturgeon in Columbia River commercial fisheries in millions of pounds from Beamesderfer et al. (2005).

Green sturgeon have been detected in a relatively narrow range of depths and temperatures while in the marine coastal environment. Adults and subadults were typically detected in depths from 20 to 70 m and temperatures from 7.3 to $16.0^{\circ} \mathrm{C}$ by Erickson and Hightower (2007) and Huff et al. (2011). In the estuarine environment, green sturgeon are exposed to varying water temperature, salinity, and DO. For example, green sturgeon in coastal estuaries have been detected in water temperatures ranging 11.9 to $21.9^{\circ} \mathrm{C}$, salinities ranging from 8.8 to 32.1 ppt , and DO ranging from 6.54 to 8.98 milligrams of oxygen per liter (Kelly et al. 2007; Moser and Lindley 2007).

Limited numbers of green sturgeon have been encountered recently in white sturgeon stock assessments in the lower Columbia River, and recapture data from the Oregon and Washington estuary study described above have not been analyzed beyond 2012 due to poor recovery of tagged fish (O. Langness, WDFW, 2015, unpublished data, see "Notes"). Several studies have involved tagging subadult and adult fish with acoustic transmitters, but most tags were implanted before 2012 and will stop transmitting when they exceed the 5 - to 10 -year battery life. Movement by tagged fish in the SFE and Washington estuaries is monitored with receiver arrays maintained by the UCD Biotelemetry Laboratory and NMFS, respectively, while nearshore marine areas are monitored with compatible receiver arrays associated with other
groups. Subadult and adult nDPS and sDPS green sturgeon are occasionally reported as bycatch in federally managed ground fisheries (Lee et al. 2015; Anderson et al. 2017). The number of sDPS green sturgeon observed in the California halibut fishery (primarily sDPS) exceeded 100 individuals in 2006 and 2015 (Lee et al. 2015; NMFS, 2016, unpublished data, see "Notes") and the only targeted monitoring of subadults and adults in the ocean involves this fishery. Some sDPS green sturgeon caught incidentally by the California halibut fishery operating out of Half Moon Bay and San Francisco since 2015 have been tagged with satellite transponders. This program aims to understand post-release mortality, but also can provide information on movement for up to three weeks after release.

Adult white sturgeon. Tagging studies show that when not migrating, adult white sturgeon are usually found in brackish and estuarine habitat (Kohlhorst et al. 1991; Gleason et al. 2008), including most tidal tributaries to the SFE (Leidy 2007). Several observations of longdistance marine migrations suggest coastal habitats may also be utilized to some degree by those older life stages (Kohlhorst et al. 1991; DeVore et al. 1999; Welch et al. 2006; Ruiz-Campos et al. 2011; E. Miller, UCD Biotelemetry Laboratory, 2017, unpublished data, see "Notes"). For example, white sturgeon tagged in the SFE have been recaptured in the Columbia River, Chehalis River, and Willapa Bay, Washington, as well as the Umpqua River, Yaquina River, and Tillamook Bay, Oregon (Kohlhorst et al. 1991). Furthermore, one adult white sturgeon tagged in the Klamath River entered the ocean and traveled approximately $1,000 \mathrm{~km}$ north to be recaptured in the Fraser River (Welch et al. 2006), and an adult white sturgeon was caught by gillnet near Todos Santos Bay in Baja California (Ruiz-Campos et al. 2011).

Daily and seasonal movement of adult white sturgeon may be influenced by salinity and tides. Kohlhorst et al. (1991) found that white sturgeon followed brackish waters upstream in dry years and remained downstream in the SFE in wet years. Foraging movements of white sturgeon in the Columbia River Estuary were also influenced by diel cycles and salinity changes associated with tides (Parsley et al. 2008).

Growth rates of Sacramento-San Joaquin River adult white sturgeon have been estimated by Pycha (1956) and Kohlhorst et al. (1980) by counting annuli in pectoral fin ray sections and fitting von Bertalanffy growth models to the resulting age-length data. The age-length relationship differs between the two studies, with higher apparent growth rates-particularly for older white sturgeon-in the Pycha (1956) study. It is uncertain whether this discrepancy reflects a biological change in growth rate over time or a reflection of the high degree of uncertainty attributable to use of this ageing technique (Rien and Beamesderfer 1994). Growth rates of male and female white sturgeon from the SFE do not appear to be significantly different (Kohlhorst et al. 1980), but a study by Chapman et al. (1996) noted that the smaller individuals they sampled tended to be males while the sex ratio in their sample of larger individuals was skewed towards females.

Sediments in the SFE are heavily contaminated with a variety of toxins. Much of the research on contaminants and fish in the SFE has focused on selenium and mercury, but a suite of other toxins are also present at high levels in sediment throughout the estuary. Some of the
highest levels of toxic trace metals in the coastal United States were recorded in SFE (Cloern et al. 2006). As long-lived, bottom-feeding fish that spend the majority of their lives in the SFE, white sturgeon are highly susceptible to bioaccumulation of those contaminants. Tissues were analyzed from recently sampled adult white sturgeon in the SFE, and selenium and mercury were measured at levels known to impair condition factor (Gundersen et al. 2017). Gundersen et al. (2017) correlated organic contaminant and selenium levels in SFE white sturgeon with altered liver and gonad physiology and plasma hormone levels. Selenium levels in SFE white sturgeon appear to be a longstanding issue. High concentrations of selenium were found in white sturgeon tissue sampled in the SFE from 1970 to 2000 (Greenfield et al. 2003).

Available benthic food items in the SFE have changed in the recent past, and invasive invertebrates have replaced native mollusks and shrimps (Cohen and Carlton 1998). High levels of selenium have been found in common white sturgeon food items (Johns and Luoma 1988; White et al. 1988). Zeug et al. (2014) documented a predominance of invasive overbite clams (Potamocorbula amurensis)—a potential vector of bioaccumulation of toxic contaminants-in white sturgeon diets in the SFE. Sturgeon consume, but do not always digest, large quantities of the invasive overbite clam (Kogut 2008; Zeug et al. 2014); consuming indigestible clams may impair digestion of other items and affect sturgeon growth. Furthermore, bivalves have potentially higher levels of contaminants than other prey items in the SFE, and digested overbite clams may be a source of internal exposure to contaminants.

The recreational harvest of white sturgeon in California has remained relatively high in spite of a diminished population, such that monitoring of recreational capture and harvest is a primary management priority. Advances in recreational fishing techniques (e.g., use of shrimp as bait) since 1964 resulted in a rapid reduction in an initially large white sturgeon population (due to harvest prohibitions in the first half of 20th century) and continuous depletion of the few strong subsequent year classes. This is illustrated by commercial passenger fishing vessel log books data summarized by the Sturgeon Study, which shows a severe decline in catch and catch-per-unit effort (CPUE) during the first decade that shrimp were widely used as sturgeon bait (CDFW, 2016e, unpublished data, see "Notes"). This time series also shows a distinct increase in catch and CPUE in the early 1990s attributable to recruitment of the large, early-1980s year classes.

Actual harvest data along with estimates of harvest rate suggest that white sturgeon populations in the SFE may be limited in-part by overfishing. These data are generated by the use of an age-length key, Sturgeon Fishing Report Card data, and mark-recapture data (reward tags) by the Sturgeon Study. Based on ongoing study of age at length, white sturgeon aged 13 to 15 years fall mainly within the current slot limit of 40 to 60 inches FL. Thus, strong mid- to late1990s year classes were heavily fished over the last 10 years and reduced to a population of fewer than 20,000 fish (Gingras et al. 2014). As part of the Sturgeon Study, reward tags with multiple values are applied to adult white sturgeon. Disproportionately high redemption of $\$ 100$ tags relative to $\$ 20$ tags and $\$ 50$ tags indicates that much of the sturgeon harvest in the SFE goes unreported such that abundance estimates have likely been biased high.

Predation by native species on sub-adult and adult white sturgeon is not well documented. The CDFW has noted an increase in the incidence of California sea lions taking white sturgeon from trammel nets in the annual Sturgeon Study and the occasional presence of damaged or dead white sturgeon from apparent predation by California sea lions. Predation by marine mammals is a significant issue in the Columbia River. The estimated consumption of white sturgeon by pinnipeds in the Bonneville Dam tailrace ranged from approximately 150 to 3,000 individuals annually between 2006 and 2014 (Stansell et al. 2012, 2013, 2014).

## Spawning

Spawning green sturgeon. Adults mature at about age 15 and spawn during spring and summer every 2 to 6 years thereafter (NMFS 2015). The putative spawning reach on the Sacramento River extends from above the GCID oxbow to the confluence of the Sacramento River and Cow Creek (Heublein et al. 2009; Klimley et al. 2015). Only two spawning events have been documented outside of the Sacramento River (2011 and 2017) and both occurred in the Feather River (Fish Barrier Dam and the Thermalito Outlet). The Sunset Pumps Weir (33-km downstream of Thermalito Outlet) limits sturgeon passage in certain flows. Modifications of the Sunset Pumps Weir are being designed in hopes that their implementation would improve migration conditions for sturgeon. There are also periodic reports of adult green sturgeon at the base of Daguerre Point Dam on the Yuba River (a tributary to the Feather River), but spawning in the Yuba River has never been documented.

Green sturgeon enter the Sacramento River from late winter through early summer to spawn. Stranding of green sturgeon at migration barriers in the Yolo and Sutter bypasses has only been documented in spring 2011, 2016, and 2017 (wet years). Stranding in 2011 involved at least 24 adults at stilling basins below the Yolo and Sutter Bypass Weirs (Fremont and Tisdale Weirs; Thomas et al. 2013). Entrainment and stranding of adult white sturgeon and salmonids in bypasses is fairly common when bypasses flood, but it is uncertain how bypass channels (e.g., Toe Drain, Tule Canal) affect sturgeon spawning migration during low-flow years. Gravid adult white sturgeon are collected in the Toe Drain in low-flow years when passage barriers prevent access to spawning habitat (DWR 2015). In this regard, route selection may affect inter-annual spawning success as adult sturgeon entrained in bypass channels may not reach spawning habitat within an appropriate time. Records of catch in migratory and spawning habitats are collected through CDFW creel surveys (i.e., Central Valley Angler Survey, California Recreational Fisheries Survey), Sturgeon Fishing Report Cards, and state-mandated commercial passenger fishing vessel log books. A small number of post-spawn fish are periodically collected in San Pablo and Suisun bays during late summer and early fall by the Sturgeon Study.

With the exception of small-scale tagging in the Feather River and at bypass weirs, there is no program tagging adult green sturgeon for acoustic telemetry in the Sacramento River system. An acoustic receiver network is in place throughout the California Central Valley and SFE that is operated and maintained by the UCD Biotelemetry Laboratory along with state and federal agencies. Tagged adults have been detected in spawning habitat in the Sacramento River since 2004, and the majority were tagged outside of California in Pacific Northwest coastal
estuaries. Detections in this array provide information on migration timing and routes, spawning distribution and periodicity, and model inputs for run-size estimates from DIDSON surveys. Efforts are also underway by USFWS, UCD Biotelemetry Laboratory, and NMFS to map habitat in spawning and adult aggregation sites on the Sacramento River. The DWR monitors adult sturgeon in the Feather River, in part through telemetry that began in 2011 and DIDSON surveys that began in 2010. VEMCO VR2 acoustic receivers have been in place throughout the Feather River since 2008 and have detected approximately 15 green sturgeon that were tagged outside the Feather River (primarily in 2017; M. Manuel, PSMFC, 2017, personal communication, see "Notes").

DIDSON survey and telemetry data have been used to estimate green sturgeon run-size in the Sacramento River since 2010 (Mora 2016). Data and modeling from those surveys has greatly improved understanding of adult abundance. However, due to late maturation and variability in maturation age and spawning periodicity, it is challenging to relate trends in green sturgeon run-size to historical conditions that may have affected specific cohorts. Unlike with shorter-lived semelparous species, trends in spawner abundance and historical conditions associated with early life stage success-conditions that may have occurred decades before the observed trend-are intractable without accurate abundance modeling and ageing.

During seasonal RBDD gate closures before 2012, biologists periodically observed aggregations of adult green sturgeon directly below the gates in the spring and early summer. Spawning in this area has been verified several times by the collection of eggs (Poytress et al. 2015). Historical gate operations presumably influenced downstream hydrology, spawning, and holding at RBDD because aggregations have not been observed there since 2012. Although this spawning site was lost when gates were permanently raised in 2012, a new spawning site in an area historically inundated by the impoundment above RBDD was documented with the collection of eggs in RBDD rotary screw traps in 2016 and 2017 (W. Poytress, USFWS, 2017a, unpublished data, see "Notes").

The permanent raising of RBDD gates has the potential to affect multiple life stages of green sturgeon. Historically, adult green sturgeon arriving to RBDD after gate closure were relegated to a warmer reach of the Sacramento River for spawning. In extreme low flows, spawning areas below RBDD may be unsuitable for green sturgeon spawning in early summer. However, temperature above RBDD was suitable for spawning in early summer even in the recent historic drought. Thus, raising the gates at RBDD expanded the accessible spawning range for green sturgeon and-in what may be a key element of green sturgeon life history diversityfacilitated temporal and spatial variation in spawning distribution (A. Steel et al., UCD Biotelemetry Laboratory, 2017, unpublished data, see "Notes"). Furthermore, multiple dead adult green sturgeon have been documented below RBDD after gate closure. Those mortalities were attributed to injuries sustained during attempted downstream migration through a narrow opening below the closed RBDD gates. With gates permanently raised, that source of mortality no longer exists. Finally, the diversion facility at RBDD was recently upgraded. Although sturgeon entrainment at the RBDD facility has not been directly studied, entrainment risk for
larval green sturgeon has presumably been reduced by both raising dam gates and potentially improved hydraulics associated with modern fish screens.

After spawning, green sturgeon hold in-river for varying periods of time but most commonly leave the system the following fall (Benson et al. 2007; Heublein et al. 2009). Early outmigration in late spring or summer may be related to elevated flows (Benson et al. 2007; Heublein et al. 2009). For adults that over-summer in spawning or holding habitats, outmigration occurs in the late fall or (rarely) in winter after spending more than a year in freshwater (R. Chase, USACE, 2015, unpublished data, see "Notes"). Extended occupancy of those spawning and holding habitats may be related to hydrologic cues or food availability, but it is uncertain why occupancy sometimes lasts more than a year. Normal outmigration patterns for adults in the Feather River are unknown because of flow management and the passage barrier at Sunset Pumps Weir. However, the behavior of tagged fish in the Feather River suggests that adult outmigration there is similar to adult outmigration in the Sacramento River (i.e., both spring and summer outmigration and over-summer holding behavior; DWR, 2017b, unpublished data, see "Notes").

Post-spawn outmigration through the SFE is also variable. Adults have been detected migrating to the Pacific Ocean rather quickly (2 to 10 days) and remaining in the SFE for a number of months after leaving upstream holding habitats (R. Chase, USACE, 2015, unpublished data, see "Notes"). Movement between river habitat and the SFE has only been related to spawning, and adults do not appear to migrate upstream from the SFE to forage (R. Chase, USACE, 2015, unpublished data, see "Notes").

Spawning white sturgeon. Chapman et al. (1996) estimated that SFE white sturgeon females reach maturity at 95 - to $135-\mathrm{cm}$ FL and white sturgeon males reach maturity at 75 - to $105-\mathrm{cm}$ FL. The smallest sexually mature female described in the SFE watershed was $104-\mathrm{cm}$ FL (Chapman et al. 1996) and was estimated to be 9 years old using the von Bertalanffy growth equation from Brennan and Cailliet (1989). Female white sturgeon in the SFE are estimated to first spawn between 10 and 30 years old, and spawn every 2 to 4 years thereafter (Chapman et al. 1996).

White sturgeon in the SFE spawn in freshwater between mid-February and early June (Miller 1972; Kohlhorst 1976; Schaffter 1997; Jackson et al. 2016). Based on recent acoustic detections of tagged adults, white sturgeon may stage in the Delta or begin upstream migration to spawning habitat in fall or early winter (DWR 2015; Klimley et al. 2015). Spawning is known to occur in the San Joaquin River, the Sacramento River, some Sacramento River tributaries, and some watersheds on the northern California coast (Miller 1972; Kohlhorst 1976; Schaffter 1997; Jackson et al. 2016). However, long-term monitoring of adult and juvenile distribution indicates that most SFE white sturgeon originate from the Sacramento River. Spawning locations on the Sacramento River currently appear to be located from Colusa to the Verona area (Kohlhorst 1976; Schaffter 1997), although spawning likely occurs many kilometers upriver in some years (described above in the Fertilization to Hatch section). A relic population of white sturgeon persisted in Shasta Reservoir following construction of Shasta Dam (1940 to 1945), suggesting
historical spawning in upper Sacramento River tributaries such as the Pit River (Schaffter 1997; Moyle 2002). However, historical stocking records for sturgeon are poor, so it is unclear if white sturgeon in Shasta Reservoir were entirely of natural origin. Periodic aggregations of gravid white sturgeon and telemetry data in the Yolo Bypass and Bear River indicate spawning may extend to those areas in some years.

Fishery-independent monitoring of white sturgeon has occurred since the early 1950s and has persisted at low intensity until it became a CDFW emphasis in 2006 with Federal Endangered Species Act listing of sDPS green sturgeon. Adult white sturgeon abundance, relative abundance, harvest rate, and survival rate are estimated annually in the Sturgeon Study, which uses data from commercial passenger fishing vessels, various creel surveys, and markrecapture data from annual trammel net sampling in San Pablo Bay and Suisun Bay. Based on those data, trends in adult white sturgeon abundance and juvenile recruitment have been developed. However, current life stage monitoring is insufficient to develop mechanisms for episodic recruitment success or to address bias in abundance estimates and harvest rates associated with angler reporting.

Some directed sampling or incidental capture of adult white sturgeon occurs in spawning and migratory habitats. This includes sampling with fyke traps in the Yolo Bypass and Sacramento River, stranding at bypass weirs, and gill and trammel netting in San Joaquin River spawning habitat. Recapture of externally tagged fish (tagged in the Sturgeon Study) by the recreational fishery and data from Sturgeon Fishing Report Cards include information on sturgeon in migratory and spawning habitat. Other information on spawning adult behavior comes primarily from telemetry studies conducted by multiple groups (e.g., UCD, DWR, USFWS) throughout the SFE and spawning tributaries. Tissue from recreational harvest of white sturgeon (40- to 60 -inches FL) in SFE fishing derbies was also analyzed in a recent directed study of contaminants (Gundersen et al. 2017).

Migration barriers and stranding have been documented in the Yolo Bypass, Sutter Bypass, and Bear River since the mid-1980s. Those stranding events were associated with poaching and intensive legal harvest that has since been prohibited. The most recent documented events occurred in spring 2011, 2016, and 2017 (wet years) and involved stranding of adult green and white sturgeon in stilling basins below the Fremont and Tisdale Bypass Weirs. It is uncertain how bypass channels affect sturgeon spawning migration and success during low-flow years. For example, adult white sturgeon were captured in relatively high numbers from 2012 to 2014 (dry years) in a fyke trap deployed in the Yolo Bypass Toe Drain (DWR 2015). Delays in upstream sturgeon migration due to the complexity in fish passage at various bypass barriers (e.g., Lisbon and Fremont Weirs) may reduce spawning success. Stranding of adult white sturgeon in the Bear River may occur in normal or dry years. The last documented stranding event on the Bear River was in spring 2012, the first year of the recent drought. Winter and spring flow and water surface elevation on the Bear River can be highly variable even in relatively dry years. These flashy hydrograph conditions are thought to attract spawning white sturgeon to the Bear River, then subject those fish to high stranding risk.

The direct effects of consumption of contaminated prey species are briefly described above, but white sturgeon egg and larval survival may also be impacted through maternal transfer of contaminants. Exposure of gravid female white sturgeon to a diet enriched in selenium resulted in selenium incorporation to plasma vitellogenin and egg yolk proteins (Doroshov et al. 2007). Selenium has been measured in liver and gonad tissues of white sturgeon collected in the SFE at levels known to cause reproductive toxicity (Linares-Casanave et al. 2015) and exceeded concentrations associated with severe deformity and high mortality in the egg yolks of larvae ( 15 milligrams per gram; Doroshov et al. 2007).

Increasing streamflow is believed to be an important cue for migration and spawning of white sturgeon across their range (Schaffter 1997; Hildebrand et al. 1999; Paragamian and Wakkinen 2011; Jackson et al. 2016). During a drought year, Schaffter (1997) documented white sturgeon spawning in the Sacramento River following $40 \mathrm{~m}^{3} / \mathrm{s}$ increases in streamflow, which would correspond to an increase in outflow or runoff of approximately $3,456,000 \mathrm{~m}^{3}$ or 2,800 acre-feet per day. Similarly, white sturgeon spawning was observed at all four San Joaquin River sampling locations following short-duration streamflow pulses (approximately 18 to $40 \mathrm{~m}^{3} / \mathrm{s}$ ) during otherwise low streamflow conditions in 2012 (Jackson et al. 2016) and again in 2016 (Z. Jackson, USFWS, 2017, unpublished data, see "Notes").

There is little quantitative information on spawning site habitat requirements. In the Sacramento and San Joaquin watersheds, white sturgeon spawning habitat has substrate ranging from silt-and-sand to gravel-and-cobble and water velocities ranging from low to high (Kohlhorst 1976; Schaffter 1997; Jackson et al. 2016). As indicated by the presence of eggs or larvae, white sturgeon have been observed to spawn at water temperatures of 8 to $23^{\circ} \mathrm{C}$, with peak spawning occurring around $14^{\circ} \mathrm{C}\left(8\right.$ to $19^{\circ} \mathrm{C}$, peak $14^{\circ} \mathrm{C}$, Moyle $2002 ; 7.8$ to $17.8^{\circ} \mathrm{C}$, peak $14.4^{\circ} \mathrm{C}$, Kohlhorst 1976 ; and 12 to $16^{\circ} \mathrm{C}$, Schaffter 1997). Spawning in the San Joaquin River has occurred at higher temperatures ( 14.2 to $26.7^{\circ} \mathrm{C}$, Jackson et al. 2016), although viability of larvae at the upper end of this range is unlikely. White sturgeon in the SFE typically return promptly to the estuary after spawning. However, adult holding through summer in freshwater spawning habitat has been documented on the Sacramento and San Joaquin rivers (Klimley et al. 2015; M. Manuel, PSMFC, 2016, personal communication unreferenced, see "Notes").

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# Estimating spatial-temporal differences in Chinook salmon outmigration survival with habitat and predation related covariates 

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Low survival rates of Chinook salmon smolts in California's Central Valley have been attributed to multiple biological and physical factors, but it is not clear which factors have the largest impact. We used five-years of acoustic telemetry data for 1709 late-fall Chinook salmon smolts to evaluate the effect of habitat and predation related covariates on outmigration survival through the Sacramento River. Using a Cormack-Jolly-Seber mark-recapture model, we estimated survival rates both as a function of covariates (covariate model) and as a function of river location and release year (spatial-temporal model). Our covariate model was overwhelmingly supported as the preferred model based on model selection criteria, suggesting the covariates adequately replicated spatial and temporal patterns in smolt survival. The covariates in the selected model included individual fish covariates, habitat specific covariates, and temporally variable physical conditions. The most important covariate affecting salmon survival was flow. We describe the importance of these parameters in the context of juvenile salmon predation risk and suggest that additional research on predator distribution and density could improve model estimates.

## Introduction

Salmon smoltification and outmigration from freshwater rearing habitats is a time of increased mortality as fish undergo physiological changes and encounter new stressors (Connor et al. 2003; Welch et al. 2008; Nislow and Armstrong 2012). Much of the research on outmigration mortality has examined the effect of dam passage on survival (Skalski et al. 2001; Williams et al. 2001; Welch et al. 2008; Elder et al. 2016), with relatively few studies focusing on how other environmental conditions affect survival. Environmental conditions that have been linked to outmigration mortality include flow (Connor et al. 2003; Smith et al. 2003; Michel et al. 2015; Courter et al. 2016), temperature (Connor et al. 2003; Smith et al. 2003), turbidity (Gregory and Levings, 1998; Smith et al. 2003), and predation (Beamesderfer et al. 1996; Friesen and Ward 1999; Schreck et al. 2006). Some of these factors, such as water temperature and flow, are expected to increasingly affect juvenile salmon survival and population production as the climate changes (Jonsson and Jonsson 2009; Mantua et al. 2010; Katz et al. 2013; Russell et al. 2012). Many of the published correlations between outmigration survival and environmental characteristics have examined survival over relatively large temporal and spatial scales, whereas individual fish experience mortality at a particular time and place. To better understand how habitat and predation related covariates influence salmon smolt mortality it is necessary to look at the conditions experienced by fish as they are migrating through a habitat.

Most Chinook salmon spawned in the Sacramento River have long outmigrations ( $\sim 500$ kilometers) through multiple habitats, and it is believed that the precipitous decline of multiple salmon populations in this system is partially due to anthropogenic habitat modifications and poor out-migration survival (Yoshiyama et al. 1998; Katz et al. 2013; Michel in press). Currently, survival of Chinook salmon smolts from the Sacramento River to the ocean is
markedly lower than smolt out-migration survival from the Columbia and Fraser rivers in the Pacific Northwest region of the United States and Canada (Welch et al. 2008; Michel et al. 2015; Buchanan et al. 2018), but it is unclear what factors cause this increased mortality. Previous research has found that interannual variability in smolt survival is much greater in the Sacramento River than in the Sacramento-San Joaquin delta or the San Francisco Bay, suggesting that the river has a large influence on outmigration success (Michel et al. 2015). Within the river, outmigration survival rates vary both spatially and interannually (Singer et al. 2013; Michel et al. 2013). This spatial and temporal variability is likely driven by changes in the underlying environmental and habitat features comprising the river landscape.

Identifying the main factors that affect smolt mortality is important to establish restoration priorities and give managers quantitative data on how to optimize survival of threatened salmonids. This is especially important given recent findings that suggest outmigration survival has a larger effect on smolt-to-adult ratios than marine survival (Michel in press). To identify which factors had the largest influence on outmigration survival, we developed a series of mark-recapture models using five years of acoustic telemetry data for latefall Chinook Salmon. We then used model selection to identify which covariates had the largest influence on survival. Our analysis builds upon the research conducted by Singer et al. (2013) and Michel et al. (2015), whose primary objective was to identify temporal and spatial differences in the mortality of outmigrating juveniles. In contrast, our objective was to model survival solely as a function of covariates that were hypothesized to affect salmon survival through habitat modification and increased predation risk.

## Methods

Study Area.-The northernmost extent of our study was the release location for late-fall run smolts at the Coleman National Fish Hatchery (Figure 1). We included all detections of acoustically tagged fish from the release location to the ocean, but we only included covariates for reaches between the release location and the I-80 Bridge in Sacramento. This was for two reasons: 1) hydrodynamic model estimates for temperature and flow below the city of Sacramento were not as reliable as the upstream estimates, and 2) survival variability was much larger in the reaches upstream of Sacramento than in the Sacramento-San Joaquin delta or San Francisco Bay (Michel et al. 2015). Riverine habitat varied spatially across the $\sim 300$ kilometers of Sacramento River that defined our study area. There was a general upriver to downriver gradient in habitat features associated with human influence. For example, diversion density, amount of armored bank, and agriculture/developed land use increased from the upper to lower reaches.

Acoustic tagging. - Late-fall run Chinook salmon were obtained from the United States Fish and Wildlife Service (USFWS) Coleman National Fish Hatchery, implanted with acoustic tags, and released annually during the winter months (December and January) from 2007 through 2011. Details regarding the surgical procedures and initial acoustic tag study design are documented in Michel et al. 2013 and Ammann et al. 2013. Briefly, small acoustic tags (Vemco $69 \mathrm{kHz}, 7 \mathrm{~mm}$ dia. X 20.5 mm long, weighing 1.8 g in air and 1.0 g in water) were surgically implanted into the peritoneal cavity of anesthetized fish through a 12 mm incision. The incision was then closed with two simple interrupted stitches with nonabsorbable nylon cable-type suture. All fish were allowed to recover for a minimum of 24 hours before release. During the first year of this study (2007), smolts were tagged and released directly into Battle Creek, a tributary of the Sacramento

River where the Coleman Hatchery is located (Figure 1). From 2008-2010 tagged smolts were released concurrently from three locations along the mainstem Sacramento River: Jelly's Ferry, Irvine Finch and Butte City to increase sample size of fish detected throughout the river and to estimate differences in survival between newly released fish and those released upstream (Figure 1). In 2011, all fish were released at Jelly's Ferry due to a slightly reduced sample size. In addition to the acoustic tag data ( $\mathrm{n}=1350$ ) utilized in Michel et al. (2013) and Michel et al. (2015), we used acoustic tag data provided by the USFWS ( $n=359$ ). These fish were tagged in accordance with the procedures described above, but released directly into Battle Creek in 2010 and 2011, simultaneous to the release of the remaining hatchery stock (batch released). The mean hatchery release during these dates was approximately 600,000 fish (range: $155 \mathrm{k}-889 \mathrm{k}$ ).

Acoustic receivers were located from the fish release sites in the upper Sacramento River to the Golden Gate Bridge at the entrance to the Pacific Ocean. We divided the Sacramento mainstem study region into 19 reaches demarcated by 20 acoustic receiver locations along the mainstem Sacramento River (Figure 1). These reach locations were selected based on interannual consistency in receiver location throughout the 5-year study period; however, detections from inconsistently deployed receivers were retained to improve precision of survival and detection probabilities (see 'mark-recapture analysis' section).

Acoustic telemetry data processing. -We used a series of algorithms to ensure our acoustic telemetry data did not include any false detections. The acoustic receivers automatically processed detection data by dropping incomplete codes from the detection file. To ensure that we removed any false detections due to acoustic pulse train collisions, we performed several additional quality control procedures. First, we removed all detections that occurred prior to the
release date and time. We then removed all detections from fish that had only a single detection throughout the study. We required three or more detections within 10 days at a single receiver location to verify those detections were not the result of pulse train collisions. We also examined the encounter history of each individual fish and removed any detections that indicated upstream movements. Furthermore, we calculated the transit time between receivers (number of river kilometers between receivers divided by the difference in seconds between the last upstream detection and first downstream detection) and removed any detections resulting from a fish traveling at speeds greater than $10 \mathrm{~km}_{\mathrm{km}} \mathrm{hrur}^{-1}\left(2.78 \mathrm{~m} \mathrm{~s}^{-1}\right)$. We also assumed that any tag consistently detected at a single receiver location for more than 4 weeks, and not subsequently detected downstream, was a mortality. We selected the 4 weeks cutoff after a preliminary examination of the data indicated fish detected at a single location for more than 4 weeks were never detected at another receiver. These fish ( $\mathrm{n}=58$ ) were considered known mortalities (i.e., treated the same way as a harvested fish in a standard mark-recapture model) and did not have any impact on the estimated survival or detection probabilities downstream from where the presumed mortality occurred.

Mark-recapture analysis.-To estimate survival of out-migrating late-fall run Chinook salmon, we fit a Cormack-Jolly-Seber (CJS) survival model (Cormack 1964; Jolly 1965; Seber 1965) using the marked (Laake et al. 2013) and RMark package (Laake and Rexstad 2008, Collier and Laake 2013) within the R programming language (version 3.3.1, R Core Team 2017). We used the marked package for the initial model selection due to its computational efficiency and RMark for parameter estimation due to better analytical functionality (see appendix). The CJS model was originally conceived to calculate survival of tagged animals over time by recapturing
individuals and estimating survival and recapture probabilities using maximum likelihood. A spatial form of the CJS model can be used for species that migrate unidirectionally, and are recaptured, throughout a migratory corridor (Burnham 1987). Using this space for time substitution, we used individual fish encounter histories to estimate the likelihood that a fish would survive and be detected at each receiver (Lebreton et al. 1992). In the standard formulation of the CJS model, detection probabilities are estimated for a single resampling occasion (i) in time or space. However, our encounter histories included detections both from receivers at the reach boundaries as well as receivers within the reach. Thus, our estimated detection parameter represents the probability of detection from receiver (i) to receiver (i+1).

Spatial-temporal model.- Prior to fitting a covariate model, we fit a model that estimated a different survival for every reach in every year. This spatial-temporal model provided a means to evaluate how well our covariate model replicated outmigration survival. We assumed that differences between the spatial-temporal model and the covariate model were the result of unaccounted variance due to missing covariates. Due to the inherent complexity of the Sacramento River ecosystem, it was not feasible to measure or estimate all potential covariates that influence salmon survival. For example, there is no hydrodynamic model currently capable of estimating turbidity levels throughout the river.

The spatial-temporal smolt survival estimates were converted to survival per 10 km values to allow for comparisons between reaches via:

$$
\Phi_{10}=\sqrt[1]{\Phi_{\mathrm{R}}}
$$

where $\Phi_{10}$ is the survival estimate per $10 \mathrm{~km}, \Phi_{\mathrm{R}}$ is survival per reach, and 1 is reach length divided by ten.

Covariate model.-We included multiple individual, release group, reach specific, and timevarying covariates in our analysis to identify the factors contributing to the mortality of outmigrating smolts. Each of the covariates included in the analysis had an a priori hypothesized relationship with smolt survival (Table 1).

The individual covariates we included were length, condition, and transit speed. Fish size has been known to influence juvenile salmon survival (Zabel and Achord 2004), thus we included both length and condition factor (Fulton's $k=\frac{W}{L^{3}} * 100$ ) as individual covariates.

Length was hypothesized to affect survival through predator gape limitation whereas condition factor is an indicator of fish health and stamina. We also included individual fish transit speed within each reach, which we estimated with a mixed effects model (see details below), because faster moving fish would have less exposure to predators.

Release group effects included release group size, a release reach effect, and the mean annual flow at Bend Bridge (see Figure 1 for location) in the release year. We included a binary group covariate for release group size to distinguish fish released in synchrony with thousands of other hatchery fish from those released in small (e.g. 50-100 fish) batches based on the hypothesis that large releases would result in increased survival due to predator swamping (Fritts and Pearsons 2008; Furey et al. 2016). To test the hypothesis that the potential survival advantage of large releases would diminish as fish diffused downstream, we also included an interaction between release group size and distance from release site. We included a release reach effect to test if survival in the first reach after release differed from fish released upstream of the release site. We hypothesized survival rate in the release reach would be lower because newly released hatchery fish are naïve and more susceptible to predation (Alvarez and Nicieza

2003; Huntingford 2004; Jackson and Brown 2011). The final release group specific covariate was the mean annual flow measured at the Bend Bridge gauge during the months of smolt outmigration (December-March). This covariate was included to test if survival decreased in low flow (e.g., drought) conditions. Bend Bridge was selected to represent mean annual flow because it was upstream of the major tributaries and diversions and was collinear with the flow measurements throughout the river.

The reach specific covariates included in the model were sinuosity, diversion density, adjacent cover density, and off-channel habitat density. We selected these features because we hypothesized they would influence survival by affecting predation risk. More natural habitats with increased sinuosity, adjacent cover density, and off-channel habitat density are hypothesized to provide more predator refuge (reviewed by Roni et al. 2014). Furthermore, agricultural and municipal water diversions along the Sacramento River pose a risk to outmigrating salmon through direct entrainment (Hanson 2001; Kimmerer 2008; Mussen et al. 2014), as well as indirectly by providing structure for salmonid predators (Sabal et al. 2016). We hypothesize that the latter has more of an effect on Chinook smolt survival since the diversions are typically not in operation during the months of outmigration. These reach specific data were derived from GIS layers available from multiple sources (Table 1) and plotted in a Geographic Information System (using ESRI ArcGIS 10.3). Because we were using static GIS layers, we were unable to determine if the available off-channel habitats were connected to the mainstem under different flow regimes. We were also unable to measure inter-annual differences in adjacent cover density.

The time-varying covariates we included in the model were flow and temperature, which we obtained from the River Assessment for Forecasting Temperature (RAFT) model. The RAFT
model is a 1-dimensional physical model that estimates temperature and flow every 15 -minutes at a 2 km spatial resolution (Pike et al. 2013). We included temperature as a covariate because predator metabolisms, and predation rates, increase at higher temperatures (Petersen and Kitchell 2001). We included multiple aspects of flow (see below) derived from the RAFT model because flow is important to smolt survival (Kjelson and Brandes 1989; Cavallo et al. 2013; Zeug et al. 2014; Michel et al. 2015; Courter et al. 2016). We associated values for each of these variables with each tagged fish in space and time at the $2-\mathrm{km}$ spatial resolution, and then calculated the reach-level means for each fish for each variable. We assumed that RAFT model predictions were accurate (i.e. we did not propagate RAFT model uncertainty into the mark-recapture model) based on results from model validations (Pike et al. 2013; Daniels et al. 2018).

Due to the importance of flow to outmigrating salmon survival, we fit a variety of models with different flow standardizations to test which aspects of flow had the largest influence on survival. We scaled (subtracted the mean and divided by the standard deviation) the time-varying estimates of flow in two ways: 1) by reach, and 2) by year and reach. We scaled by reach to detect within reach patterns of survival relative to inter-annual flow conditions. In other words, is reach-specific survival dependent on whether flows are above or below average compared to other years? Since this parameter could distinguish between annual differences in flow (i.e., low flow versus high flow year), we did not include the annual flow at Bend Bridge in any models that included flow scaled only by reach. Thus, we could test if the spatially explicit estimates of flow added any additional information beyond a single measure of mean annual flow. The year and reach scaling tested whether intra-annual changes in flow within a reach were important to salmon survival. In other words, we wanted to determine if periods of higher flows within a reach, such as those after large precipitation events, would increase survival relative to periods of
lower flows within the same year. This hypothesis was based on previous studies that have observed large increases in survival due to controlled changes in flow rate (Cavallo et al. 2013; Courter et al. 2016). Scaling by both year and reach removes the effect of annual differences in flow such that it is impossible to distinguish high flow years from low flow years with this parameter. Thus, models in which flow was scaled by year and reach could also include the mean annual flow at Bend Bridge. We also fit models that included an interaction between the mean annual flow and the time-varying flow standardized by year and reach to test the hypothesis that precipitation events would have a larger impact on survival in years with lower flows. We tested this hypothesis based on work by Courter et al. (2016) that suggested flow has a large impact on survival in reaches with relatively low flow but has a negligible impact in reaches with high flow.

To estimate the effect of a covariate (e.g. flow) on fish survival throughout a reach, it is necessary to have a covariate value for every fish in every reach. When we did not detect an individual fish at a receiver there was uncertainty as to when that fish might be within that reach and, thus, what covariate value should be used. To impute covariate data in locations where fish were not detected, we fit a mixed-effects model where the response was transit speed of individual fish detected at both upstream and downstream acoustic receivers of a single reach. Our independent covariates were release year, release week, reach, and fish condition. We also included a random intercept for each individual fish to account for individual behavioral variability. We fit the model using the 'Ime4' package (Bates et al. 2015) and selected the model with the lowest Akaike's Information criterion (AIC; Burnham and Anderson 2002). To verify that the mixed-effects model did not unduly violate any assumptions, we examined model diagnostics (QQplot and residuals) using the DHARMa package (Hartig 2018). We then used the
results from the mixed-effects model based on detected fish to estimate the dates and times undetected fish were present within each reach.

Prior to fitting the CJS models, all continuous covariates were standardized by subtracting the mean and dividing by the standard deviation. Standardized coefficients could then be interpreted as the estimated change in survival predicted from one standard deviation increase in the covariate value. We also conducted pairwise comparisons of all continuous individual, habitat, and physical covariates to determine if any covariates were collinear (Supplemental figure S1). From pairs that had correlation coefficients greater than 0.7 (Dormann et al. 2012), we selected a single covariate that we hypothesized would have the largest influence on survival based on results from previous studies.

Model selection.-We fit a series of CJS models to determine which covariates (individual, release group, reach specific, or time varying) had the greatest impact on out-migrating smolt survival. With the exceptions of collinear variables and the restrictions noted above, we fit models with all possible combinations of covariates and selected the most appropriate models with adequate support using Quasi-Akaike's information criterion (QAICc) (Burnham and Anderson 2002). QAIC adjusts the AIC value based on an overdispersion parameter ( $\hat{c}$ ), which we estimated using the median ĉ method for the spatial temporal model within program MARK (White and Burnham 1999). If the observed data has no overdispersion, $\hat{c}$ will be approximately equal to 1 . Values of c greater than 4 indicates the model structure is inadequate and does not account for a sufficient amount of variation in the data (Burnham and Anderson 2002). Our median $\hat{c}$ was 1.45 , indicating the model was satisfactory but slightly overdispersed. We selected the most appropriate model by examining the difference in QAIC values between each model
and the model with the lowest QAIC ( $\triangle$ QAIC). We assumed models with $\triangle$ QAIC $<2$ had equal support (Burnham and Anderson 2002); thus, if multiple models had a $\Delta$ QAIC $<2$, we selected the one with the fewest parameters.

Covariate plots.-To determine which covariates had the largest influence on survival, we plotted the $\triangle$ QAIC between the selected covariate model and the same model without a single covariate. In the case of covariates that were included as main effects and in an interaction, we also removed the interaction. We will refer to these models as our covariate importance analysis.

We used marginal model plots to evaluate the effect of individual covariates on outmigrating smolt survival. To produce these plots, the $\beta$ parameter coefficients from the selected covariate model were used to simulate what survival would be for the $95 \%$ observed range of a single covariate. With the exception of reach length, covariates not included in the individual response plots were set to zero for binomial covariates or to their mean for continuous covariates. Reach length was set to 10 km for all plots except the one that explicitly focused on the effect of reach length.

## $<A>$ Results

## $<\mathrm{B}>$ Spatial-temporal model

Based on the model that included a reach by year interaction, we observed that survival was not consistent spatially or temporally. We saw a general trend of lower per-reach survival in the upper and middle reaches, compared to the more downstream reaches, but the location and severity of mortality varied inter-annually (Figure 2). The high flows in 2011 negatively
impacted our detection efficiencies, rendering 12 receivers without reliable detection data; however, the detection efficiencies in the lower river and the estuary remained high and provided sufficient data to estimate out-migration survival through the river. The receiver locations with low detection efficiencies often resulted in survival estimates of 1 due to numerical boundary issues.

## $<\mathrm{B}>$ Covariate model

The selected covariate model had 15 survival parameters and fit the data nearly as well as the spatial-temporal model that had 110 survival parameters. As a result, the covariate model had a much lower QAICc value $(\triangle \mathrm{QAICc}=55.90)$, implying it was more parsimonious. Although the covariate model showed some deviation from the spatial-temporal model, especially in the most upstream reaches, these tended to be relatively small and not significantly different from zero (Figure 3).

The top covariate model included a combination of an individual covariate (transit speed), group covariates (batch release, interaction between batch release and distance from release site, release reach, and the mean annual flow recorded at Bend Bridge), reach specific covariates (reach length, sinuosity, and diversion density), and time-varying covariates that were estimated for when a fish passed through a specific reach (reach flow, interaction between reach flow and annual flow, and water temperature). Based on the standardized beta coefficients for the covariates (Table 2) and the results from the covariate importance analysis (Figure 4), annual flow and reach length had the largest influence on survival. Flow was the most important covariate in predicting outmigration success, with increased levels of annual flow correlating to increased smolt survival (Figure 5a). Above average reach flows within a year (e.g., large
precipitation events) helped improve survival much more in low flow years than in high flow years. As would be expected, longer reaches had lower survival rates (Figure 5b). Based on the covariate importance analysis, the next most important variables affecting survival were diversion density, release reach, and the interaction between release group size and distance from release location. Survival increased relative to diversion density (Figure 5c), was lower in the first reach after release (Figure 5d), and increased for approximately the first 200 km from the release site when fish were released concurrently with thousands of hatchery fish (Figure 5e). Finally, the covariates that had the least effect on survival were sinuosity (increase), transit speed (increase), and water temperature (decrease) (Figure 5 f-h).

## <a>DISCUSSION

Conservation of salmonid populations depends on understanding what physical and biological factors have the largest impact on mortality during different life history stages. Recent research has shown that the outmigration period may have the largest influence on smolt to adult survival rates and cohort strength (Michel in press). Therefore, identifying the primary factors that affect survival of outmigrating smolts can help prioritize management actions that will be most beneficial to the conservation of imperiled populations. While we could not include all possible sources of mortality in our analysis, we conclude that flow remains the single most influential factor for determining survival of late-fall Chinook salmon smolts outmigrating from California's Central Valley.

## $<\mathrm{B}>$ Spatial and temporal survival heterogeneity

The spatial-temporal model indicated that survival through different reaches varied interannually, which is likely a result of the dynamic nature of the Sacramento River system. Overall, we can conclude from our reach-specific survival estimates that increased mortality rates occurred most frequently in the upper and middle regions of the Sacramento River, and decreased rates occurred through the lower reaches. We compared the observed values for the covariates included in the selected model to determine if fish had different behaviors in the upper reaches and if any aspects of the physical habitat differed. The most striking difference between the upper reaches and the lower reaches was the diversion density. This implies that increased diversion desity, and the coincident anthropogenic habitat modifications of the lower river, reduced mortality of outmigrating smolts. Much of the previous work that has examined the effects of habitat modification and restoration on salmonid populations has focused on egg incubation, freshwater/estuarine rearing, and available spawning habitat (reviewed in Roni et al. 2014) or the effects of fish passage on outmigration mortality (Skalski et al. 2001; Williams et al. 2001; Welch et al. 2008; Elder et al. 2016). We do not know of any studies that have explicitly looked at the effect of channel alteration on salmon outmigration survival. A valuable future study would be to examine if channelized habitats have lower predator densities or if the deeper waters make it easier for salmon to avoid predators.

In addition to the higher mortality rates in the upper reaches, the biggest discrepancies between the spatial-temporal model and the covariate model also occurred in the upper reaches. This suggests our covariate model would benefit from including additional covariates that contributed to smolt mortality in the upper reaches. Based on previous research, we believe that including covariates such as turbidity and predator density would likely improve our explanatory power. Turbidity likely improves salmon survival by decreasing predation risk (Gregory and

Levings 1998). Likewise, high predator densities in the upper and middle reaches may partially explain the increased mortality rates in these locations. Naïve, hatchery raised fish, are more susceptible to predation after release (Alvarez and Nicieza 2003; Huntingford 2004; Jackson and Brown 2011). This was reflected in our covariate model where newly released fish had a lower survival rate than fish released upstream. Including turbidity and predator density in a mark recapture model could improve model fit and provide important information necessary to develop a purely mechanistic model to estimate outmigration mortality.

## $<\mathrm{B}>$ Time-varying covariates

Model selection for the covariate model provided insight into which time-varying physical covariates had the largest influence on survival of out-migrating late-fall Chinook Salmon. Flow exerted the greatest overall effect on outmigration success, with increased annual flow positively related to increased smolt survival. Studies have repeatedly demonstrated that flow is the most important factor affecting survival of Chinook salmon (Conner et al. 2003, Smith et al. 2003; Zeug et al. 2014; Michel et al. 2015). In addition to the effect of annual flow, we also found that variability in flow within a reach affected survival rates, particularly in low flow years. If flow within a reach was well above the annual average, as it would be after a precipitation event, there was relatively little ( $1.6 \%$ per 10 km ) difference between survival in a low and high flow years. In contrast, below average flows within a reach resulted in large (5\% per 10 km ) differences in survival between low and high flow years. This provides a potential explanation for results observed by Courter et al. (2016), where survival was highly dependent on flow within a low flow ( $<125 \mathrm{cms}$ ) reach, but had no effect in a reach with higher flows (100-300 cms).

Our study also builds on previous work by including measurements of both spatially explicit flow and transit speed as covariates in our model. This allowed us to separate the effect of flow from transit speed, suggesting that there are features inherent to flow itself, not just its effect on travel time, which affects survival. Flow has been significantly reduced and homogenized in the Sacramento River system from historic levels (Buer et al. 1989), in particular during the winter months when runoff from storm events is captured behind dams. Flow magnitude affects the amount of off-channel and floodplain habitat available for juvenile salmon rearing (Nislow and Armstrong, 2012; Merenlender and Matella 2013). Fish residing in these habitats have accelerated growth rates that may aid individuals in predator avoidance and survival (Sommer et al. 2001; Limm and Marchetti 2009). Furthermore, the highest sediment loads for the Sacramento River were observed with the highest peak flows (Stern et al. 2016), which can increase turbidity rates and decrease predation rates (Gregory and Levings 1998). Whatever the specific mechanism, flow was clearly the most important factor influencing the outmigration success of late-fall run Chinook smolts in 2007-2011. Perhaps more importantly, the effect of flow propagates throughout a cohort's life history and can be used to estimate smolt-to-adult ratios (Michel in press). Threshold flow values could be determined through combined controlled-release and tagged-release studies in the Central Valley.

We also found survival was higher at lower water temperatures. We hypothesize that this effect was the result of increased predator metabolism, and thus consumption, at increased temperatures (Petersen and Kitchell 2001). This effect was relatively minor (1.3\% per 10 km ) over the small range of temperatures we observed during the fall-run winter outmigration months. However, we expect this effect will be more pronounced for fall and winter run fish that
are outmigrating during warmer months and may exhibit adverse responses to warmer temperatures (Baker et al. 1995; Lehman et al 2017).

## $<\mathrm{B}>$ Release group covariates

Acoustically tagged fish had higher survival rates when they were released concurrently with thousands of hatchery fish. Based on the interaction between release size and distance from release location, this effect persisted for approximately 200 km from the release location. One explanation for this improved overall survival is the theory of "predator swamping;" whereby predators, inundated by prey, pose less of a threat to individual smolts. This effect has been demonstrated for Chinook salmon in the Yakima River (Fritts and Pearsons 2008) and juvenile sockeye salmon in British Columbia (Furey et al. 2016). We examined the difference in arrival times at the acoustic receiver locations for each of the release groups, and found that fish from the same release group arrived at the same location within approximately 24 hours for the first 100 km (Supplemental fig S2). After the first 100 km the river has more channel alterations and fish arrival times were more dispersed. However, fish survival rates in these lower sections of the river were generally higher than in the upstream reaches, most likely due to decreased predation rates in the channelized portions of the river.

## $<\mathrm{B}>$ Individual covariates

Predicted transit speeds were also an important factor, with increasing transit speeds corresponding to increased survival. For out-migrating yearling smolts, it is likely that transit speed in the context of our study is a proxy for duration of exposure to mortality factors. Previous studies have found that survival rates decline over longer migration distances (Bickford
and Skalski 2000; Muir et al 2001; Smith et al. 2002). However, these studies have primarily found that survival was related to distance traveled but not to travel time. Anderson et al. (2005) explained this apparent discrepancy by suggesting that survival was a function of both migration distance and predation risk. This provides further motivation to study the factors that influence the spatial distribution, and density, of salmon predators throughout the Sacramento River.

## $<\mathrm{B}>$ Reach specific characteristics

Model selection results provided evidence that reach length, diversion density, and sinuosity were associated with outmigrating smolt survival. After accounting for all other covariates, survival was higher with increasing sinuosity, suggesting that more natural river conditions were better for smolt survival than the deeper and more armored portions of the river. This result is in contrast to our other finding that the highest survival rates were in the lower, more channelized sections of the river. We suspect that the larger covariate effect of diversion density accounts for the variation associated with increased survival in the lower reaches. Because the diversions are typically not operational during the period when late-fall Chinook are outmigrating, we suspect this effect is more a function of the habitat conditions in locations where diversions are more abundant. Diversions were highly correlated to other habitat variables typical of agricultural zones; namely depth, armored banks and agricultural and developed land use (Supplemental figure S 1 ). Because we did not wish to obfuscate the results of our analysis, we withdrew these collinear factors from our modeling efforts, but the role of "diversions" on survival could be equally viewed as the role of depth, agriculture and developed land, and armored banks. These modified habitats may result in reduced predator densities and predation mortality.
$<B>$ Conclusions
Flow, diversion density, and release strategy had the strongest influence on survival of out-migrating, hatchery origin, late-fall run Chinook salmon during the 2007-2011 water years. For years with high flow, gains in in-river survival can lead to a three-fold increase in total outmigration survival, while survival in the delta and estuary remain the same (Michel et al. 2015). There is limited natural habitat remaining for Chinook salmon in the Central Valley as a result of human activities, and increasingly managers are turning to habitat restoration efforts to restore salmon populations. When we compare physical covariates, metrics for habitat features and individual covariates, flow remains the most important factor affecting out-migration survival of late-fall run hatchery raised smolts. Although our study used hatchery fish, which have limitations as wild fish surrogates, these results suggest that maintaining flow during periods of salmon outmigration is an important step towards conserving Chinook salmon in the Central Valley.

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Figure Captions
Figure 1: Map of the mainstem Sacramento River. Our study area extended from above Red Bluff in the north to the city of Sacramento in the south. Late-fall run Chinook salmon yearling smolts were released at Battle Creek, Jelly's Ferry, Irvine Finch or Butte City during the winter (Dec-Jan) of each of our study years. The locations of the 20 acoustic receivers that delineated our 19 river reaches are shown as red stars.

Figure 2: Map depicting reach-specific survival estimates (per 10km) for 2008-2010. Colors represent per reach survival risk and standard error is represented as the grey buffer surrounding each reach. The values adjacent to each reach represent the survival estimate for a given reach (per 10 km ) from our full survival model.

Figure 3. Difference between survival estimates in the spatial-temporal model and the covariate model for each reach (labeled as the distance (River km) between the upstream boundary and the Golden Gate Bridge). Negative values represent occasions when the covariate model had a larger estimate of survival and was presumably missing covariates that increased smolt mortality. Error bars represent the $95 \%$ confidence interval estimated with the delta method.

Figure 4: A barplot depicting the results of covariate removal analysis to determine the importance of each variable to the final model. Delta QAIC values represent the change in QAIC when specific variables are removed from the full model.

Figure 5: Covariate response plots showing the effect of the individual covariates on the apparent survival rate through a 10 km reach. The grey shaded region represent the $95 \%$ confidence interval.


Fig 1: Map of the mainstem Sacramento River. Our study area extended from above Red Bluff in the north to the city of Sacramento in the south. Late-fall run Chinook salmon yearling smolts were released at Battle Creek, Jelly's Ferry, Irvine Finch or Butte City during the winter (Dec-Jan) of each of our study years. The locations of the 20 acoustic receivers that delineated our 19 river reaches are shown as black triangles.


Figure 2: Map depicting reach-specific survival estimates (per 10km) for 2007-2011. Colors represent survival per 10 km for each reach and standard error is represented as the grey buffer surrounding each reach. The values adjacent to each reach are the survival estimates and detection probabilities.


Figure 3. Difference between survival estimates in the spatial-temporal model and the covariate model for each reach (labeled as the distance (River km) between the upstream boundary and the Golden Gate Bridge). Negative values represent occasions when the covariate model had a larger estimate of survival and was presumably missing covariates that increased smolt mortality. Error bars represent the $95 \%$ confidence interval estimated with the delta method.
$1164 \times 846 \mathrm{~mm}(72 \times 72 \mathrm{DPI})$


Figure 4: A barplot depicting the results of covariate importance analysis to determine how removing a single covariate influenced the fit of the selected model. Delta QAIC values represent the change in QAIC when specific variables are removed from the full model.


Figure 5: Covariate response plots showing the effect of the individual covariates on the apparent survival rate through a 10 km reach. The grey shaded region represent the $95 \%$ confidence interval.
$1058 \times 846 \mathrm{~mm}$ ( $72 \times 72 \mathrm{DPI}$ )

Table 1: A description of the covariates included in the mark recapture model.

| Category | Covariate | Range | Definition | Hypothesized relationship with survival |
| :---: | :---: | :---: | :---: | :---: |
| Individual | Fish Length ${ }^{1}$ | 135-204 mm | Fork length | Larger fish may exceed gape width of predators |
|  | Fish Condition ${ }^{1}$ | 0.59-1.32 | Fulton's K | Increased condition improves predator escape capability |
|  | Transit speed ${ }^{2}$ | $0.02-8.25 \mathrm{~km} \mathrm{~h}^{-1}$ | Reach specific transit speed | Faster moving fish have less exposure to predators |
| Release group | Batch release ${ }^{2}$ | Binary | Tagged fish released concurrently with large hatchery releases. | Predator swamping |
|  | Release reach ${ }^{1}$ | Binary | Difference in survival between newly released fish and those released upstream. | Newly released hatchery fish are naïve and susceptible to predation |
|  | Annual flow ${ }^{3}$ | 179-499 cms | Mean flow measured at Bend Bridge throughout outmigration (December-March). | Increased flows produce more habitat and predator refugia throughout the river |
| Reach specific | Sinuosity ${ }^{4}$ | 1.04-2.74 | River distance divided by Euclidean distance. | More natural habitats have more predator refugia |
|  | Diversion density ${ }^{5}$ | 0-1.05 num km ${ }^{-1}$ | Number of diversions per reach length. | Increased predator densities near diversions |
|  | Adjacent cover density ${ }^{6}$ | 0.2-0.76 \% | Percent of non-armored river bank with adjacent natural woody vegetation. | Increased cover produces more predator refugia |
|  | Off-channel habitat density ${ }^{6}$ | 0-1.62\% | Off-channel habitat within 50 m of river expressed as percentage of river area | Increased off-channel habitat produces more predator refugia |
| Time varying | Temperature ${ }^{7}$ | 6.2-12.9 ${ }^{\circ} \mathrm{C}$ | Mean water temperature per reach | Increased temperatures results in increased predation due to higher metabolic demands of predators |
|  | Inter-annual Reach flow ${ }^{7}$ | $215-447 \mathrm{cms}$ | Mean water flow per reach | Higher flows within a reach will produce more habitat and predator refugia within that reach |
|  | Intra-annual <br> Reach flow ${ }^{7}$ | $129-902 \mathrm{cms}$ | Mean water flow per reach and year | Higher intra-annual flows (e.g., precipitation or dam releases) decreases predation due to increased turbidity and increased predator refugia. |

${ }^{1}$ Measured during tagging and release; ${ }^{2}$ Observed travel times and mixed effects model estimates; ${ }^{3}$ California Water Data Library; ${ }^{4}$ National Hydrography Dataset; ${ }^{5}$ Passage Assessment Database - verified by field survey; ${ }^{6}$ Department of Water Resources; ${ }^{7}$ River Assessment for Forecasting Temperature (RAFT) model

Table 2. Beta estimates (standard errors) of covariates included in mark recapture models with a delta QAICc $<2$. The Battle Creek, Sacramento-San Joaquin Delta (Sac-SJ Delta), and San Francisco Bay (SF Bay) covariate are beta estimates for the three reaches where habitat and predation related covariates were not included in the model. See Table 1 for definitions of the other covariates. The selected model is in bold.

| Covariate | $\begin{gathered} \hline \text { Model } \\ 1 \\ \hline \end{gathered}$ | $\begin{gathered} \hline \text { Model } \\ 2 \\ \hline \end{gathered}$ | $\begin{gathered} \hline \text { Model } \\ 3 \\ \hline \end{gathered}$ | $\begin{gathered} \text { Model } \\ 4 \\ \hline \end{gathered}$ | $\begin{gathered} \text { Model } \\ 5 \\ \hline \end{gathered}$ | $\begin{gathered} \hline \text { Model } \\ 6 \\ \hline \end{gathered}$ | $\begin{gathered} \text { Model } \\ 7 \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Intercept | $\begin{gathered} \hline 2.918 \\ (0.050) \end{gathered}$ | $\begin{gathered} \hline 2.900 \\ (0.049) \end{gathered}$ | $\begin{gathered} \hline 2.899 \\ (0.049) \end{gathered}$ | $\begin{gathered} \hline 2.917 \\ (0.050) \end{gathered}$ | $\begin{gathered} \hline 2.943 \\ (0.049) \end{gathered}$ | $\begin{gathered} 2.942 \\ (0.049) \end{gathered}$ | $\begin{gathered} 2.936 \\ (0.049) \end{gathered}$ |
| Battle Creek | $\begin{aligned} & -2.000 \\ & (0.141) \end{aligned}$ | $\begin{gathered} -1.986 \\ \mathbf{( 0 . 1 4 0 )} \end{gathered}$ | $\begin{aligned} & -1.957 \\ & (0.141) \end{aligned}$ | $\begin{gathered} -1.969 \\ (0.142) \end{gathered}$ | $\begin{gathered} -2.023 \\ (0.141) \end{gathered}$ | $\begin{gathered} -1.992 \\ (0.142) \end{gathered}$ | $\begin{aligned} & -2.013 \\ & (0.141) \end{aligned}$ |
| Sac-SJ Delta | $\begin{aligned} & -2.673 \\ & (0.096) \end{aligned}$ | $\begin{aligned} & -2.656 \\ & (0.095) \end{aligned}$ | $\begin{aligned} & -2.659 \\ & (0.095) \end{aligned}$ | $\begin{aligned} & -2.678 \\ & (0.096) \end{aligned}$ | $\begin{aligned} & -2.695 \\ & (0.096) \end{aligned}$ | $\begin{aligned} & -2.698 \\ & (0.096) \end{aligned}$ | $\begin{aligned} & -2.691 \\ & (0.096) \end{aligned}$ |
| SF Bay | $\begin{aligned} & -2.888 \\ & (0.260) \end{aligned}$ | $\begin{aligned} & -2.868 \\ & (0.259) \end{aligned}$ | $\begin{aligned} & -2.899 \\ & (0.261) \end{aligned}$ | $\begin{aligned} & -3.042 \\ & (0.240) \end{aligned}$ | $\begin{aligned} & -2.926 \\ & (0.259) \end{aligned}$ | $\begin{gathered} -2.959 \\ (0.261) \end{gathered}$ | $\begin{aligned} & -2.913 \\ & (0.259) \end{aligned}$ |
| Reach length | $\begin{aligned} & -0.463 \\ & (0.047) \end{aligned}$ | $\begin{aligned} & -0.446 \\ & (0.046) \end{aligned}$ | $\begin{gathered} -0.444 \\ (0.045) \end{gathered}$ | $\begin{gathered} -0.461 \\ (0.047) \end{gathered}$ | $\begin{aligned} & -0.445 \\ & (0.049) \end{aligned}$ | $\begin{aligned} & -0.442 \\ & (0.049) \end{aligned}$ | $\begin{aligned} & -0.457 \\ & (0.049) \end{aligned}$ |
| Sinuosity | $\begin{gathered} 0.168 \\ (0.050) \end{gathered}$ | $\begin{gathered} 0.147 \\ \mathbf{( 0 . 0 4 9 )} \end{gathered}$ | $\begin{gathered} 0.145 \\ (0.049) \end{gathered}$ | $\begin{gathered} 0.167 \\ (0.050) \end{gathered}$ | $\begin{gathered} 0.181 \\ (0.051) \end{gathered}$ | $\begin{gathered} 0.181 \\ (0.051) \end{gathered}$ | $\begin{gathered} 0.188 \\ (0.051) \end{gathered}$ |
| Adjacent cover |  |  |  |  | $\begin{gathered} 0.073 \\ (0.053) \end{gathered}$ | $\begin{gathered} 0.076 \\ (0.053) \end{gathered}$ | $\begin{gathered} 0.089 \\ (0.052) \end{gathered}$ |
| Diversion density | $\begin{gathered} 0.421 \\ (0.057) \end{gathered}$ | $\begin{gathered} 0.382 \\ (0.052) \end{gathered}$ | $\begin{gathered} 0.379 \\ (0.052) \end{gathered}$ | $\begin{gathered} 0.418 \\ (0.057) \end{gathered}$ | $\begin{gathered} 0.423 \\ (0.056) \end{gathered}$ | $\begin{gathered} 0.421 \\ (0.056) \end{gathered}$ | $\begin{gathered} 0.419 \\ (0.056) \end{gathered}$ |
| Off-channel habitat | $\begin{gathered} 0.118 \\ (0.062) \end{gathered}$ |  |  | $\begin{gathered} 0.120 \\ (0.062) \end{gathered}$ | $\begin{gathered} 0.143 \\ (0.065) \end{gathered}$ | $\begin{gathered} 0.147 \\ (0.065) \end{gathered}$ | $\begin{gathered} 0.147 \\ (0.065) \end{gathered}$ |
| Fish condition |  |  | $\begin{gathered} 0.050 \\ (0.030) \end{gathered}$ | $\begin{gathered} 0.054 \\ (0.030) \end{gathered}$ |  | $\begin{gathered} 0.054 \\ (0.030) \end{gathered}$ |  |
| Annual flow | $\begin{gathered} 0.404 \\ (0.039) \end{gathered}$ | $\begin{gathered} 0.406 \\ (0.039) \end{gathered}$ | $\begin{gathered} 0.405 \\ (0.039) \end{gathered}$ | $\begin{gathered} 0.402 \\ (0.039) \end{gathered}$ | $\begin{gathered} 0.387 \\ (0.038) \end{gathered}$ | $\begin{gathered} 0.387 \\ (0.038) \end{gathered}$ | $\begin{gathered} 0.396 \\ (0.038) \end{gathered}$ |
| Reach flow (year) | $\begin{gathered} 0.320 \\ (0.047) \end{gathered}$ | $\begin{gathered} 0.320 \\ (0.047) \end{gathered}$ | $\begin{gathered} 0.315 \\ (0.047) \end{gathered}$ | $\begin{gathered} 0.314 \\ (0.047) \end{gathered}$ | $\begin{gathered} 0.309 \\ (0.047) \end{gathered}$ | $\begin{gathered} 0.304 \\ (0.047) \end{gathered}$ | $\begin{gathered} 0.327 \\ (0.046) \end{gathered}$ |
| Annual flow: Reach flow | $\begin{aligned} & -0.112 \\ & (0.046) \end{aligned}$ | $\begin{aligned} & -0.113 \\ & (0.046) \end{aligned}$ | $\begin{gathered} -0.107 \\ (0.046) \end{gathered}$ | $\begin{aligned} & -0.106 \\ & (0.046) \end{aligned}$ | $\begin{aligned} & -0.115 \\ & (0.046) \end{aligned}$ | $\begin{aligned} & -0.109 \\ & (0.046) \end{aligned}$ | $\begin{aligned} & -0.106 \\ & (0.046) \end{aligned}$ |
| Temperature | $\begin{aligned} & -0.079 \\ & (0.041) \end{aligned}$ | $\begin{aligned} & -0.080 \\ & (0.041) \end{aligned}$ | $\begin{aligned} & -0.078 \\ & (0.041) \end{aligned}$ | $\begin{aligned} & -0.077 \\ & (0.041) \end{aligned}$ |  |  |  |
| Transit speed | $\begin{gathered} 0.079 \\ (0.034) \end{gathered}$ | $\begin{gathered} 0.078 \\ (\mathbf{0 . 0 3 4}) \end{gathered}$ | $\begin{gathered} 0.081 \\ (0.034) \end{gathered}$ | $\begin{gathered} 0.083 \\ (0.035) \end{gathered}$ | $\begin{gathered} 0.069 \\ (0.035) \end{gathered}$ | $\begin{gathered} 0.073 \\ (0.035) \end{gathered}$ |  |
| Release reach | $\begin{aligned} & -0.821 \\ & (0.131) \end{aligned}$ | $\begin{aligned} & -0.857 \\ & (\mathbf{0 . 1 3 0}) \end{aligned}$ | $\begin{aligned} & -0.865 \\ & (0.130) \end{aligned}$ | $\begin{aligned} & -0.829 \\ & (0.131) \end{aligned}$ | $\begin{aligned} & -0.781 \\ & (0.135) \end{aligned}$ | $\begin{aligned} & -0.787 \\ & (0.135) \end{aligned}$ | $\begin{aligned} & -0.781 \\ & (0.135) \end{aligned}$ |
| Batch release | $\begin{gathered} 0.694 \\ (0.147) \end{gathered}$ | $\begin{gathered} 0.701 \\ (0.146) \end{gathered}$ | $\begin{gathered} 0.689 \\ (0.147) \end{gathered}$ | $\begin{gathered} 0.679 \\ (0.147) \end{gathered}$ | $\begin{gathered} 0.637 \\ (0.143) \end{gathered}$ | $\begin{gathered} 0.625 \\ (0.143) \end{gathered}$ | $\begin{gathered} 0.651 \\ (0.143) \end{gathered}$ |
| Batch release: Distance | $\begin{gathered} -0.003 \\ (0.000) \\ \hline \end{gathered}$ | $\begin{gathered} \mathbf{- 0 . 0 0 3} \\ \mathbf{( 0 . 0 0 0 )} \\ \hline \end{gathered}$ | $\begin{gathered} -0.003 \\ (0.000) \\ \hline \end{gathered}$ | $\begin{array}{r} -0.003 \\ (0.000) \\ \hline \end{array}$ | $\begin{aligned} & -0.003 \\ & (0.000) \\ & \hline \end{aligned}$ | $\begin{gathered} -0.003 \\ (0.000) \\ \hline \end{gathered}$ | $\begin{array}{r} -0.003 \\ (0.000) \\ \hline \end{array}$ |
| Survival covariates | 16 | 15 | 16 | 17 | 16 | 17 | 15 |
| Delta QAICc | 0 | 0.29 | 0.71 | 1.20 | 1.27 | 1.38 | 1.63 |

# Decoupling outmigration from marine survival indicates outsized influence of streamflow on cohort success for California's Chinook salmon populations 

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#### Abstract

Historically, marine survival estimates for salmon have been confounded with freshwater seaward migration ("outmigration") survival. Telemetry studies have revealed low and variable survival during outmigration, suggesting marine mortality may not be the primary source of variability in cohort size as previously believed. Using a novel combination of tagging technologies, survival during these two life stages was decoupled over five years for Sacramento River Chinook salmon. Outmigration survival ranged from $2.6 \%$ to $17 \%$, marine survival ranged from $4.2 \%$ to $22.8 \%$. Influential environmental drivers in both life stages were also compared to smolt-to-adult ratios (SAR) for three Chinook salmon populations over 20 years. Streamflow during outmigration had higher correlation with SAR ( r -squared $>0.34$ ) than two marine productivity indices (r-squared $<0.08$ ). The few SAR estimates that were poorly predicted by flow occurred during years with the lowest marine productivity, suggesting most inter-annual SAR fluctuations are explained by outmigration survival, but abnormally poor marine conditions also reduce SAR. The outsized influence of flow on SAR provides managers with a powerful mitigation tool in a watershed where flow is tightly regulated.


## Keywords

## -Chinook salmon

-Survival

## -California

## -River regulation

## -Marine productivity

## -Smolt-to-adult

## -Acoustic telemetry

## Introduction

Convention is that variability in salmon cohort success is set during the early marine residence period. To date, direct evidence of how outmigration (freshwater plus estuarine) survival might be affecting overall cohort success has been scarce throughout the range of salmon populations. Historically, it has been difficult to parse out outmigration survival from marine survival, further obfuscating the causes and magnitude of outmigration mortality. Recent telemetry studies have estimated very low survival during the outmigration life stage of certain salmon stocks (Buchanan et al. 2013; Michel et al. 2015; Clark et al. 2016), suggesting that marine survival is likely higher than what the literature indicates. Many models attempting to explain marine survival using marine environmental indicators suffer from large amounts of unexplained variation in some years (Koslow et al. 2002; Logerwell et al. 2003; Sharma et al. 2013); and there is potential that variation due to outmigration survival has been incorrectly attributed to marine survival in these models. Through the accurate partitioning of outmigration and marine survival, it may be possible to identify new survival bottlenecks which will require new and different management solutions.

Marine conditions are often blamed for poor cohort success of California's Central Valley Chinook salmon (Oncorhynchus tshawytscha) populations, but there is a building body of evidence to suggest that outmigration survival may be playing a large role (Buchanan et al. 2013; Michel et al. 2015). Gross et al. (1988) posited that anadromous life history strategies evolve in fishes when migration to the ocean provides gains to individual fitness that outweigh the costs of the migration itself. It is believed that salmon have evolved this life history strategy because the ocean provides a more favorable tradeoff between abundant food and predation risk. However, the Central Valley may be an example of a system where the costs of outmigration are high
enough that the anadromous life history strategy is no longer sustainable, and is only persisting through the assistance of humans (such as through hatcheries, or transporting outmigrants past regions of poor survival). Three of the four distinct salmonid Evolutionarily Significant Units (ESUs) that are found there are listed under the U.S. Endangered Species Act (ESA), and the fourth is a "species of concern". Many inland stressors have been identified that have led to the decline of these populations, including the loss of $47 \%$ of spawning and rearing habitat due to dams without fish passage (Yoshiyama et al. 2001) and $97 \%$ of the productive floodplain rearing habitat to diking (Whipple et al. 2012). These dams and levees are one-time historical perturbations, but have ongoing impacts and will likely never be completely reversed. While it is almost certain that populations will not return to pre-dam and pre-diking levels without reversing these habitat changes, studies must also concentrate on the contemporary stressors that are governing annual outmigration survival dynamics, such as warm stream and estuary temperatures during outmigration, slow water velocities, low turbidity, and abundant predators (Baker et al. 1995; Newman and Rice 2002; Grossman 2016). However, these are just the symptoms of a larger problem: the fundamental alteration of the Central Valley hydrological regime. The dams and diversions of the Central Valley have resulted in the reduction and homogenization of river flows (Buer et al. 1989), which in turn can alter water temperatures, slow water velocities associated with large flow events, lower turbidity and provide more suitable habitat for warm-water predator species. These same dams and diversions give resource managers tight control over streamflow and associated covariates. In contrast, managers have no control over the environmental variables that are thought to govern marine survival. Therefore, if outmigration survival is found to have a large influence on the magnitude and variability in cohort success, this suggests that managers can likely do more to help these populations.

A novel method of pairing outmigration survival estimates derived from an acoustic tagging study with smolt-to-adult ratio (SAR) estimates derived from coded-wire tag (CWT) recoveries from the same cohorts was used to investigate the relative importance of (1) freshwater and estuarine outmigration (hereafter simply termed "outmigration") survival versus (2) marine survival rates for Central Valley Chinook salmon over the 5-year time series of the acoustic tagging study. Expanding beyond this time series, many additional years of SAR estimates were regressed against environmental drivers that are believed to be influential on survival in each region to investigate the importance of these environmental drivers on smolt-toadult dynamics and ultimately gain insights on where the majority of mortality might be occurring every year.

## Methods

## Study system

California's Central Valley includes the two largest rivers in the state. In the northern portion of the valley, the Sacramento River flows north to south and in the southern portion of the valley, the San Joaquin River flows south to north (Fig. 1). These two rivers meet to create the freshwater portion of their shared estuary: the Sacramento-San Joaquin River Delta (hereafter "Delta"), an expansive and complex network of tidal freshwater river channels and sloughs. It is connected to the west by a series of increasingly saline bays, most notably the San Francisco Bay, which comprise the brackish portion of the estuary ("Bays" in Fig. 1). The estuary connects to the Pacific Ocean at the narrow passage at the Golden Gate, beyond which salmon have access to the productive waters of the Gulf of the Farallones.

## Outmigration survival estimates

In an attempt to decouple outmigration and marine survival of Central Valley Chinook salmon, cohorts that were tagged using both acoustic tags (for estimation of outmigration survival) and coded-wire tags ("CWT"; for estimation of overall cohort success) were identified. Outmigration survival estimates were used from two acoustic tagging studies conducted on hatchery-origin late-fall-run Chinook salmon from 2007 to 2011 (Michel et al. 2015; Iglesias et al. 2017). These studies released their acoustic tagged fish as part of larger hatchery releases that were also coded-wire tagged. CWTs are tiny, injectable, magnetized wire segments that are embossed with a release group serial code, with release groups of thousands of fish often sharing the same serial code. Recovery of tagged adults allows the estimation of smolt-to-adult ratio (SAR) of these larger release groups. SAR represents the proportion of fish of a harvestable size recovered from the total number of juveniles released into the wild and was therefore used as an index of cohort success.

To assess the contribution of outmigration survival to overall SAR, and to factor out estimates of marine survival, outmigration survival from acoustic tagged release groups were associated to the SAR estimates from the most appropriate CWT release groups. However, some of the acoustic tagged release groups were not released in exact synchrony with a respective CWT release group. For these, if one or more CWT release groups were released within 7 days of the acoustic tag group's release date, that acoustic tag group's outmigration survival was associated to the respective CWT release group(s). For the purposes of these studies, outmigration survival was estimated as total survival from release to the Golden Gate Bridge, thereby including river and estuarine survival. For more information on the acoustic tagging, tracking, and estimation of survival for the acoustic tagging studies, refer to Michel et al. (2015).

## Smolt-to-Adult estimates

SAR is a survival metric often used for hatchery fish because of the fairly accurate estimates of how many smolts are released. Hatchery Chinook salmon are often raised up to the smolting stage before release, which is the beginning of the SAR period. The end of the SAR period is when a fish either returns to the spawning grounds or hatchery, or is captured by commercial or recreational fisheries. These various recapture scenarios ("strata"), and their associated CWT recoveries, occur after Chinook salmon have spent at least one year in the ocean ( $2+$ year old), and can commonly occur for salmon that have spent as many as 3 years in the ocean (4+ year old; Fig. 2). SAR therefore represents the survival of a cohort from smolting to the point at which they reach harvestable and minimum reproductive (i.e. "adult") size. Thus, survival during the SAR period for a CWT group will be the product of 1) "outmigration survival" $\left(S_{O}\right)$ and 2) "marine survival" $\left(S_{M}\right)$, survival during the first year at sea plus an amalgamation of year 2, 3, and 4 survival depending on recapture time of individuals within the CWT group. Due to this complexity, SAR should be treated as an index of survival that primarily represents survival from hatchery release to age 2 , with some additional mortality from latter periods (but that are thought to be relatively small contributions compared to critical survival bottlenecks of outmigration and the first year at sea [Magnusson and Hilborn 2003; Quinn 2005 and references therein]).

The SAR in the Central Valley is most often calculated using CWT recoveries $\left(C W T_{R}\right)$. Approximately $25 \%$ of all hatchery-origin fall-run Chinook salmon (since 2007) and $100 \%$ of all hatchery-origin late-fall-run and winter-run Chinook salmon (since 1992) in the Central Valley have CWTs inserted into their snouts as juveniles. Once the salmon attain harvestable size (hereafter "adults"), the CWTs are recovered from the fisheries through creel surveys, from the spawning grounds through carcass surveys, and through the hatcheries (for additional details on
recovery sources, refer to Table 1). All CWT data were downloaded from the Pacific States Marine Fisheries Commission's Regional Mark Processing Center's Regional Mark Information System database (http://www.rmpc.org/).

The first brood year (i.e., the year the eggs were spawned; "BY" hereafter) for which SAR could be accurately estimated was 1999 for both winter and fall-run Chinook salmon, and 1993 for late-fall-run Chinook salmon (despite the absence of spawning ground and recreational river fishery recoveries until the late 1990s). Since an estimated 61 to $97 \%$ (mean $80 \%$ ) of late-fall-run Chinook salmon escapement are counted at hatcheries (using CWT data from recovery years 2000-2016 when spawning ground and recreational river fishery recoveries occured), using only hatchery returns in years prior to the late 1990s could bias SAR estimates low for those years, but would likely still capture the major population trends.

For creel and carcass surveys, full coverage of all fishing areas and spawning grounds is not possible; sampling fractions $(r)$ are therefore estimated per stratum (i.e., unique recovery type, area and year combinations). Sampling fractions are the fraction of estimated total number of salmon caught (if a fishery) or that returned (if a hatchery or spawning area) that were examined for presence of a CWT per stratum, with some additional nuances outlined in PalmerZwahlen and Kormos (2015). Details on how total number of salmon per stratum were estimated can be found in O'Farrell et al. (2012). Expansion factors, the reciprocal of sampling fractions, are applied to the total CWTs observed per CWT release group that are recovered from that respective stratum to produce expanded CWT recoveries $\left(e C W T_{R}\right)$. Finally, since Chinook salmon spawning age is variable (minimum age 2 years), SAR for the full cohort cannot be estimated until the CWTs from the fifth year after release are processed. Thus, SAR estimates
beyond BY 2012 are not reported. Total expanded recoveries for each release group $\left(N_{e}\right)$ is therefore estimated as:

$$
\begin{equation*}
N_{e}=\sum_{y=1}^{Y}\left[\mathrm{eCWT}_{R} \text { Ocean Fishery }+\mathrm{eCWT}_{R} \text { River Fishery }+\mathrm{eCWT}_{R} \text { Spawning Grounds }+\mathrm{CWT}_{R} \text { Hatchery }\right] \tag{1}
\end{equation*}
$$

where $Y$ is total number of return years for which CWTs are observed for that CWT release group. Note that hatchery CWT recoveries are not expanded because all CWTs are presumed to be recovered from hatchery returns.

SAR is expressed as the proportion of expanded recoveries $\left(N_{e}\right)$ out of all smolts released from the hatchery for that CWT release group $\left(N_{r}\right)$ :
$S A R=\frac{N_{e}}{N_{r}}$
The standard error (SE) of the SAR for a CWT release group is a function of $N_{e}, N_{r}$, and the total number of observed CWTs (before expansion, $N_{d}$ ) (Skalski and Townsend 2005):
$S E(S A R)=\sqrt{\frac{N_{e}}{N_{r}}\left(1-\frac{N_{e}}{N_{r}}\right)} \frac{N_{r}}{N_{r}}+\frac{\left(\frac{1-r}{r^{2}}\right) N_{d}}{N_{r}^{2}}$
For proper variance calculation, sampling fractions are needed per stratum. However, protocols for estimating sampling fractions differed substantially by year and recapture type. Overall, the sampling fraction for all CWTs recovered (across the strata) per brood year and per population in this analysis was never below 0.21 , and the mean was 0.35 for winter-run, 0.49 for fall-run, and 0.63 for late-fall-run. Therefore, a global sampling fraction $(r)$ was applied to equation 3 using a conservative estimate of 0.2 :
$S E(S A R)=\sqrt{\left.\frac{\frac{N_{e}}{N_{r}}}{\frac{N_{r}}{}} 1-\frac{N_{e}}{N_{r}}\right)}+\frac{\left(\frac{1-0.2}{0.2^{2}}\right) N_{d}}{N_{r}{ }^{2}}$

When calculating SAR and standard error for more CWT release groups that were released on the same day, $N_{e}, N_{r}$, and $N_{d}$ were totaled among those CWT release groups. However, because there can be large heterogeneity in SAR estimates for different CWT release groups released in the same year, annual SAR and standard errors are calculated differently (Skalski and Townsend 2005). Annual SAR is a weighted average across CWT release groups:
$\widehat{S A R}=\frac{\sum_{k=1}^{K} N_{e_{k}}}{\sum_{k=1}^{K} N_{r_{k}}}$
Where K is the number of CWT release groups in a year. Standard error of the annual SAR is estimated as:
$S E(\widehat{S A R})=\sqrt{\frac{\sum_{k=1}^{K} N_{r_{k}}\left(S A R_{k}-S A R\right)^{2}}{(K-1) \sum_{k=1}^{K} N_{r_{k}}}}$
For the late-fall-run and winter-run populations, the only hatcheries that release smolts in the Central Valley are the United States Fish and Wildlife Service's (USWFS) Coleman National Fish Hatchery (CNFH) and Livingston Stone National Fish Hatchery (LSH), respectively. Both of these hatcheries release the majority of their fish into the uppermost portions of the Sacramento River that is available to anadromy, more than 500 river km from the Pacific Ocean. Because multiple hatcheries in the Central Valley release fall-run smolts, to compare fall-run release groups over the same outmigration corridor as the late-fall-run and winter-run, only fallrun CWT recoveries from CNFH release groups were used. All CWT release groups that were trucked and released downstream, a management strategy intended to artificially increase SARs (by reducing outmigration mortality) of hatchery smolts, were also excluded. This is because one of the main objectives of this study was to explicitly measure the magnitude and variability in natural outmigration survival.

SAR estimates are the combination of survival over a finite outmigration period and nondiscrete marine period (due to various CWT recapture times). To ascertain the magnitude of the bias introduced by the latter periods of the non-discrete marine period on overall SAR, SAR estimates were compared to survival rates from hatchery release to the end of age 2 for winterrun Chinook salmon for the same brood years, as estimated from a Sacramento River winter-run Chinook salmon cohort reconstruction model ([O'Farrell et al. 2012]; data provided by M. O'Farrell, NOAA-NMFS). This was done using a linear regression model fitted between the two variables, after logit-transformation (due to the range of both variables being bound by 0 and 1 ). Currently, a salmon cohort reconstruction model does not exist for Central Valley fall or late-fall-run Chinook salmon.

## Outmigration vs. Marine Survival comparison

The outmigration survival component of SAR, as estimated from acoustic telemetry, was factored out to get an estimate of marine survival for those brood years:
$S_{M}=\frac{S A R}{S_{o}}$
To incorporate error in estimates of both $S A R$ and $S_{O}$, parametric bootstrapping was employed. $S A R$ was assumed to have a normal distribution on the real scale and $S_{O}$ was assumed to have a normal distribution on the logit scale. Given these distributions, $S A R$ and $S_{O}$ were generated 1000 times each and transformed back to the real scale, such that $\left(S A R_{1}^{*}, S A R_{2}^{*}, \ldots\right.$, $\left.S A R_{1000}^{*}\right)$ and $\left(S_{O_{1}}^{*}, S_{O_{2}}^{*}, \ldots, S_{O 1000}^{*}\right)$ yielded $S_{M 1}{ }_{1}^{*}, S_{M 2}{ }^{*}, \ldots, S_{M 1000}^{*}$. Mean $S_{M}$ and standard error of the mean were estimated from these values on the logit scale and back transformed to the real scale. The 95\% confidence intervals were also generated given:
$\operatorname{logit}^{-1}\left[\operatorname{logit}\left(\widehat{S_{M}}\right) \pm 1.96 \times \operatorname{SE}\left[\operatorname{logit}\left(\hat{S_{M}}\right)\right]\right]$

This was done for late-fall-run Chinook salmon only, and not for fall-run or winter-run Chinook salmon due to the lack of acoustic tag data old enough to estimate respective SAR values.

## Freshwater outmigration survival vs. SAR

Michel et al. (2015) demonstrated that much of the annual variability in outmigration survival may be occurring during the freshwater portions of the outmigration. To evaluate the effect of annual freshwater outmigration survival $\left(S_{F W}\right)$ dynamics on SAR, a linear model was fitted to survival rates estimated from acoustic tags and the CWT-based SAR. The acoustic tagestimated survival rates encompassed the river and Delta regions combined (i.e., from release to Chipps Island; data from Michel et al. [2015]).

In order to incorporate error, parametric bootstrapping was employed for both $S A R$ and $S_{F W}$. SAR data was generated 1000 times on the real scale, then transformed to the logit scale due to $S A R$ being bounded by 0 and 1 , such that $\left[\operatorname{logit}\left(S A R_{1}^{*}\right), \operatorname{logit}\left(S A R_{2}^{*}\right), \ldots, \operatorname{logit}\left(S A R_{1000}^{*}\right.\right.$ )] datasets were created. $S_{F W}$ was generated 1000 times on the logit scale, again because $S_{F W}$ is bounded by 0 and 1 , such that $\left[\operatorname{logit}\left(S_{F W}{ }_{1}^{*}\right), \operatorname{logit}\left(S_{F W}{ }_{2}^{*}\right), \ldots, \operatorname{logit}\left(S_{F W}{ }_{1000}^{*}\right)\right]$ datasets were created. The $S A R$ datasets were fitted to their respective $S_{F W}$ datasets per iteration of 1000 different linear models, such that 1000 estimates of $r$-squared values were generated. The median, $5 \%$ and $95 \%$ percentile values (i.e., $95 \%$ confidence intervals) of the r-squared estimates were then calculated.

## Environmental covariates vs. SAR

The relationship between SAR and variables that characterize the river and ocean environments were evaluated for each of the three Chinook salmon populations. Linear regression models were fitted between logit-transformed SAR estimates and environmental
indices. Because extreme outliers can mask strong and persistent trends, Cook's distances were estimated for all points in all models (Cook 1977) to determine if any annual SAR values exert excessive leverage on the linear regressions. The linear regression model was fitted with and without any annual SAR value with a Cook's distance $>1$.

Environmental covariates thought to influence survival during the outmigration and marine survival life stages were selected in an attempt to determine the relative contribution of these factors on cohort success. For the river environment, the literature suggests that flow may have the greatest influence on outmigration survival (Newman and Rice 2002; Smith et al. 2003; Michel et al. 2015). Flow values (cubic feet per second) were used from the United States Geological Survey's Bend Bridge gauging station on the Sacramento River (USGS station number 11377100). This gauge is located approximately 20 and 60 river kilometers downstream from the release locations used by the CNFH and LSFH, respectively. Distribution of flow values were right-skewed, and thus log-transformed for normality.

A single variable (upwelling) and a multivariate index of productivity were chosen for the marine environment. Upwelling is a key variable in determining the quality of marine conditions for salmon (Kope and Botsford 1990; Scheuerell and Williams 2005; Wells et al. 2016). Mean monthly coastal upwelling index as computed by the National Oceanic and Atmospheric Administration's National Marine Fisheries Service for the $39^{\circ} \mathrm{N} 125^{\circ} \mathrm{W}$ station, the closest station to the Gulf of the Farallones
(https://www.pfeg.noaa.gov/products/PFEL/modeled/indices/upwelling/upwelling.html) was used as the single covariate. The upwelling index represents wind-driven cross-shore transports computed from surface pressure analyses (in cubic meters per second along each 100 meters of coastline). The Multivariate Ocean Climate Indicator (MOCI) as described in Garcia-Reyes and

Sydeman (2017) was used as the multivariate index of productivity. This unitless environmental indicator, specific to California's continental shelf, synthesizes numerous ocean and atmospheric variables to give an index of the state of the ecosystem productivity (http://www.faralloninstitute.org/moci). The MOCI is estimated for both the Northern California region ( 38 to $42^{\circ} \mathrm{N}$ latitude) and the Central California region ( 34.5 to $38^{\circ} \mathrm{N}$ latitude). Since juvenile salmon from the Central Valley are known to occupy both these regions (MacFarlane 2010), the mean seasonal MOCI between these regions was used. Low MOCI values represent high marine productivity, and high MOCI values represent low marine productivity.

Daily mean flow at Bend Bridge was averaged over a 14-day window, starting the day of release, for each CWT release group, to represent the mean river travel time from release to Delta entry (as estimated for acoustic tagged hatchery-origin late-fall-run Chinook salmon smolts [Michel et al. 2012]). These release group-specific 14-day mean flows were then averaged per year and weighted to the size of each CWT release group. For the marine environment, the first few months at sea is the most critical survival period of the marine phase of a salmon's life history (Kilduff et al, 2014), specifically during the first spring at sea for Central Valley salmon stocks and mediated through environmental drivers such as upwelling (Wells et al. 2012; Woodson et al. 2013). Therefore, the mean monthly upwelling index across the months of March, April, and May for the year of outmigration were used, as well as the mean of the Northern and Central California spring MOCI.

The residuals of the flow linear models were graphically compared to upwelling and MOCI to evaluate if any variability in SAR that was unexplained by flow could be explained by the marine environmental covariates. Two contour plots were generated by interpolating the known SAR values (all three salmon populations combined, to increase resolution) across a grid
of flow and either upwelling or MOCI values (using Akima interpolation [Akima 1970]), bounded by the limits of the current dataset. Because SAR values could be influenced by population-specific life history strategies, annual logit-scale SAR values were standardized within populations (i.e., z-score: subtracted the mean and divided by the standard deviation for each SAR value), and then combined. All analyses were performed using program R (version 3.5.1, R Core Team 2016) along with the "akima" package (Akima and Gebhardt 2016).

## Results

## Smolt-to-Adult estimates

Annual SAR values were estimated for 20 consecutive years for late-fall-run, and 14 consecutive years for winter-run and fall-run Chinook. The number of CWTs released per run and per year ranged from 30,451 to $3,128,686$. Annual SAR ranged from $0.02 \%$ to $3.29 \%$ overall, and mean annual SAR for these years were $1.00 \%(0.1 \mathrm{SE})$ for late-fall-run Chinook salmon, $0.64 \%(0.18 \mathrm{SE})$ for winter-run Chinook salmon, and $0.81 \%(0.26 \mathrm{SE})$ for fall-run Chinook salmon (Table 2).

There was a strong positive relationship between the winter-run Chinook salmon SAR values and hatchery release to end of age- 2 survival, as estimated by cohort reconstruction (rsquared 0.95 ; Fig. 3). Because the two variables are approximately equal under the same conditions ( $95 \%$ confidence intervals of the linear model between these two variables overlap the 1:1 line), SAR was used to represent the combined outmigration and marine survival during the first year at sea.

## Outmigration vs. Marine Survival comparison

Overall, outmigration survival ranged from $2.6 \%$ to $17 \%$, and marine survival ranged from $4.2 \%$ to $19 \%$ for eight late-fall-run Chinook salmon CWT release groups (or cluster of
release groups) from brood years 2007 through 2010 (Fig. 4). For the eight CWT release groups, five were estimated to have higher marine survival than the respective outmigration survival estimate, two groups had the opposite pattern, and one group had approximately equal survival in both periods. SAR estimates were distributed above and below the BY 1993-2012 long-term median SAR ( $0.81 \%$; represented by the black dashed line in Fig. 4), suggesting that these release groups experienced overall survival that was roughly representative of the larger pool of CWT release group SAR estimates.

## Freshwater outmigration survival vs. SAR

Freshwater survival had a strong positive relationship with overall SAR for these same eight CWT release group clusters (r-squared 0.62; Fig. 5), indicating freshwater outmigration survival was an important factor in overall SAR for those cohorts.

## Environmental covariates vs. SAR

Flow during outmigration was a strong predictor of SAR in all three of the Chinook salmon runs (r-squared 0.45 for late-fall-run, 0.57 for winter-run, and 0.35 for fall-run Chinook salmon, after removing the extreme outliers identified by Cook's distance), while both upwelling and MOCI during the first spring at sea had little influence over SAR (Fig. 6). All points in all linear models had Cook's distances $<1$ with the exception of 20.0 and 1.9 for outmigration year (i.e., brood year +1 ; "OY" hereafter) 2006 in both the fall-run and winter-run Chinook salmon linear models between SAR and flow (red labeled points in Fig. 6. d and g). The r-squared of the linear regressions with the outlier included was 0.08 for fall-run and 0.16 for winter-run (linear regressions shown in Fig. 6. d and g do not include the OY 2006 year). In both cases, these outliers had lower SAR than what would be predicted by flow during outmigration given the remainder of the datasets.

The residuals from the three flow regressions were plotted against spring upwelling index, and spring MOCI. For fall-run and winter-run Chinook salmon OY 2006, the residual was predicted based on the linear regression that was fitted to the dataset that did not include OY 2006 (due to having a Cook's distance > 1). Model performance was poorest in predicting annual SAR in years with some of the lowest upwelling and MOCI indices (Fig. 7). Specifically, for late-fall-run Chinook salmon, model performance was poor in OYs 1998 and 2005; years with the lowest spring upwelling indices and the highest MOCI indices (i.e., low productivity) from the 20-year time series. For winter-run Chinook salmon, the flow model performed poorly in explaining the low SAR that occurred for salmon outmigrating during OY 2005 and 2006; these same years also had the first and third lowest spring upwelling index values and the highest MOCI index values for the 14 -year time series. For fall-run Chinook salmon, the model poorly explained the low SAR for outmigrating salmon in OY 2006; the year with the third lowest spring upwelling index and the second highest MOCI index for the 14 -year time series.

For all three runs, flow was the primary driver of year-to-year variation in SAR for the variables tested (Fig. 6), with marine productivity only playing a major role in annual dynamics when productivity was at low levels (Fig. 7). High SAR values tended to only occur when flow was higher than average and productivity was not near abnormally low levels (Fig. 8). The OY 2014-2017 cohorts (for which SAR values are not yet available) are predicted to have poor SAR based on the trends seen in the existing data with the exception of the OY 2015 late-fall Chinook salmon and all three runs in OY 2017 as predicted by the upwelling contour plot (Fig. 8). The MOCI contour plot has all three runs in OY 2017 falling outside the bounds of the contour plot.

## Discussion

This study indicates that outmigration survival, and the conditions that affect it, are the primary drivers of SAR dynamics, and marine survival likely only plays a critical role in years with abnormally unfavorable marine conditions for salmon. Lindley et al. (2009) also suggested that ocean conditions can have infrequent and yet drastic effects on salmon cohorts, while the long-term, steady degradation of the freshwater environment likely plays a larger role in population health of Central Valley Chinook salmon populations. In a sense, these populations are extremely stressed due to the degraded freshwater environment, and cumulative to this, poor marine conditions can then result in extremely low survival rates.

This study used a novel combination of short-term acoustic tagging data paired with long-term coded-wire tag recovery data to estimate marine survival rates for California Chinook salmon populations. The results indicated that marine survival for California Chinook salmon populations is similar in scale to outmigration survival. Given that these marine survival estimates are confounded with return river survival, net marine survival is likely higher than outmigration survival in most years. Two studies have found exceptionally low outmigration survival rates for California Central Valley Chinook salmon stocks compared to other large West Coast rivers (Buchanan et al. 2013; Michel et al. 2015). Given these low outmigration survival rates, it would be mathematically impossible for these fished populations to be sustainable if marine survival was much lower than outmigration survival and hatchery propagation did not exist (Michel et al. 2015). Indeed, the average annual SAR estimates in this study were below $1 \%$ for all three populations; for Upper Columbia and Snake River Chinook salmon populations, the Columbia River Basin Fish and Wildlife Program suggests that a minimum of $2 \%$ SAR is required for population survival and $4 \%$ for population recovery (NPCC 2009). This study is an additional line of evidence suggesting that for California Central Valley Chinook salmon
populations, the risks of outmigration may now be too high and these populations are likely no longer sustainable.

That the contribution of marine survival to cohort success has been overestimated over the past decades of salmon research is an emerging concept, and one that is not unique to California or Chinook salmon. It has been suggested for Atlantic salmon (Salmo salar) in the Bay of Fundy, Canada (Lacroix 2008), for steelhead (Oncorhynchus mykiss) in the Cheakamus River, British Columbia (Melnychuk et al. 2014), and for sockeye salmon (Oncorhynchus nerka) in the Fraser River, British Columbia (Clark et al. 2016). The emergence of this concept is fundamentally linked to the advent of acoustic tags small enough for tagging juvenile salmon; because accurate estimates of outmigration survival before acoustic tags was difficult if not impossible. Without an estimate of outmigration survival, outmigration survival and marine survival cannot be parsed, which may lead researchers to believe that marine survival was driving population declines. Potential factors leading to this misconception include the fact that less is known about marine survival dynamics, marine residency is substantially longer in duration than the outmigration period, and recruitment is set during early marine residence for many strictly marine fishes and this concept was transferred to salmon. Managers and biologists should ensure that salmon life-cycle and forecast models incorporate some index of outmigration survival.

Streamflow during outmigration was found to have a large influence on SAR dynamics. Over $35 \%$ of all variability in annual SAR dynamics can be explained by flow during outmigration for three different Chinook salmon populations (after removal of an extreme outlier). Flow has been found by numerous studies to have strong influences on outmigration survival of salmon populations worldwide, including Central Valley Chinook salmon
populations (Kjelson and Brandes 1989; Zeug et al. 2014). Increases in flow usually cause or are coincident with changes in many other river conditions that are beneficial to the survival of outmigrating salmon, such as increased water velocities (Hogasen 1998), decreased water temperatures (Smith et al. 2003), increased turbidity (Gregory and Levings 1998), and increases in habitat area that reduce exposure to predators and increase growth opportunities (Sommer et al. 2001). Among existing studies, this is one of only a few studies have demonstrated that flow can ultimately have a strong influence on overall cohort success in the Central Valley (Sturrock et al. 2015; Wells et al. 2017).

These results demonstrate that marine survival is also a major contributor to overall cohort strength. While the indices used for marine productivity in this analysis did not show strong relationships with SAR, this is not evidence of a lack of influence of marine survival on SAR variability, as they cannot capture all the relevant factors (e.g., abundance of predators, alternative prey, etc.). Moreover, the magnitude of marine survival was found to be as large a contributor to SAR as outmigration survival. Furthermore, three of the study years showed evidence of poor marine productivity leading to low SAR, all of which were corroborated with existing literature. The first of these three years, 1998, was a record El Nino-Southern Oscillation (ENSO) event with drastic effects on the California marine ecosystem (Lynn et al. 1998), which likely had a strong negative impact on marine survival of salmon (Pearcy and Schoener 1987; Johnson 1988). In 2005, during the well-documented delayed spring upwelling and resulting poor productivity of the northern California Current (Schwing et al. 2006; Barth et al. 2007), there was evidence of strong size and growth-rate selective early-marine mortality of Central Valley Chinook salmon (Woodson et al. 2013). In 2006, spring upwelling was similarly delayed as in 2005, especially off the coast of Central California where juvenile Central Valley Chinook
salmon first recruit to after leaving the San Francisco Bay, leading to a similar situation of poor productivity (Lindley et al. [2009] and references therein). It is widely accepted that the poor early-marine survival of Central Valley fall-run Chinook salmon in the springs of 2005 and 2006 were the proximate causes of the collapse of that stock and the temporary closure of the fishery (Lindley et al. 2009), and in this analysis, the otherwise strong positive relationship between flow and SAR for fall-run and winter-run Chinook salmon was likely overshadowed by abnormally poor early-marine survival in OY 2006, as demonstrated by the high Cook's distances of those points.

These results also provide insights into how river and marine conditions might have varied influences on different salmon populations. High flows during outmigration benefited all three populations, despite the juveniles leaving at different sizes and at different times of the year. However, marine productivity seems to have affected the different runs differently in some years. For example, the late-fall-run Chinook salmon did not experience the OY 2006 crash, while the winter-run and fall-run did, despite all three benefitting from relatively high flows during outmigration. This could be due to the late-fall-run's predisposition to a larger size at ocean entry, especially if size-selective mortality is at play (which is often seen during poor ocean conditions [Holtby et al. 1990; Saloniemi et al. 2004; Woodson et al. 2013]). Lindley et al. (2009) reported on this discrepancy between the fall-run and late-fall-run Chinook salmon in those years: "Curiously, Sacramento River late-fall-run Chinook salmon escapement has declined only modestly since 2002, while the [Sacramento River fall-run] in the same river basin fell to record low levels." This is strong support for the concept of allowing Central Valley salmon to exhibit many life-history strategies and thereby diversifying the Central Valley
salmon's portfolio and increasing population stability (Schindler et al. 2010; Carlson and Satterthwaite 2011).

As with many large-scale correlative survival studies, there are noteworthy caveats. Firstly, the survival estimates used in this analysis are for hatchery-origin fish only. While the trends discovered in this analysis likely effect wild populations similarly, empirical estimates of SAR for wild Central Valley Chinook salmon do not currently exist. Secondly, the effects of acoustic tagging on juvenile salmon can bias survival estimates low, through mortality related to the tag or surgery, mortality due to behavioral changes, or tag shedding. A subset of the fish used to generate the acoustic tag survival estimates used here from Michel et al. (2015) were also submitted to a laboratory tag effects study. In that study, no fish shed their tags over 160 days (exceeding the maximum outmigration time) and tagged fish growth and survival was not significantly different than untagged fish (Ammann et al. 2013). However, no tests were conducted to address mortality related to behavioral changes, and therefore it is conceivable that outmigration survival estimates used in this study were biased low. Thirdly, the strong relationship between flow during outmigration and SAR may be mediated in some part through marine survival. Climatic dynamics that led to increases or decreases in precipitation over the inland portions of the salmon's range may have also influenced marine conditions in a manner not captured by the marine productivity indices, but had an influence on SAR nonetheless. A similar scenario was demonstrated by Lawson et al. (2004) with coho salmon populations in the Pacific Northwest. One potential avenue for a post-hoc investigation of this concept would be to look for correlation between flow during outmigration and the marine productivity indices. Using the combined datasets, the r-squared for a linear model between flow during outmigration and spring upwelling was 0.07 , and 0.19 between flow and spring MOCI, showing some
evidence of relationships between these freshwater and marine indices. These relationships are likely driven by the trend that years with extremely high flows typically have low spring productivity (see conspicuous lack of points in upper-right quadrant of figure 8 a and lower-right quadrant of figure 8 b ). This phenomenon may be in part explained by the effects of ENSO, which often manifests itself in California with heavy precipitation and low productivity of coastal waters (Schonher and Nicholson 1989; Jacox et al. 2015). In the one year that contradicted this trend in this dataset, OY 2005, when flow during outmigration and ocean productivity were both extremely low, SAR values were at their lowest levels ( $1^{\text {st }}$ lowest for late-fall-run, $2^{\text {nd }}$ lowest for winter-run, and $3^{\text {rd }}$ lowest for fall-run). For salmon, it is perhaps a fortunate climatic concurrence that low marine productivity seems to be frequently associated with high outmigration flows in California.

The management implications of this study are important: while we do not have the luxury of mitigation actions when it comes to marine conditions, we have some control over conditions in the freshwater environment, and therefore potentially control over $35 \%$ of the annual variability in salmon population abundances, and thus can somewhat buffer these populations from the negative effects of poor marine conditions. Managers should explore approaches to increase river flow and other associated beneficial river conditions during the outmigration season of Central Valley Chinook salmon populations.

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Table 1. The different sources of CWT recoveries and the agency, method, and first collection year for each. In the last column, years highlighted in black represent the first brood year for which SAR was estimated.

| Chinook salmon run | Recovery type | Recovery <br> Agency ${ }^{\text {c }}$ | Collection methods | Brood Year when <br> first available |
| :---: | :---: | :---: | :---: | :---: |
| Winter | Ocean recreational fishery ${ }^{\text {a }}$ | CDFW | Creel surveys | $1991{ }^{\text {b }}$ |
| Winter | Ocean commercial fishery ${ }^{\text {a }}$ | CDFW | Creel surveys | $1991{ }^{\text {b }}$ |
| Winter | River recreational fishery | CDFW | Creel surveys | No fishery |
| Winter | Spawning ground | USFWS | Carcass surveys | 1999 |
| Winter | Hatchery | USFWS | Hatchery returns | $1991{ }^{\text {b }}$ |
| Late-fall | Ocean recreational fishery ${ }^{\text {a }}$ | CDFW | Creel surveys | $1993{ }^{\text {b }}$ |
| Late-fall | Ocean commercial fishery ${ }^{\text {a }}$ | CDFW | Creel surveys | $1993{ }^{\text {b }}$ |
| Late-fall | River recreational fishery | CDFW | Creel surveys | 1998 |
| Late-fall | Spawning ground | CDFW | Carcass surveys | 1999 |
| Late-fall | Hatchery | USFWS | Hatchery returns | $1993{ }^{\text {b }}$ |
| Fall | Ocean recreational fishery ${ }^{\text {a }}$ | CDFW | Creel surveys | $1979{ }^{\text {b }}$ |
| Fall | Ocean commercial fishery ${ }^{\text {a }}$ | CDFW | Creel surveys | $1979{ }^{\text {b }}$ |
| Fall | River recreational fishery | CDFW | Creel surveys | 1998 |
| Fall | Spawning ground | CDFW | Carcass surveys | 1999 |
| Fall | Hatchery | USFWS | Hatchery returns | $1979{ }^{\text {b }}$ |

[^3]Table 2. The estimated annual SAR (\%), standard errors (SE), and total number of release days for each run and each brood year. Standard errors were calculated using equation 6 .

| Brood Year | Late-fall run |  |  | Winter run |  |  | Fall run |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | SAR <br> (\%) | SE | Total Release Days | SAR <br> (\%) | SE | Total Release Days | SAR <br> (\%) | SE | Total Release Days |
| 1993 | 0.50 | 0.07 | 3 |  |  |  |  |  |  |
| 1994 | 1.80 | 0.42 | 5 |  |  |  |  |  |  |
| 1995 | 1.02 | 0.13 | 5 |  |  |  |  |  |  |
| 1996 | 1.64 | 0.23 | 5 |  |  |  |  |  |  |
| 1997 | 0.69 | 0.10 | 6 |  |  |  |  |  |  |
| 1998 | 0.85 | 0.08 | 3 |  |  |  |  |  |  |
| 1999 | 1.03 | 0.14 | 5 | 2.23 | 0.21 | 1 | 3.29 | 0.14 | 3 |
| 2000 | 0.77 | 0.11 | 4 | 0.34 | 0.03 | 1 | 0.78 | 0.05 | 4 |
| 2001 | 1.10 | 0.19 | 4 | 0.24 | 0.02 | 1 | 0.70 | 0.06 | 5 |
| 2002 | 1.44 | 0.25 | 4 | 1.88 | 0.09 | 1 | 0.94 | 0.12 | 2 |
| 2003 | 1.44 | 0.16 | 4 | 1.38 | 0.07 | 1 | 0.30 | 0.04 | 1 |
| 2004 | 0.26 | 0.07 | 4 | 0.08 | 0.01 | 1 | 0.10 | 0.03 | 2 |
| 2005 | 1.72 | 0.24 | 3 | 0.11 | 0.01 | 1 | 0.02 | 0.01 | 2 |
| 2006 | 0.87 | 0.16 | 3 | 0.29 | 0.04 | 1 | 0.04 | 0.01 | 4 |
| 2007 | 0.79 | 0.16 | 3 | 0.28 | 0.05 | 1 | 0.13 | 0.01 | 4 |
| 2008 | 0.56 | 0.05 | 4 | 0.05 | 0.01 | 1 | 0.59 | 0.04 | 3 |
| 2009 | 0.58 | 0.10 | 3 | 0.59 | 0.04 | 2 | 2.39 | 0.09 | 3 |
| 2010 | 1.21 | 0.14 | 3 | 0.43 | 0.06 | 1 | 1.46 | 0.08 | 4 |
| 2011 | 0.91 | 0.09 | 5 | 0.42 | 0.03 | 1 | 0.45 | 0.04 | 3 |
| 2012 | 0.88 | 0.10 | 4 | 0.62 | 0.07 | 1 | 0.15 | 0.02 | 3 |

## Figure Captions

Figure 1. Map of the Central Valley, including portions of major rivers accessible to Chinook salmon populations delineated by major regions, major cities and points of interest, and salmon hatcheries relevant to this study.

Figure 2. A schematic representing the various recapture points for CWTs along the salmon life cycle that contribute to the estimation of a SAR for a given CWT group. The colored arrows represent life stage transitions, each with inherent levels of natural mortality. The circle shape represents hatchery release and rectangles represent CWT recoveries. Green-filled shapes represent events that occur in freshwater, and blue-filled shapes represent events that occur in the ocean. While recoveries of $5+$ year old salmon are possible, they are extremely rare and therefore not represented in this schematic.

Figure 3. The relationship between winter-run Chinook salmon SAR values (\%) and survival from hatchery release to the end of age- $2(\%)$. The solid black line represents the $1: 1$ line. The black dotted line represents the linear model between these two variables, and the grey shaded area the $95 \%$ confidence interval around the linear model. The intercept, slope, r-squared and significance of the linear model is provided in the top left corner of the plot frame.

Figure 4. The range of possible relationships between outmigration survival and marine survival given known CWT release group SAR values for late-fall-run Chinook salmon. Each grey line represents the SAR value for a specific CWT release group, and the point along each line that represents the actual outmigration and marine survival for each release group is unknown, with the exception of the years for which acoustic tagging data outmigration survival estimates existed (black points, respective marine survival estimates with bootstrapped $95 \%$ confidence
intervals represented alongside). The black dashed line represents the median SAR for all CWT release groups. The black dotted line represents the location where outmigration and marine survival are equal (i.e., 1:1 line): if a point falls above this line, marine survival was higher than outmigration survival.

Figure 5. The relationship between freshwater outmigration survival (i.e., release to Chipps Island) for acoustic-tagged late-fall-run Chinook salmon release groups and their associated \% SAR. The red lines represents 1000 linear models between 1000 parametric bootstrapped samples of these two variables, with the mean r-squared (and bootstrapped $95 \%$ confidence intervals) of these models represented in the top left corner of the plot frame.

Figure 6. The relationship between annual SAR and (1) flow during outmigration (a, d, g), (2) upwelling during the first spring at sea (b, e, h), and (3) MOCI during the first spring at sea (c, f, i), for late-fall-run Chinook salmon ( $a, b, c$ ) winter-run Chinook salmon ( $d, e, f$ ) and fall-run Chinook salmon ( $\mathrm{g}, \mathrm{h}, \mathrm{i}$ ). The solid lines in all panels represent the linear model for that relationship, as well as the $r$-squared value. Note that the $r$-squared values in plots $d$ and $g$ did not include the OY 2006 because it was determined to be an outlier (datapoint represented in red).

Figure 7. The relationship between the residuals from the flow versus SAR linear model and spring upwelling during the first spring at sea ( $\mathrm{a}, \mathrm{c}, \mathrm{e}$ ), and between the residuals from the flow versus SAR linear model and spring MOCI during the first spring at sea ( $\mathrm{b}, \mathrm{d}, \mathrm{f}$ ). The dotted lines in all panels represent the zero line for residuals. The points with the largest negative residual values have been labeled with their year of ocean entry. The closer points fall to the zero line, the better they were predicted by the flow model. The three different runs of Chinook salmon are represented: late-fall-run $(a, b)$; winter-run $(c, d)$ and fall-run $(e, f)$.

Figure 8. The influence of (1) flow during outmigration and spring upwelling during the first year at sea on SAR (a), and (2) flow during outmigration and spring MOCI during the first year at sea on SAR (b). Logit-scale SAR values have been standardized; yellow colors represent low SAR values, and blue colors represent high SAR values. Empty symbols represent the location of actual data that were interpolated across; size of these symbols increase proportionally with standardized SAR values. Solid black symbols represent conditions experienced by cohorts for which SAR values are not yet available, spanning OY 2014-2017. Square symbols represent late-fall-run Chinook salmon, circle symbols are for fall-run Chinook salmon, and triangle symbols are for winter-run Chinook salmon (no point exists for OY 2015 fall-run because no CNFH salmon were released in the river that year).


Figure 1.


Figure 2.


Figure 3.


Figure 4.


Figure 5.


Figure 6.


Figure 7.

SAR z-score


Figure 8.

Research Press

# Fishery collapse, recovery, and the cryptic decline of wild salmon on a major California river 

Malte Willmes, James A. Hobbs, Anna M. Sturrock, Zachary Bess, Levi S. Lewis, Justin J.G. Glessner, Rachel C. Johnson, Ryon Kurth, and Jason Kindopp


#### Abstract

Fall-run Chinook salmon (Oncorhynchus tshawytscha) from the Sacramento-San Joaquin River system form the backbone of California's salmon fishery and are heavily subsidized through hatchery production. Identifying temporal trends in the relative contribution of hatchery- versus wild-spawned salmon is vital for assessing the status and resiliency of wild salmon populations. Here, we reconstructed the proportion of hatchery fish on natural spawning grounds in the Feather River, a major tributary to the Sacramento River, using strontium isotope $\left(^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}\right)$ ratios of otoliths collected during carcass surveys from 2002 to 2010 . Our results show that prior to the 2007-2008 salmon stock collapse, $55 \%-67 \%$ of in-river spawners were of hatchery origin; however, hatchery contributions increased drastically ( $89 \%$ ) in 2010 following the collapse. Data from a recent hatchery marking program corroborate our results, showing that hatchery fish continued to dominate ( $\sim 90 \%$ ) in 2011-2012. Though the rebound in abundance of salmon in the Feather River suggests recovery of the stock postcollapse, our otolith chemistry data document a persistent decline of wild spawners, likely leading to the erosion of locally adapted Feather River salmon populations.


Résumé : Les saumons quinnats (Oncorhynchus tshawytscha) à migration automnale du réseau du fleuve Sacramento et de la rivière San Joaquin forment l'épine dorsale de la pêche aux saumons en Californie et sont fortement soutenus par la production en alevinières. La détermination des tendances dans le temps des apports relatifs de saumons issus d'alevinières et de saumons nés dans la nature est cruciale pour évaluer l'état et la résilience des populations de saumons sauvages. Nous avons reconstitué la proportion de poissons issus d'alevinières dans des aires de frai naturelles dans la rivière Feather, un important affluent du fleuve Sacramento, en utilisant les rapports d'isotopes de strontium $\left({ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}\right)$ d'otolites prélevés durant des relevés de carcasses de 2002 à 2010. Les résultats montrent que, avant l'effondrement des stocks de saumons de 2007-2008, de $55 \%$ à $67 \%$ des frayeurs dans la rivière provenaient d'alevinières; toutefois, l'apport d'alevinières a connu une augmentation très marquée ( $89 \%$ ) en 2010 dans la foulée de l'effondrement. Les données tirées d'une campagne récente de marquage en alevinière corroborent ces résultats, démontrant que les poissons issus d'alevinières sont toujours prédominants ( $\sim 90 \%$ ) en 2011-2012. Si la remontée de l'abondance des saumons dans la rivière Feather semble indiquer un rétablissement du stock à la suite de l'effondrement, nos données sur la chimie des otolites documentent un déclin soutenu des géniteurs sauvages, qui mène vraisemblablement à l'érosion des populations de saumons adaptées aux conditions locales de la rivière Feather. [Traduit par la Rédaction]

## Introduction

The Sacramento-San Joaquin River system in California's Central Valley (CV) is the foundation of California's water supply, providing water for approximately 35 million residents and supporting a multibillion dollar agriculture industry, and is home to the southernmost spawning runs of Chinook salmon (Oncorhynchus tshawytscha) in the Northern Hemisphere (Fisher 1994; Yoshiyama et al. 1998; Moyle 2002; Williams 2006). Chinook salmon populations have persisted in California's highly variable Mediterranean climate by exhibiting a diverse portfolio, expressed as distinct run types (spring, fall, late-fall, winter) and plastic life history strategies (Yoshiyama et al. 1998; Hilborn et al. 2003; Williams 2006), which buffers population abundance against stochastic environmental variability. However, habitat loss and degradation, water
diversions, fish harvest, and the construction of dams, which blocked large areas ( $>80 \%$ ) of spawning habitat and rearing grounds, have resulted in population decline threatening the long-term survival of salmon in the CV (Yoshiyama et al. 2000, 2001). Spring- and winter-run Chinook salmon are listed as threatened and endangered, respectively, under the federal Endangered Species Act (NMFS 1999, 2005), while fall-late-fall-run salmon are considered species of concern and are targeted for harvest in the ocean fishery.

Hatcheries were built along CV tributaries to mitigate for dam construction and habitat loss, and many salmon populations in the CV are heavily subsidized by hatchery production (HSRG 2012, 2014; Palmer-Zwahlen and Kormos 2015). Fall-run Chinook salmon from the Sacramento-San Joaquin River system form the backbone of California's ocean salmon fishery, contributing sub-

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Fig. 1. Fall-run escapement estimates for the Sacramento River System (grey) and the Feather River (hatchery and in-river spawning population; black) from 1975 to 2016. Data from GrandTab2017.04.07, California Central Valley Chinook Population Database Report. An asterisk (*) indicates that there is no in-river escapement data available for 1990, 1998, and 1999 for the Feather River.

stantially to the fisheries off Oregon and Washington (Lindley et al. 2009; Satterthwaite et al. 2015), and are an integral part of the present and past culture in this region (Yoshiyama et al. 1998). However, wild stocks in several California rivers are now dominated by hatchery fish (Barnett-Johnson et al. 2007; Johnson et al. 2012; Quiñones and Moyle 2014), potentially eroding the longterm resiliency of wild, locally adapted populations by disrupting selection for heritable traits that improve lifetime reproductive success in variable environments.

In 2007, record low numbers of adult salmon returned to the CV (Fig. 1), and forecasted low escapement resulted in the closure of the commercial ocean fishery off the coast of California and Oregon in 2008 and 2009 for the first time in over 100 years, causing major economic impact (Schwarzenegger 2008; Michael 2010). While the proximate cause of this stock collapse was attributed to low food availability in the coastal ocean in spring 2005 and 2006 (Lindley et al. 2009), the effect of hatchery practices likely contributed to a weakened CV salmon portfolio through increasing synchrony in fall-run population dynamics, further exacerbating the impact of climatic variability (Satterthwaite and Carlson 2015). After 2009, Chinook salmon fall-run escapement numbers rebounded, suggesting a quick and successful recovery of the salmon stock, before the decreases in 2014 and 2015 that were potentially linked to the recent prolonged drought period (Dettinger and Cayan 2014). Owing to the continued declines in wild salmon abundance, lack of hatchery management reform, and impending climate change, the fate of wild salmon in California are in jeopardy, and extinction in the wild is deemed likely if drastic management actions are not taken (Katz et al. 2012; Franks and Lackey 2015; Moyle et al. 2017).

Effective management, monitoring, and status assessment of wild salmon populations require reliable estimates of hatchery fish abundance on the natural spawning grounds (Araki et al. 2008; HSRG 2012, 2014; Christie et al. 2014; Quiñones et al. 2014). However, only recently have hatchery fish been consistently physically marked (adipose fin clip) and tagged (coded wire tags) in California rivers (Lindley et al. 2007; HSRG 2014). The California Constant Fractional Marking Program (CFM) began marking 25\% of fall-run hatchery releases in 2007, providing a method to estimate hatchery contributions to natural spawning grounds since 2010 (Kormos et al. 2012; Palmer-Zwahlen and Kormos 2013, 2015). The results of the CFM program showed that most CV rivers have
a high contribution of hatchery-origin fish in their natural spawning grounds, particularly those co-located with hatcheries, such as the Feather River (Palmer-Zwahlen and Kormos 2013). However, the CFM data cannot provide estimates prior to 2010 and thus does not inform stock composition prior to and during the salmon stock collapse. Without such estimates, it is impossible to quantify trends in the abundance of natural-origin fish (Johnson et al. 2012) or to evaluate population extinction risk (Lindley et al. 2007; Katz et al. 2012). Ultimately, understanding the extent of gene flow between hatchery- and natural-origin spawners is critical, since the influence of too many hatchery adults can reduce the fitness of subsequent generations in the wild (Waples 1991; McGinnity et al. 2003; Araki et al. 2008; Christie et al. 2014).

A variety of methods have been used to discriminate between hatchery- and natural-origin salmonids, including physical tags (Palmer-Zwahlen and Kormos 2015), genetics (Hauser et al. 2006), otolith microstructure (Barnett-Johnson et al. 2007), and otolith chemistry (Johnson et al. 2012, 2016). Otoliths (i.e., ear stones) consist of calcium carbonate and are found in the inner ear of bony fishes. They are metabolically inert and accrete continuously, producing a unique record of fish age and growth. Chemical elements from the environment are incorporated into the otolith, resulting in a chemical composition that can reflect the habitat occupied during deposition (Campana 1999). Strontium $(\mathrm{Sr})$ is readily substituted for calcium in the mineral lattice, resulting in element concentrations and isotope ratios $\left({ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}\right)$ that reflect environmental abundances and are frequently used to reconstruct individual movements (Rooker et al. 2001; Walther and Limburg 2012). In the CV, ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ can be a powerful natural tag of fish origin, because the water ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope ratios vary among many of the salmon-producing rivers and hatcheries (Ingram and Weber 1999; Barnett-Johnson et al. 2008; Sturrock et al. 2015).

Here, we use otolith ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope ratios to identify natal origin of adult Chinook salmon spawning in the Feather River to determine the annual contribution of hatchery-produced fish to escapement years 2002-2010, encompassing the years of stock collapse and recovery. We refer to individuals that reared in the river as wild fish and fish that reared in the hatchery as hatchery fish, independent of their parental or genetic origin. Furthermore, we focused only on phenotypic fall-run salmon, defined as returning to in-river spawning grounds after 1 September, and did not examine the genetic run identity of these fish.

Fig. 2. Overview map of the study region and the hatcheries producing Chinook salmon. The Feather River is further divided into the High Flow and Low Flow channels. Data from the National Hydrography Dataset, US Geological Survey.


## Materials and methods

## The Feather River

The Feather River Basin is located in the foothills of the western Sierra Nevada (Fig. 2). The basin is a major contributor to the California State Water Project, and Lake Oroville, created by the completion of Oroville Dam in 1967, plays an important role in flood management, water storage, water quality, power generation, and recreation. The Fish Barrier Dam represents the uppermost barrier to upstream fish migration, as well as the location of the fish ladder entering the Feather River Hatchery (FRH). In addition to the main hatchery, there is an annex hatchery located along Highway 99, about 2 km south of Oroville Dam Blvd., with warmer water temperatures that provide opportunities for increased growth rates.

The Feather River Hatchery is one of the largest producers of CV Chinook salmon, supporting both spring- and fall-run populations (Fisher 1994; Yoshiyama et al. 1998). Historically, spring-run salmon returned from the ocean in spring-early summer and
then held over summer and spawned in the uppermost reaches of small tributaries to the Feather River, while fall-run fish returned later in the fall and spawned in the lower foothill reaches of the mainstem river. Spawning for both populations is concentrated in approximately 12 river kilometres below Oroville Dam (Mercer and Kurth 2014). Hatchery broodstock management has attempted to separate the two runs; however, considerable mixing has occurred, resulting in substantial genetic introgression (Clemento et al. 2014; Meek et al. 2016b), and spring-run fish have hybridized with fall-run fish, resulting in overlapping run timings and frequent examples of "run-switching" between parents and offspring (Sommer et al. 2001).

The FRH maintains an integrated hatchery program resulting in considerable mixing of hatchery and wild fish in the fall-run hatchery broodstock and in-river spawning population (Williamson and May 2005; HSRG 2012). To reduce in-river mortality during seaward migration, juveniles produced by the FRH (and many other hatcheries in the CV) are trucked directly to San Pablo Bay

Table 1. Number of otoliths sampled and analyzed by recovery location (High Flow and Low Flow channels).

| Year | Population estimates |  |  | Analyzed otoliths |  |  | Proportion of population sampled by otoliths |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | High Flow | Low Flow | Total | High Flow | Low Flow | Total | High Flow | Low Flow | Total |
| 2002 | 34125 | 71038 | 105163 | 41 | 70 | 111 | 0.15 | 0.12 | 0.13 |
| 2003 | 37643 | 52303 | 89946 | 41 | 54 | 95 | 0.11 | 0.10 | 0.11 |
| 2004 | 17113 | 37058 | 54171 | 35 | 64 | 99 | 0.24 | 0.17 | 0.19 |
| 2005 | 12583 | 36577 | 49160 | 33 | 70 | 103 | 0.27 | 0.19 | 0.21 |
| 2006 | 16990 | 59424 | 76414 | 23 | 49 | 72 | 0.14 | 0.08 | 0.09 |
| 2007 | 876 | 21033 | 21909 | 33 | 76 | 109 | 3.78 | 0.36 | 0.50 |
| 2008 | 297 | 5642 | 5939 | 27 | 62 | 89 | 9.09 | 1.10 | 1.50 |
| 2009 | 223 | 4624 | 4847 | 7 | 38 | 45 | 3.14 | 0.80 | 0.91 |
| 2010 | 2201 | 42713 | 44914 | 4 | 28 | 32 | 0.18 | 0.07 | 0.07 |

Note: Population estimates are from GrandTab2017.04.07, California Central Valley Chinook Population Database Report and California Department of Water Resources (unpublished data, contact Jason.Kindopp@water.ca.gov).
and acclimatized in net pens prior to release. Since 2000, FRH has released $80 \%-100 \%$ of its fall-run production directly into the San Francisco Estuary or Bay (Huber and Carlson 2015). Because of a lack of olfactory imprinting during juvenile emigration, juveniles that are trucked stray disproportionately as adults to other rivers, resulting in increased gene flow among salmon populations (Palmer-Zwahlen and Kormos 2013, 2015; Huber and Carlson 2015; Meek et al. 2016a). Hatchery fish are also released over a relatively short time window, leading to reduced diversity in emigration phenology and increased risk of mismatch with optimal ocean conditions (Satterthwaite et al. 2014; Huber and Carlson 2015).

## Otolith collection

Otoliths were collected from postspawned Chinook salmon between 2002 and 2010 as part of the annual carcass survey. For this survey, the Feather River is divided into 40 stream sections, each section corresponding to a single riffle-pool complex. The Low Flow Channel includes the Feather River from the Fish Barrier Dam to the Thermalito Outlet, and the High Flow Channel extends from the Thermalito Outlet downstream to the Gridley Bridge (Fig. 2). Otoliths were collected from a total of 50 fish per week, among 10 river sections randomly selected each week, five in the Low Flow Channel and five in the High Flow Channel. To ensure that these fish are representative of the overall population, the first five salmon carcasses, irrespective of size, sex, and presence or lack of adipose fin, were sampled within each of these randomly selected locations. This stratification of the river into sections ensures that the entire river's spawning grounds are surveyed equally (Table 1).

From this set of collected fish, a subset $(n=755)$ was selected for otolith analysis in proportion to fish abundances in the High Flow and Low Flow channels for each year. However, this was not possible for all years, leading to an uneven sample distribution between the High Flow and Low Flow channels. We used the escapement estimates for each recovery location as a weighting factors to incorporate this variation in subsampling into our estimations of hatchery and wild contributions to the overall escapement. In addition to these stratified samples, fish with coded wire tags (CWT, $n=110$ ) were randomly selected and used to validate the otolith isotope assignments (natal origins).

## Otolith sample preparation

Sagittal otoliths were extracted from each fish, cleaned, dried, labeled, and transferred to the Department of Wildlife, Fish, and Conservation Biology, University of California Davis. Otoliths were mounted in Epocure (Buehler Scientific) epoxy resin and thin-sectioned with an Isomet diamond cutting saw in the transverse plane. Thin sections were adhered to glass microscope slides with Crystal Bond thermoplastic resin (Crystalbond 509, Ted Pella Inc., Redding, California), sanded to the core on both sides with $1200-2000$ grit sandpaper, and polished with $0.3 \mu \mathrm{~m}$ alumina and a polishing cloth, following methods from Wells et al. (2003).

Digital images of otoliths were taken at $6 \times$ magnification on a CH30 Olympus compound microscope. Otoliths sections were washed with $1 \mathrm{~mol} \cdot \mathrm{~L}^{-1}$ chemical grade nitric acid for 5 to 10 s , rinsed in an ultrasonic water bath for 5 min , and dried under a class 100 laminar flow hood.

## Otolith aging

Otoliths contain a unique time series of opaque and translucent bands that are deposited on a daily and seasonal basis in response to photoperiod, temperature, diet, and endogenous rhythms (Neilson and Geen 1982; Campana and Neilson 1985). On adult otoliths, an opaque zone followed by a translucent zone (Fig. 3) represents 1 year of otolith growth (Welch et al. 1993). Annual ages were estimated from digital images along the transverse plane of the ventral lobe (Fig. 3), counting the summer bands. We used the transverse section instead of the sagittal preparation typically used for juvenile habitat use and growth reconstructions (Woodson et al. 2013; Sturrock et al. 2015) to preserve the outer rings in the convex adult otoliths. Otoliths that were completely vateritic, or broken along the ventral lobe by the sanding process, were not aged. Ages were validated by comparing age counts with those of hatchery fish with physical tag information and known age ( $n=$ 74) and between two age readers following the methods proposed by Campana (2001) and using the FSA package (Ogle 2018) in R ( R Core Team 2017). All fish were aged as either 2-, $3-, 4-$, or 5 -yearolds, consistent with the currently understood life history of the species (Fisher 1994).

## Otolith Sr isotopic analysis

For laser ablation, otoliths were remounted on petrographic glass slides, with 20 individual otoliths per slide. ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope ratios were measured at the University of California Davis Interdisciplinary Center for Plasma Mass Spectrometry. For the in situ Sr isotope analysis, an Nd:YAG 213 nm laser (New Wave Research UP213) was coupled to a Nu Plasma HR MC-ICP-MS (Nu032). A laser beam of $55 \mu \mathrm{~m}$ diameter was traversed across the otolith from the core to the edge at $10 \mu \mathrm{~m} \cdot \mathrm{~s}^{-1}$, with the laser pulsing at 10 Hz frequency and $5-15 \mathrm{~J} \cdot \mathrm{~cm}^{-2}$ photon output. The ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope ratio was normalized for instrumental mass discrimination by monitoring the ${ }^{86} \mathrm{Sr} /{ }^{88} \mathrm{Sr}$ isotope ratio (assuming ${ }^{86} \mathrm{Sr} /{ }^{88} \mathrm{Sr}=$ 0.1194 ), and ${ }^{87} \mathrm{Rb}$ was corrected by monitoring the ${ }^{85} \mathrm{Rb}$ signal. Krypton interference originating in the argon supply ( ${ }^{86} \mathrm{Kr}$ ) was subtracted using the on peak zero method before each analysis. Kr contribution was monitored throughout the analyses, as increasing amounts of Kr would lead to an increased uncertainty of the individual measurement. Operating conditions and reproducibility of standards on the LA-MC-ICP-MS were evaluated using a modern marine coral from the South China Sea and a modern marine otolith from a white seabass (Atractoscion nobilis) collected offshore of Baja California. Replicate analyses for the coral yielded a mean $( \pm 2 \sigma)^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope ratio of $0.70921 \pm 0.00008(n=61)$ and for the

Fig. 3. (A) Image of a transverse section of an adult otolith in transmitted light, showing growth increments used to estimate ages as well as the laser trajectory (white dotted line) from the strontium isotopic analysis. (B) Corresponding ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope profile with a loess smooth (span $=0.1$ ) applied; grey bands represent the $95 \%$ confidence intervals. Green box indicates the approximate region used for the natal origin assignment. [Colour online.]

otolith of $0.70919 \pm 0.00003(n=63)$. These values are in good agreement with the mean modern ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope value of seawater, 0.70918 (McArthur et al. 2001).

## Baseline Sr isotope data

The ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope ratio baseline for the Sacramento River System was compiled from published data (Ingram et al. 1999; Barnett-Johnson et al. 2008; Sturrock et al. 2015) and newly collected water and otolith samples of known origin from five locations along the Feather River, from the FRH, and from the Thermalito Annex.

Water samples were collected at base flow conditions during the fall of 2013, in the direct flow of water using 50 mL polypropylene tubes, acidified with $1 \mathrm{~mol} \cdot \mathrm{~L}^{-1}$ nitric acid and filtered with a $0.45 \mu \mathrm{~m}$ filter. The samples were transported to a class 100 clean room facility at the UC Davis Interdisciplinary Center for Plasma Mass Spectrometry. An aliquot of each water sample was made at a volume totaling approximately $1 \mu \mathrm{~g}$ of Sr . This volume was evaporated to dryness in an acid-leached polytetrafluoroethylene
(Teflon) vial on a hotplate, and Sr was isolated from all other aqueous constituents by selective ion exchange chromatography (Horwitz et al. 1992). Sr separates were reconstituted in $2 \% \mathrm{HNO}_{3}$ and introduced in the MC-ICP-MS (Nu Plasma HR) using a desolvating nebulizer introduction system (Nu Instruments DSN-100). ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ data were internally normalized by the measured ${ }^{86} \mathrm{Sr} /$ ${ }^{88} \mathrm{Sr}$ ratio (assuming ${ }^{86} \mathrm{Sr} /{ }^{88} \mathrm{Sr}=0.1194$ ). ${ }^{85}$ Rubidium was monitored to correct for ${ }^{87} \mathrm{Rb}$ if present, but all were well below the Rb correction threshold due to the selective ion exchange chromatography beforehand. ${ }^{84} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ was monitored to estimate the ${ }^{84} \mathrm{Kr} /{ }^{86} \mathrm{Kr}$ isotope ratio. ${ }^{86} \mathrm{Kr}$ was subtracted until the ${ }^{84} \mathrm{Sr} /{ }^{88} \mathrm{Sr}$ ratio equalled the canonical value of 0.006755 , while iterating the mass-bias correction. Procedural blank was measured and contributed $<0.002 \%$ of total Sr processed per sample. Replicated analyses of NIST SRM 987 were conducted every six samples, normalizing for instrument drift over the course of the day and for analytical artifacts among sessions. An in-house modern marine coral standard was processed in parallel with water samples and
resulted in a mean $(2 \sigma)^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope ratio of $0.70918 \pm 0.00002$ ( $n=8$ ).

Sr isotope ratios from the natal region of hatchery-reared otoliths (Fig. 3) were analyzed by LA-MC-ICP-MS, following the same protocols as the samples collected from the spawning grounds, to determine the range of isotope ratios indicative of rearing in the FRH. We examined otoliths from 10 fall-run Chinook salmon reared at the Feather River's main hatchery facility in two raceways collected on 19 April 2012 and ranging in fork length from 43 to 78 mm and 10 fish reared at the Thermalito Annex from two raceways removed on 19 March 2012 and ranging in fork length from 88 to 90 mm .

## Natal origin assignments

To assign natal origin, the otolith material deposited immediately following onset of exogenous feeding (i.e., with no isotopic influence of the maternal yolk) was visually identified in the Sr isotope profile and matched to the distance from the otolith core (typically $\sim 250 \mu \mathrm{~m}$ ). The mean ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotopic ratio for this natal portion of the profile (Fig. 3) was then assigned to a source location by matching it to the established Sr isotopic baseline for the Sacramento River System (Ingram and Weber 1999; Barnett-Johnson et al. 2008) using single-factor quadratic discriminant function analysis (QDFA) in $R$ ( $R$ Core Team 2017). We used the quadratic function instead of a linear function because it relaxes the assumption that all the variances of the ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ values from groups are the same. There are some ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ overlaps between natal sources in the San Joaquin and Sacramento basins (particularly Mokelumne versus Feather River Hatcheries and Merced versus Yuba rivers; Sturrock et al. 2015), which could result in strays from the San Joaquin basin potentially being misclassified by our Sacramento-focused QDFA. However, given the large production differences among basins, San Joaquin origin strays have little numerical effect, with the combined contribution of Merced and Mokelumne Hatchery strays to the Feather River and FRH escapements being $<1 \%$ in 2010-2012 (Kormos et al. 2012; Palmer-Zwahlen and Kormos 2013, 2015). Given that CV hatchery fish are likely to stray at much higher rates than natural-origin fish due to the extensive trucking program, we assumed that the potential error rate attributable to misclassified San Joaquin origin strays was below $1 \%$. We made no attempt to adjust prior probabilities in the QDFA based on annual hatchery production estimates.

Otolith subsample sizes among weekly surveys and channel strata were too small to calculate proportion of hatchery fish at fine spatial and temporal scales; therefore, samples were pooled by survey year retaining the stratification by High Flow and Low Flow channels. To provide a robust estimate of the proportion of hatchery-origin fish on the spawning grounds, we estimated the mean and standard deviation using bootstrapping with 1000 iterations and sample sizes equal to the number of otoliths collected in each channel for each year. The annual contribution and number of hatchery-origin fish on the spawning grounds was estimated by expanding the bootstrapped mean proportion by the escapement estimates in the High Flow and Low Flow channels for each year. The accuracy of our hatchery classifications was evaluated using otoliths of known hatchery origin $(n=110)$ that were included in the sample set without prior decoding of their origin.

## Emigration timing

To investigate the difference in timing of ocean entry of wild- and hatchery-origin fish, we compared fall-run hatchery release data from Huber and Carlson (2015) with catch data from the USFWS Chipps Island Midwater Trawl Survey. The hatchery release data were filtered to include only fall-run FRH-produced fish that were released into San Pablo Bay from 2002 to 2012 and normalized for each day of the year by the total number of fish released that year. Chipps Island Survey data was filtered to include only fall-run sized, unmarked fish captured in 2002-2012 and then normalized
for each Julian day using catch per unit effort and total catch for that year. Note that the latter will therefore include unmarked hatchery fish released upstream of Chipps Island, for example from the Coleman National Fish Hatchery (75\% unmarked), and thus likely represents a lower estimate for the true emigration timing variability of wild fish. This allows us to compare emigration timing irrespective of differences in interannual abundance.

## Results

## Ages

Otolith annual band counts provided a reliable determination of fish age. Otolith age estimates of known-age fish showed high accuracy: $92 \%(n=74)$. Fish incorrectly aged were $\pm 1$ year of known age, $5 \%$ were estimated to be 1 year older, and $3 \%$ were estimated to be 1 year younger than their known age. Age estimations between the two age readers across all otoliths $(n=755)$ reached an agreement of $92 \%(\mathrm{ACV}=1.771, \mathrm{APE}=1.252)$. This level of agreement between readers is comparable to that of other otolith aging studies (Flain and Glova 1988; Murray 1994; Secor et al. 1995). Of the individuals aged, $6.1 \%, 68.6 \%, 24.6 \%$, and $0.7 \%$ were estimated as 2-, 3-, 4-, and 5-year-old fish, respectively. This age distribution was similar to age class distribution estimates for the Feather River (Grover and Kormos 2008). Similarly, of the CWT retrieved from fish at FRH from 2002 through 2007, 15.58\% were 2-year-olds, $56.07 \%$ were 3 -year-olds, $28.23 \%$ were 4 -year-olds, and $0.13 \%$ were 5-year-olds (Mesick et al. 2009).

## Natal origins

The baseline ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope ratios for the Sacramento River System varied significantly between different rivers and hatcheries (Fig. 4; Table 2). Using QDFA, we achieved an overall classification success rate of $96 \%$, providing a robust baseline to determine the natal origins of Chinook salmon in this river system. Furthermore, $95 \%$ of known-origin (CWT) fish were correctly assigned to the FRH (Table 3). Classification success rate varied by collection year, ranging from $75 \%$ in 2002 to 100\% in 2006, 2009, and 2010.

Sr isotope profiles from the core to the edge of otoliths for 755 individuals revealed distinct patterns in natal origins and life histories (Fig. 5). All otoliths examined reached ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope ratios equivalent to the global ocean value 0.70918 (McArthur et al. 2001) prior to the first annual band, indicating that all individuals had entered the ocean in their first year of life. Each otolith was classified based on its natal origin, escapement year, and recovery location (Table 4). A majority of fish were classified as FRH fish ( $n=$ 373), with a small contribution from the Thermalito Annex ( $n=15$ ) or as wild fish from the Feather River $(n=292)$. The remaining fish were classified as strays within the Sacramento River System, most originating from the Yuba River $(n=32)$ and the Nimbus Fish Hatchery ( $n=35$ ) on the American River, with minor contributions from the Coleman National Fish Hatchery $(n=6)$ and the Northern Tributaries $(n=2)$. The relatively large presence of Yuba River strays is likely explained by the fact that the Feather-Yuba confluence is only about 40 river miles ( 1 mile $=1.609 \mathrm{~km}$ ) downstream from the FRH, with well-documented exchange occurring between the two tributaries (Yuba Accord RMT 2013).

## Changes in spawning composition over time

The proportion of hatchery- and wild-origin fish varied throughout the time series (2002-2010; Fig. 6; Table 5). Since stray fish are included in the overall CFM estimates of hatchery and wild contributions in the CV, we combined them with the FRH- or wildorigin fish, based on their natal assignment (hatchery strays or wild strays).

The contribution of hatchery origin fish on the Feather River spawning grounds before the stock collapse (2002-2006) varied from $55 \% \pm 7 \%( \pm 1 \sigma)$ to $67 \% \pm 9 \%( \pm 1 \sigma)$ (Fig. 6). During the collapse (2007-2008), the proportion of hatchery fish decreased to $40 \% \pm 7 \%$ $( \pm 1 \sigma)$ in 2008. After the collapse (2009), the contribution of hatchery

Fig. 4. Boxplot showing the ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope ratios of natal sources in the Sacramento River System assumed to potentially contribute to the Feather River escapement. Note that the northern tributaries (Upper Sacramento, Mill Creek, Deer Creek, Battle Creek, Butte Creek) were combined and treated as a single source. The acronyms are NT (northern tributaries), CNH (Coleman National Fish Hatchery), THE (Thermalito Annex), FEA (Feather River), FRH (Feather River Hatchery), YUB (Yuba River), NIH (American River Nimbus Fish Hatchery), AME (American River). [Colour online.]


Table 2. Summary statistics for the baseline ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope ratios for the Sacramento River System.

|  | Northern <br> tributaries | Coleman National <br> Fish Hatchery | Feather <br> River | Feather River <br> Fish Hatchery | Thermalito <br> Annex | Yuba <br> River | Nimbus Fish <br> Hatchery |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Min. | 0.70382 | 0.70481 | 0.70594 | 0.70655 | 0.70569 | 0.70762 | 0.70968 |
| American |  |  |  |  |  |  |  |
| $Q_{1}$ | 0.70407 | 0.70519 | 0.70612 | 0.70700 | 0.70578 | 0.70815 | 0.70971 |
| Median | 0.70411 | 0.70532 | 0.70618 | 0.70717 | 0.70582 | 0.70822 | 0.70971 |
| Mean | 0.70421 | 0.70533 | 0.70623 | 0.70712 | 0.70584 | 0.70823 | 0.70974 |
| $Q_{2}$ | 0.70420 | 0.70541 | 0.70630 | 0.70726 | 0.70591 | 0.70832 | 0.70975 |
| Max. | 0.70489 | 0.70615 | 0.70684 | 0.70756 | 0.70599 | 0.71022 |  |
| $n$ | 41 | 13 | 31 | 37 | 10 | 0.71025 |  |

Note: Data from Ingram and Weber (1999), Barnett-Johnson et al. (2008), Sturrock et al. (2015), and this study.

Table 3. Classification success rate of known-origin (coded wire tags, CWT) Feather River Hatchery fish using quadratic discriminant function analysis and the Sacramento River basin ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope baseline.

|  |  | Assigned to |  |  |
| :--- | :---: | :---: | :--- | :---: |
| Collection <br> year | No. of fish <br> with CWT | Hatchery | River |  | | Correct |
| :--- |
| assignment (\%) |$|$|  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
| 2002 | 8 | 6 | 2 | 75 |
| 2003 | 7 | 6 | 1 | 86 |
| 2004 | 8 | 7 | 1 | 88 |
| 2005 | 16 | 15 | 1 | 94 |
| 2006 | 6 | 6 | 0 | 100 |
| 2007 | 11 | 10 | 1 | 91 |
| 2008 | - | - | - | - |
| 2009 | 5 | 49 | 0 | 100 |
| 2010 | 49 | 104 | 0 | 100 |
| Total | 110 |  | 6 | 95 |

fish increased rapidly to $89 \% \pm 8 \%( \pm 1 \sigma)$ in 2010 and, according to CFM data, remained at 90\% for 2011-2012 (Kormos et al. 2012; Palmer-Zwahlen and Kormos 2013, 2015).

For 2010, both CFM and microchemistry data are available, with CFM estimates yielding $78 \%$ hatchery fish and otolith estimates $89 \% \pm 8 \%( \pm 1 \sigma)$. However, CFM data from 2010 is thought to underestimate the proportion of hatchery fish due to problems in the identification of hatchery fish from decayed carcasses (Mohr and Satterthwaite 2013). Given the directionality of bias in the CFM estimate, the otolith microchemistry and CFM estimates appear compatible.

## Emigration timing

During this time series (2002-2012), Feather River fall-run hatchery fish were almost exclusively ( $\sim 95 \%$ ) trucked and released directly into San Pablo Bay (Huber and Carlson 2015). The timings of the hatchery fish releases overlap within their interquartile range with the timing of wild emigration, suggesting that most of the fish enter the ocean at a similar time. However, the

Fig. 5. Example otolith ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope profiles of Chinook salmon with different natal origins. The cores show an influence of the marine strontium isotopic signature, indicating that the parents of these fish matured in the ocean and that this marine signature was incorporated into the organism prior to hatching, typical for fall-run Chinook salmon. Shaded bars are the ${ }^{87} \mathrm{Sr} /{ }^{86} \mathrm{Sr}$ isotope ranges of the different freshwater source regions from Fig. 4. The acronyms are NT (northern tributaries), CNH (Coleman National Fish Hatchery), THE (Thermalito Annex), FEA (Feather River), FRH (Feather River Hatchery), YUB (Yuba River), NIH (American River Nimbus Fish Hatchery), AME (American River). No fish were classified as American River natal origin. [Colour online.]

variance of the wild emigrating fish is larger, with both earlier and later emigrants, than the hatchery releases, which occur over a shorter time period (Fig. 7).

## Discussion

Identifying the contribution of hatchery-origin fish to a population is essential for assessing the status, fitness, and resilience of locally adapted wild salmon populations (Lindley et al. 2007; Williams et al. 2016). Numerous studies have documented reduced fitness (Araki et al. 2008; Christie et al. 2014) and loss of diversity (portfolio) in populations with supplementation from hatchery-reared fish, causing overall reduced resilience (Schindler et al. 2010). In this study, otolith Sr isotope analysis was highly successful in identifying natal origin and discerning hatchery from wild origins of Chinook salmon in the Feather River. Combining our otolith-based approach with CFM data allowed us to reconstruct an 11-year record of hatchery contributions to the in-river escapement. Temporal trends in the contribution of hatchery- and wild-origin fish in our time series document an increase in the proportion of hatchery-origin fish on the natural spawning grounds after the salmon stock collapse (Figs. 6 and 8).

This substantial change in the proportion of hatchery fish was likely the result of relatively stable hatchery production from the small numbers of returning fish during stock collapse and either poor production or survival of wild offspring in-river during the 2007-2009 drought (DWR 2010).

The proximate cause of the 2007-2009 salmon stock collapse has been attributed to poor ocean conditions and food availability in the ocean after emigration (Lindley et al. 2009; Wells et al. 2016). Ocean conditions off the Californian coast are highly variable, with wind patterns driving the intensity and timing of coastal upwelling, influencing food production and survival of young salmon during ocean entry (Mantua et al. 1997; Satterthwaite et al. 2014; Sabal et al. 2016; Wells et al. 2016). Hatchery-produced smolts entering the ocean in 2005 and 2006 experienced weakly upwelled ocean conditions with variable timing of the spring transition (Lindley et al. 2009), which led to elevated ocean mortality and decline in hatchery contributions and stock abundance on the Feather River in 2008 (Huber and Carlson 2015). Meanwhile, wild fish exhibited a broader ocean arrival window during emigration (Fig. 7), increasing the odds that at least part of the population would be matched with optimal feeding opportunities

Table 4. Number of otoliths by escapement year and recovery location classified to their natal origin.

| Escapement year | Recovery location | Hatchery-origin fish (otolith $n$ ) |  |  |  |  | Wild-origin fish (otolith $n$ ) |  | Total (otolith $n$ ) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | FRH | THE | NIH | CNH | FEA | YUB | NT | Hatchery | Wild |
| 2002 | High Flow | 18 | 2 | 2 | - | 19 | - | - | 22 | 19 |
|  | Low Flow | 25 | 3 | 11 | - | 28 | 2 | 1 | 39 | 31 |
| 2003 | High Flow | 18 | 3 | 2 | 1 | 17 | - | - | 24 | 17 |
|  | Low Flow | 29 | 1 | 2 | - | 20 | 2 | - | 32 | 22 |
| 2004 | High Flow | 17 | 1 | 2 | - | 12 | 3 | - | 20 | 15 |
|  | Low Flow | 31 | - | 5 | - | 22 | 6 | - | 36 | 28 |
| 2005 | High Flow | 12 | 1 | - | - | 16 | 3 | 1 | 13 | 20 |
|  | Low Flow | 47 | 1 | 1 | 1 | 17 | 3 | - | 50 | 20 |
| 2006 | High Flow | 13 | - | - | - | 9 | 1 | - | 13 | 10 |
|  | Low Flow | 34 | - | - | - | 14 | 1 | - | 34 | 15 |
| 2007 | High Flow | 11 | 1 | - | 1 | 19 | 1 | - | 13 | 20 |
|  | Low Flow | 42 | - | 3 | 3 | 24 | 4 | - | 48 | 28 |
| 2008 | High Flow | 7 | - | - | - | 20 | - | - | 7 | 20 |
|  | Low Flow | 21 | 1 | 3 | - | 35 | 2 | - | 25 | 37 |
| 2009 | High Flow | 2 | - | - | - | 5 | - | - | 2 | 5 |
|  | Low Flow | 19 | 1 | 3 | - | 11 | 4 | - | 23 | 15 |
| 2010 | High Flow | 3 | - | - | - | 1 | - | - | 3 | 1 |
|  | Low Flow | 24 | - | 1 | - | 3 | - | - | 25 | 3 |

Note: The northern tributaries (Upper Sacramento, Mill Creek, Deer Creek, Battle Creek, Butte Creek) were combined and treated as a single source. The acronyms are NT (northern tributaries), CNH (Coleman National Fish Hatchery), THE (Thermalito Annex), FEA (Feather River), FRH (Feather River Hatchery), YUB (Yuba River), and NIH (American River Nimbus Fish Hatchery).

Fig. 6. Time series of the proportions of fall-run Chinook salmon on the Feather River assigned to each of the seven natal habitats. No fish were assigned to the American River. [Colour online.]

(Wells et al. 2016). We hypothesize that this diversity in outmigration timing may have led to the observed increase in the wild component of the 2008 returns. However, even with this apparent difference in the resiliency of hatchery and wild populations, all populations declined precipitously during this period. Thus, the population collapse can be attributed to poor ocean survival for
both hatchery- and wild-origin smolts from the 2005 and 2006 emigration years (Lindley et al. 2009).

Total escapement to the Feather River and the Sacramento Valley increased rapidly from 2009 to 2013 (Fig. 8), reaching near prestock collapse abundances in just 3 years. Our analysis showed that $\mathbf{\sim 9 0 \%}$ of these fish returning to the Feather River were of

Table 5. Contributions of hatchery origin fish on the Feather River spawning grounds from 2002 to 2010.

| Escapement year | Recover location | Population estimate | Hatchery origin (\%) | Hatchery origin (1 $\sigma$ ) | Hatchery <br> fish ( $n$ ) | Hatchery <br> fish ( $1 \sigma$ ) | Combined hatchery origin $(\% \pm 1 \sigma)$ | Combined hatchery fish $(n \pm 1 \sigma)$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2002 | High Flow | 34125 | 54 | 8 | 18348 | 2700 |  |  |
| 2002 | Low Flow | 71038 | 56 | 6 | 39508 | 4295 | $55 \pm 7$ | $57863 \pm 7106$ |
| 2003 | High Flow | 37643 | 59 | 8 | 22115 | 2951 |  |  |
| 2003 | Low Flow | 52303 | 59 | 7 | 31014 | 3506 | $59 \pm 7$ | $53127 \pm 6483$ |
| 2004 | High Flow | 17113 | 57 | 8 | 9781 | 1451 |  |  |
| 2004 | Low Flow | 37058 | 56 | 6 | 20748 | 2239 | $56 \pm 7$ | $30505 \pm 3700$ |
| 2005 | High Flow | 12583 | 39 | 9 | 4923 | 1078 |  |  |
| 2005 | Low Flow | 36577 | 71 | 6 | 26118 | 2034 | $63 \pm 15$ | $31117 \pm 7576$ |
| 2006 | High Flow | 16990 | 56 | 10 | 9578 | 1753 |  |  |
| 2006 | Low Flow | 59424 | 70 | 6 | 41317 | 3853 | $67 \pm 9$ | $50930 \pm 7075$ |
| 2007 | High Flow | 876 | 40 | 8 | 347 | 72 |  |  |
| 2007 | Low Flow | 21033 | 62 | 5 | 13011 | 1134 | $61 \pm 7$ | $13358 \pm 1544$ |
| 2008 | High Flow | 297 | 26 | 9 | 78 | 25 |  |  |
| 2008 | Low Flow | 5642 | 40 | 6 | 2269 | 357 | $40 \pm 7$ | $2347 \pm 424$ |
| 2009 | High Flow | 223 | 28 | 17 | 63 | 39 |  |  |
| 2009 | Low Flow | 4624 | 60 | 8 | 2777 | 376 | $59 \pm 11$ | $2841 \pm 532$ |
| 2010 | High Flow | 2201 | 75 | 22 | 1650 | 473 |  |  |
| 2010 | Low Flow | 42713 | 89 | 6 | 38152 | 2494 | $89 \pm 8$ | $39802 \pm 3613$ |

Note: Population estimates from GrandTab2017.04.07, California Central Valley Chinook Population Database Report, and California Department of Water Resources (unpublished data, contact Jason.Kindopp@water.ca.gov). Mean values and standard deviation (1 $1 \sigma$ ) were calculated using bootstrapping for each recovery location (High Flow and Low Flow channels) and then expanded by the population estimate and combined for each year.

Fig. 7. Timing of ocean entry of fish released from the Feather River hatchery (blue) and wild out-migrating (red) from 2002 to 2010 . The area of each violin represents the proportion of fish out-migrating at that Julian day and is normalized to the total abundance of outmigrants for that year. The black lines represent the interquartile range (first to third quantiles). Hatchery release data for the Feather River Hatchery (FRH) are from Huber and Carlson (2015). Data for "wild" (unmarked) fall-run sized outmigrants are from the USFWS Chipps Island Midwater Trawl. [Colour online.]

hatchery origin. This could in part be attributed to hatchery fish benefitting from improved ocean conditions as well the fishery closure, showing that hatchery and fishery management actions were highly effective at recovering fish stocks following the stock collapse. During this period, FRH release practices were relatively
unchanged, with comparable numbers of juveniles produced and trucked directly to the bay relative to escapement (Fig. 8) (Huber and Carlson 2015). Wild populations likely take longer to recover from stock collapse because their population dynamics are reliant on spawning stock size more so than that of the hatchery. Fur-

Fig. 8. Feather River fall-run in-river escapement estimates split between hatchery- (blue) and wild-origin (red) fish based on otolith (20022010) and California Constant Fractional Marking Program (CFM) data (2011-2012). The time of the salmon stock collapse (2007-2009) is marked by low escapement numbers. Top panel shows the relative stability of hatchery releases over the time series. Escapement data are taken from the GrandTab2016.04.11, California Central Valley Chinook Population Database Report, and hatchery releases numbers are from Huber and Carlson (2015). [Colour online.]

thermore, wild populations are subject to high early life mortality during emigration from CV rivers, which is exacerbated by periods of low-flow conditions during droughts (Zeug et al. 2014). Inriver spawner abundance was greatly reduced during the stock collapse (J. Kindopp, unpublished data, Jason.Kindopp@water.ca.gov), which coincided with a period of drought in California (20072009), and their offspring likely experienced high mortality. Continued monitoring of proportion of hatchery-origin fish as part of this time series will be critical to determine the extent to which the observed pattern represents a fundamental shift towards hatchery dominance. The overwhelming presence of hatchery-origin adults on the spawning grounds would suggest that fall-run Chinook on the Feather River may now be dependent on hatchery fish. Thus, while management strategies for population supplementation of salmon have been successful, they may also facilitate synchronization of hatchery and wild populations (Satterthwaite and Carlson 2015), eroding their resilience to ocean and climate variability.

The dominance of hatchery fish on the Feather River spawning grounds in recent years suggests that interbreeding of hatchery and wild fish is likely and pervasive, particularly given that we did not consider cross-generational hatchery influence (treating all juveniles produced in-river as "wild", independent of parental origin). This is supported by unusually high hatchery contributions on most natural spawning grounds in the CV (Kormos et al.

2012; Palmer-Zwahlen and Kormos 2013, 2015) and a lack of genetic structuring in fall-run CV Chinook salmon, both hatchery and "wild" (Williamson and May 2005). This introgression of hatchery- and wild-origin fish may have reduced fitness (Araki et al. 2008) and weakened the Chinook salmon population portfolio, increasing synchrony among populations and eroding life history diversity and resilience (Carlson and Satterthwaite 2011; Satterthwaite and Carlson 2015). Given the environmental variability inherent to California and predictions of increased frequency of extreme events with future climate change (Cloern et al. 2011; Dettinger et al. 2011), loss of phenotypic diversity could have serious impacts on salmon stock resilience, increasing ecological and economic uncertainty.

The dominance of hatchery-origin fish is not limited to the Feather River. For example, $90 \%$ of in-river spawners on the Mokelumne River in 2004 were classified as hatchery fish (Johnson et al. 2012), and CFM data indicate high hatchery contributions ( $\sim 80 \%-$ 90\%) to natural spawning grounds on Battle Creek, the Merced River, and the Stanislaus River (Kormos et al. 2012; Palmer-Zwahlen and Kormos 2013, 2015). Furthermore, the majority of the ocean fishery is supported by hatcheries, with $90 \%$ of the fishery supported by hatchery fish in 2001 (Barnett-Johnson et al. 2007). There is a growing concern that salmon populations in the CV of California are becoming dependent upon hatchery supplementation, a conservation status recently identified as "mitigated extinction"
(Baumsteiger and Moyle 2017). Further studies are needed basinwide to better understand the role that hatcheries may be playing in "reseeding" in-river populations, masking their declines, and (or) depressing natural production.

It is likely that a number of factors have resulted in hatchery fish effectively replacing wild stocks in the CV, including high and sustained smolt production, largely independent of spawner abundance and freshwater conditions (Huber and Carlson 2015), increased straying rates of trucked fish onto natural spawning grounds (Palmer-Zwahlen and Kormos 2013, 2015), and inflated survival of hatchery smolts as a result of their larger size and the reduction in freshwater mortality for trucked individuals. Such management practices have synchronized the CV salmon stock complex, contributing to a weakened portfolio (Carlson and Satterthwaite 2011; Huber and Carlson 2015; Satterthwaite and Carlson 2015), increased genetic homogenization, and potentially reduced population productivity (Williamson and May 2005). Owing to the reliance on hatchery fish and the high synchrony among the hatcheries in the CV, salmon stock collapses are likely in the future, and compensating for these collapses by increasing hatchery salmon production is likely to prove ineffective (Lindley et al. 2009). We recommend implementation of hatchery practices designed to promote population diversity, such as varying the timing, size, and location of releases to facilitate greater expression of life history diversity in this region and, in turn, its productivity and resiliency (Greene et al. 2010). Moving forward, to reduce the vulnerability of the fishery to over-reliance on hatchery fish and reduce overall extinction risk to wild CV fall-run Chinook salmon, production hatcheries could implement practices that (i) reduce domestication selection through balanced gene flow between hatchery- and natural-origin fish in hatchery broodstock and in rivers (HSRG 2014), (ii) minimize the numbers of hatchery-origin fish interbreeding with wild fish on spawning grounds, and (iii) reduce straying of hatchery adults to support local adaptation in natural salmon populations. CV salmon are at a critical juncture, with many populations close to extinction and facing an increasingly volatile climatic future (Greene et al. 2010; Cloern et al. 2011). Hatcheries can play a key role in the recovery of wild stocks, supplementing the fishery, and the reestablishment of natural areas, but only with cautious and appropriate management.

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# Flow-mediated effects on travel time, routing, and survival of juvenile Chinook salmon in a spatially complex, tidally forced river delta 

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#### Abstract

We evaluated the interacting influences of river flows and tides on travel time, routing, and survival of juvenile late-fall Chinook salmon (Oncorhynchus tshawytscha) migrating through the Sacramento-San Joaquin River Delta. To quantify these effects, we jointly modeled the travel time, survival, and migration routing in relation to individual time-varying covariates of acoustic-tagged salmon within a Bayesian framework. We used observed arrival times for detected individuals and imputed arrival times for undetected individuals to assign covariate values in each reach. We found travel time was inversely related to river inflow in all reaches, yet survival was positively related to inflow only in reaches that transitioned from bidirectional tidal flows to unidirectional flow with increasing inflows. We also found that the probability of fish entering the interior Delta, a low-survival reach, declined as inflow increased. Our study illustrates how river inflows interact with tides to influence fish survival during the critical transition between freshwater and ocean environments. Furthermore, our analytical framework introduces new techniques to formally integrate over missing covariate values to quantify effects of time-varying covariates.


## Introduction

Anadromous salmonids have evolved diverse life history strategies that capitalize on spatial and temporal variation in their habitat to maximize productivity. Understanding how salmonids use habitat over space and time can provide insight into population dynamics and help to identify particularly sensitive stages in their life history. Regulated rivers influence migrations of anadromous salmonids by altering the timing, magnitude, variation, and constituents of river discharge (e.g., temperature, turbidity), which in turn can affect their survival (Raymond 1988; Smith et al. 2003). Thus, interest often centers on how regulation of river flow affects survival of juvenile salmonids at different locations and times (Skalski et al. 2002; Michel et al. 2015).

Juvenile Chinook salmon (Oncorhynchus tshawytscha) in the Central Valley of California, USA, emigrate from natal tributaries of the Sacramento River through the Sacramento-San Joaquin River Delta (henceforth, "the Delta"), a network of natural and manmade channels linking the Sacramento and San Joaquin rivers to San Francisco Bay and the Pacific Ocean (Fig. 1). The Delta is the hub of California's water delivery system, providing agricultural and domestic water that supports California's economy, the eighth largest in the world (Healey et al. 2016). Water from the Sacramento River is diverted from the North through natural channels and gated man-made channels to the South where large pumping stations "export" water from the Delta in canals (Fig. 1). As juvenile salmon enter the Delta, they distribute among its complex channel network where they are subject to channel-specific abiotic and biotic factors that influence their migration timing, growth, and survival. For example, fish that enter the interior Delta, the region to the south of the mainstem Sacramento River (reach 8 in Fig. 1), survive at lower rates than fish migrating through northerly routes, likely owing to longer
travel times, longer travel distances, higher predation rates, and entrainment at the pumping stations (Brandes and McLain 2001; Newman and Brandes 2010; Perry et al. 2010, 2013).

Survival of juvenile salmon has been positively related to river discharge at the Delta-wide scale (Kjelson et al. 1982; Kjelson and Brandes 1989; Newman and Rice 2002; Newman 2003), but the underlying factors driving this relationship remain unclear. Low river discharge has been associated with a high proportion of fish entering the interior Delta, thereby decreasing overall survival by subjecting a larger fraction of the population to low survival probabilities (Perry et al. 2015). What remains unclear is the extent to which within-reach survival contributes to the overall flow-survival relationship. Is survival related to discharge in all reaches, or do a few key reaches drive the overall flow-survival relationship? Given that the Delta transitions from unidirectional flow in its upper reaches to tidally driven bidirectional flows in lower reaches, we hypothesized that the reach-specific relationship between inflow and survival could vary along this gradient. Understanding exactly which reaches contribute to the overall flow-survival relation will help researchers to focus on specific mechanisms driving this relationship and help managers to target specific actions to increase survival.

Here, we analyze acoustic telemetry data on juvenile Chinook salmon from 17 distinct release groups collected from two studies conducted between 2007 and 2011 (Table 1) to understand how reach-specific travel time, migration routing, and survival vary among reaches in the Sacramento-San Joaquin River Delta. Because each release group spread out over time as they migrated through the Delta, individuals entered a given reach over a wide range of environmental conditions. Our interest therefore centered on quantifying factors affecting individual variation in survival. However, time-varying individual covariates are a vexing problem in conventional mark-recapture models (e.g., maximum likelihood estimation performed
in Program Mark, White and Burnham 1999) because the value of the covariate is unknown when an individual is undetected, rendering the likelihood analytically intractable in most cases (but see Catchpole et al. 2008). Therefore, we developed a Bayesian hierarchical model that jointly modeled both travel times and survival of juvenile salmon. The travel time model was used to impute arrival times of undetected fish in each reach, which allowed us to define covariate values based on imputed arrival times for undetected individuals. We then used Markov Chain Monte Carlo techniques to integrate the likelihood over the missing covariate values while simultaneously estimating parameters associated with both travel time and survival.

## Methods

## Study area and telemetry system

The telemetry system was designed to accommodate requirements of a multistate markrecapture model that estimated reach- and route-specific survival for nine discrete reaches and four primary migration routes through the Delta (Perry et al. 2010; Figs. 1 and 2). The nine reaches separate the Delta into the three hydrodynamic zones: 1) riverine reaches with unidirectional flows and the least influence of tidal forcing (reaches $0-2$ ), 2) transitional reaches that shift from unidirectional flow to tidally driven bidirectional flows as river flow entering the Delta decreases (reaches 3-6), and 3) tidal reaches with bidirectional flows regardless of the amount of river flow entering the Delta (reaches 7-8, Figs. 1 and 3). These nine reaches comprise four distinct migration routes that constitute the states of the multistate model: the Sacramento River (Route A = reaches 1, 2, 4, and 7), Sutter and Steamboat Slough (Route B = reaches 1, 3, and 7), the Delta Cross Channel (Route $C=$ reaches $1,2,6$, and 8 ), and Georgiana Slough (Route D = reaches 1, 2, 5, and 8; Figs. 1 and 2).

Each telemetry station consisted of single or multiple tag-detecting monitors (Vemco Model VR2, Amirix Systems, Inc., Halifax, Nova Scotia, Canada), depending on the number of monitors needed to maximize detection probabilities at each station. Migration routes $\mathrm{A}, \mathrm{B}, \mathrm{C}$, and D were monitored with $7,1,1$, and 2 telemetry stations, respectively, labeled according to migration route $r$ at sampling occasion $j$ (Figs. 1 and 2). Sampling occasion was defined based on the $j$ th telemetry station within the mainstem Sacramento River, with the upstream release site defined as occasion one. Migrating juvenile salmon first arrive at Sutter and Steamboat Slough $\left(B_{3}\right)$, which diverges from the Sacramento River at the first river junction and converges again with the Sacramento River upstream of $\mathrm{A}_{5}$ (Figs. 1 and 2). Fish remaining in the Sacramento River then pass the Delta Cross Channel $\left(\mathrm{C}_{4}\right)$, a man-made gated canal that diverts fish, when its gates are open, into reach 6 and subsequently into the interior Delta (reach 8). The Delta Cross Channel is used to control salinity at the water pumping stations, undergoes mandatory closures for fisheries protection in mid-December each year, and also closes when Sacramento River flow exceeds $708 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}\left(25,000 \mathrm{ft}^{3} \cdot \mathrm{~s}^{-1}\right)$. Fish then pass Georgiana Slough $\left(\mathrm{D}_{4}\right)$, a natural channel (reach 5) that also leads to the interior Delta (reach 8). All routes then converge at Chipps Island $\left(\mathrm{A}_{6}\right)$, the terminus of the Delta. With this configuration, survival to site $\mathrm{A}_{6}$ is confounded with detection probability at the last telemetry station. Therefore, to estimate survival to $\mathrm{A}_{6}$, we pooled detections from numerous tag detecting monitors downstream of $\mathrm{A}_{6}$ in San Francisco Bay for estimating detection probability at Chipps Island.

Although there are numerous possible migration pathways, we focused on these four routes because management actions likely have the largest influence on movement and survival among these routes. For example, fish may enter the interior Delta from the Sacramento River through either Georgiana Slough or the Delta Cross Channel, where they subsequently become
vulnerable to migration delays and entrainment at the water pumping projects (Perry et al. 2010; Newman and Brandes 2010). Steamboat and Sutter Slough is an important migration route because fish using this route bypass the Delta Cross Channel and Georgiana Slough (Fig. 1) thereby avoiding the interior Delta. Thus, monitoring these primary migration routes provides information about the likely ultimate fate of individuals.

## Fish tagging and release

All juvenile late fall Chinook salmon were obtained from the Coleman National Fish Hatchery in Anderson, California. Release groups were defined based on release timing and data source, with the exception of release group 3, which was pooled over a longer period of release times owing to small sample size (Table 1). All fish other than release group 1 were tagged with a 69 kHz acoustic tag weighing 1.58 g (Vemco Model V7-2L-R64K, Amirix Systems, Inc., Halifax, Nova Scotia, Canada) transmitting either every $30-90$ s (release groups 1-3) or 15-60 s (release groups 4-17). Battery life of these transmitters ranged from 98-749 d based on tests conducted by Michel et al. (2015). Fish from release group 1 were tagged with an acoustic tag weighing 1.44 g , which had an expected battery life of 70 d (Vemco Model V7-2L-R64K).

Most juvenile salmon were surgically tagged at the hatchery and then transported to release sites, but fish from release groups 8 and 11 were tagged at release sites. Fish were randomly selected and those $\geq 140 \mathrm{~mm}$ fork length were retained for tagging to maintain tag burden below $6 \%$ of the fish weight. Fish tagged by Michel et al. (2015) were held at the hatchery for 24 h following surgery, transported to release sites, and held in-river for $1-3 \mathrm{~h}$ prior to release. Fish tagged by Perry et al. $(2010,2012,2013)$ fish were transported to release sites, held in-river at release sites for 24 h , and then released into either the Sacramento River near

Sacramento, CA ( $n_{\mathrm{A} 1}$ ) or Georgiana Slough ( $n_{\mathrm{D} 4} ;$ Fig. 1 and 2). Fish were released into Georgiana Slough to increase the number of fish entering the interior Delta (reach 8) and improve precision of survival estimates for that region. For the Michel et al. (2015) study, fish were released well upstream of the Delta, at four locations in the Sacramento River (Table 1). In most migration years, two releases were made; one in December and another in January. Releases in December occurred prior to seasonal closure of the Delta Cross Channel gates, which typically occurs on December 15; whereas the Delta Cross Channel gates were closed for all January releases. Further details of tagging and release protocols can be found in the citations listed in Table 1.

## Screening for false positive detections and predators

Telemetry data were screened for false positive detections by first summarizing data into detection events defined by the number of consecutive detections from an individual tag within a 30-minute period at a given telemetry station. Any detection event with at least two detections at a given location was considered as valid. Detection events with a single detection were considered valid if the detection was consistent with the entire spatiotemporal detection history of the individual's tag (e.g., a single detection was preceded by an upstream detection and proceeded by a downstream detection). Otherwise, single detections were considered false positives and removed from analysis.

Tags that may have been consumed by predators were identified by adapting the methods of Gibson et al. (2015), which consisted of several steps. First we calculated five movement metrics from tag detections that quantified differences in behavioral patterns between live tagged smolts and tagged smolts that had been consumed by predators such as striped bass (Morone
saxatilis), smallmouth bass (Micropterus dolomieu), largemouth bass (Micropterus salmoides), and spotted bass (Micropterus punctulatus). The metrics included: 1) the mean rate of downstream movement calculated as the shortest channel distance between consecutive detections of downstream movements divided by the elapsed time between detections, 2) the number of consecutive detection events occurring at the same location, 3 ) the cumulative distance travelled divided by the total number of days spent in the study area, 4) the number of transitions between telemetry stations that were deemed to be only possible by a predator (i.e., movement upstream against the flow), and the total time in the array from the time of release to the time of last detection.

Next we used hierarchical cluster analysis to group each tag by the multivariate characteristics of the five metrics. We used the hclust package in $R$ ( R Core Team 2015) and divided the tags in three groups based on the dendrogram resulting from hierarchical clustering using Ward's minimum variance method (Ward 1963; Gibson et al. 2015). We then selected the group whose movement characteristics were most consistent with that of predator-like behavior (i.e., upstream movement against flow, long residence times near receivers, and low average distance travelled per day). We examined each tag's time series of movement metrics to identify if and when the tag transitioned from smolt-like to predator-like behavior. The detection history was then truncated at this point in the detection history. Overall, $17 \%$ percent of tags were flagged for review based on the movement metrics, and $11 \%$ percent exhibited predator-like behavior that required truncation of their capture history.

## Structure of the mark-recapture model

The multistate mark-recapture model estimates three types of parameters from detections of acoustic-tagged juvenile Chinook Salmon: $S_{r, j}$ is the probability of surviving from a telemetry station within route $r$ at sampling occasion $j$ to the next downstream telemetry station; $\Psi_{r, s, j}$ is the probability of entering route $s$ from route $r$ at sampling occasion $j$, conditional on surviving to occasion $j$ (henceforth, routing probability); and $P_{r, j}$ is the probability of detecting a tagged fish at a telemetry station on sampling occasion $j$ within route $r$, conditional on fish surviving to occasion $j$ (Fig. 2). In the parlance of multistate mark-recapture models, the routes constitute the states, the routing probabilities represent the state transition probabilities, and survival and detection probabilities are conditioned on migration route (i.e., conditioned on state).

In addition, our modeling framework includes an auxiliary model for travel times, which we used to impute arrival times of undetected individuals in each reach for the purposes of assigning daily covariate values. This model estimates two travel time parameters associated with lognormally distributed travel times: $\mu_{r, j}$ is the mean of log-travel times from a telemetry station in route $r$ at sampling occasion $j$ to the next downstream telemetry station, and $\sigma_{r, j}^{2}$ is the variance of the travel times. Because reaches $1-8$ are associated with a unique $r, j$ combination (route, sampling occasion) we generally refer to travel time and survival parameters as being reach-specific (Fig. 1 and 2).

To understand how both migration routing and reach-specific survival contribute to overall survival through the Delta, we model the underlying parameters as functions of covariates and then reconstruct the overall relationship from these component parts. Overall survival through the Delta was reconstructed from the individual components as:

$$
\begin{equation*}
S_{\text {Detla }}=\sum_{r \in\{A, B, C, D\}} \Psi_{r} S_{r} \tag{1}
\end{equation*}
$$

where $S_{r}$ is the survival from telemetry stations $\mathrm{A}_{2}$ to $\mathrm{A}_{6}$ (i.e., from the entrance to the exit of the Delta) for fish taking migration route $r$, and $\Psi_{r}$ is the total probability of a fish taking route $r$.

Thus, $S_{r}$ is the product of reach-specific survival probabilities that trace a unique migration route through the Delta (e.g., $S_{\mathrm{D}}=S_{\mathrm{A} 2} S_{\mathrm{A} 3} S_{\mathrm{D} 4} S_{\mathrm{D} 5}$ ), and $\Psi_{r}$ is the product of routing probabilities along that route (e.g., $\Psi_{\mathrm{D}}=\Psi_{\mathrm{AA} 3} \Psi_{\mathrm{AD} 4}$, Perry et al., 2010).

## Time-varying individual covariates

We hypothesized that river discharge affected migration routing, travel times, survival, and detection probabilities. Mean daily discharge varies among the nine reaches owing to the distribution of total discharge among the Delta's channel network. However, tidally averaged net discharge in most reaches is a direct function of 1) river flows entering the Delta (as measured in the Sacramento River at Freeport located near telemetry station $\mathrm{A}_{2}$ in Fig. 1), and 2) whether the Delta Cross Channel Gate is open or closed (Fig. A1 ${ }^{1}$ ). Furthermore, as river inflow increases, tidal fluctuations are dampened in all but reaches 7 and 8 (Fig. 3). Therefore, we used river discharge at Freeport $(Q)$ and the position of the Delta Cross Channel gate ( $G=1$ or 0 for gates open or gates closed, respectively) as an index of variation in reach-specific mean discharge affecting migration routing, travel times, survival, and detection probabilities. Specifically, time-varying individual covariates $Q_{d}$ and $G_{d}$ were assigned based on the day $d$ when the $i$ th individual passed a telemetry station in route $r$ at sampling occasion $j$.

[^4]We modeled $\mu$, the log-mean of the travel time distribution, as a linear function of individual time-varying covariates:

$$
\begin{equation*}
\mu_{i, r, j}=\alpha_{0, r, j}+\alpha_{1, r, j} Q_{d}+\alpha_{2, r, j} G_{d}+z_{\mu, n, r, j} \xi_{\mu, r, j} \tag{2}
\end{equation*}
$$

where $r, j$ indexes the route and occasion where individuals entered reaches $0, \ldots, 8$ (Fig. 1 and 2), $\mu_{i, r, j}$ is the log-mean travel time for individual $i$ in each reach, $\alpha_{0, r, j}$ is the intercept, $\alpha_{1, r, j}$ is the slope for the effect of discharge on $\mu$, and $\alpha_{2, r, j}$ is the effect of Delta Cross Channel gate position on $\mu$. We modeled $\sigma_{r, j}$, the variance parameter of the log-normal travel time distribution, as a constant for all individuals within a reach. In addition, $\alpha_{2, r, j}$ was set to zero for reaches located upstream of the of the Delta Cross Channel (i.e., for reaches $0,1,2,3$, and 6 ).

Given that discrete groups of fish were released in different months, years, and locations, we expected considerable variation in release-specific travel time, survival, and routing over and above variation that could be accounted for by covariates in the model. Extra variation among release groups was structured as a non-centered random effect where $z_{\mu, n, r, j}$ in Eq. 2 is a standard normal deviate for the $n$th release group entering each reach, $\xi_{\mu, r, j}$ is the standard deviation of the random effect in each reach, and their product is the deviation of each release group from the mean, conditional on the covariates. We used a non-centered random effect to reduce autocorrelation and speed convergence of the model fitting routine (Papaspiliopoulos et al. 2007; Monnahan et al. 2017).

Reach-specific survival was modeled as a logistic function using the same linear structure as travel time:

$$
\begin{equation*}
\operatorname{logit}\left(S_{i, r, j}\right)=\beta_{0, r, j}+\beta_{1, r, j} Q_{d}+\beta_{2, r, j} G_{d}+\beta_{3} \ell_{i}+z_{S, n, r, j} \xi_{S, r, j} \tag{3}
\end{equation*}
$$

where $\operatorname{logit(}(\cdot)$ is the logit link function, $\ell_{i}$ is the fork length of individual $i, \beta_{3}$ is the slope for the effect of fork length on survival, and all other coefficients are defined as in Eq. 2 except with respect to survival. In this model, survival is constant among individuals that enter a given reach on a particular day. Travel time influences survival only through its effect on arrival times to a given telemetry station, which determines the discharge that individuals experienced when they entered a given reach.

We modeled three routing probabilities as a function of covariates: $\Psi_{\mathrm{AB3}}, \Psi_{\mathrm{AC}}$, and $\Psi_{\mathrm{AD} 4 \mid \mathrm{C}^{\prime}}$. Here, $\Psi_{\mathrm{AB3}}$ is the probability of entering Sutter and Steamboat Slough (route B) from the Sacramento River (route A) at sampling occasion 3, $\Psi_{\mathrm{AC} 4}$ is the probability of entering the Delta Cross Channel (route C) from the Sacramento River at sampling occasion 4, and $\Psi_{\mathrm{AD}^{\prime} / \mathrm{C}^{\prime}}$ is the probability of entering Georgiana Slough (route D) from the Sacramento River, conditional on not having entered the Delta Cross Channel ( $\mathrm{C}^{\prime}$ ). Since routing probabilities must sum to one at each of the two river junctions, the unconditional probability of entering Georgiana Slough ( $\left.\Psi_{\mathrm{AD} 4}\right)$ at sampling occasion 4 is $\left(1-\Psi_{\mathrm{AC} 4}\right) \Psi_{\mathrm{AD} 4 \mid \mathrm{C}^{\prime}}$.

We model routing probabilities using a generalized logistic function:

$$
\begin{equation*}
\Psi_{i}=L+\frac{U-L}{1+\exp \left(-\left(\gamma_{0}+\gamma_{1} Q_{d}+\gamma_{2} G_{d}+z_{\gamma, n} \xi_{\Psi}\right)\right)} \tag{4}
\end{equation*}
$$

where $\Psi_{i}$ is one of the three routing probabilities described above for individual $i, L$ is the lower limit of $\Psi_{i}, U$ is the upper limit of $\Psi_{i}$, and all other parameters are described as in Eq. 2 except with respect to routing. The parameters $U$ and $L$ allow the logistic function to take on values other than one or zero for upper and lower limits, respectively. This equation reduces to the standard inverse logit function by setting $U=1$ and $L=0$. We used the generalized logistic
function because we expected routing probabilities to follow a relationship similar to that between total discharge $(Q)$ and the fraction of discharge entering each route. As these channels transition from bidirectional tidal flows to unidirectional flows with increasing total discharge, the fraction of discharge entering a route either increases (Sutter and Steamboat sloughs) or decreases (Georgiana Slough) with discharge before leveling off at a constant fraction of discharge (Fig. A2). Therefore, for Sutter and Steamboat Slough ( $\Psi_{\text {AB3 }}$ ) we set $L=0$ and $\gamma_{2}=0$; for the Delta Cross Channel $\left(\Psi_{\mathrm{AC} 4}\right)$ we set $L=0, U=1$, and $\gamma_{2}=0$; and for Georgiana Slough ( $\Psi_{\mathrm{AD} 4^{\prime} \mid \mathrm{C}^{\prime}}$ ) we set $U=1$.

We hypothesized that increases in discharge could reduce detection probabilities by increasing acoustic noise and by increasing the speed at which juvenile salmon pass telemetry stations. In addition, many telemetry stations were monitored each year with different hydrophones, varying numbers of hydrophones, and different spatial configurations that could have influenced detection probability. Therefore, we modeled these effects on detection probability as linear on the logit scale:

$$
\begin{equation*}
\operatorname{logit}\left(P_{i, r, j}\right)=\theta_{0, r, j, y}+\theta_{1, r, j} Q_{d} \tag{5}
\end{equation*}
$$

where $\theta_{0, r, j, y}$ is an intercept for year $y$ at occasion $j$ within route $r$, and $\theta_{1, r, j}$ is the slope for the effect of river discharge on detection probability at occasion $j$ in route $r$.

## Complete data likelihood

To estimate model parameters as a function of time-varying individual covariates, we used the complete data likelihood of the multistate model within a Bayesian framework. The complete data likelihood proceeds as if there were no missing values by augmenting the observed data with the unobserved missing data and treating the missing data as additional model
parameters to be estimated (King et al. 2010; Link and Barker 2010). This approach relies on using an appropriate probability model for imputing missing covariate values and then constructing the joint likelihood of the mark-recapture model parameters, the covariate model parameters, and the missing data (Bonner and Schwarz 2004). To impute missing covariate values for non-detected individuals whose arrival times are unknown, we model arrival times by estimating parameters of the distribution of travel times through each reach.

The observed data for each individual required to estimate model parameters include 1) the detection history, 2) cumulative travel times, 3) reach-specific travel times, and 4) covariates linked to the fish's arrival time in each reach. A "detection history" is the alpha-numeric vector $h_{i}$ indicating whether individual $i$ was detected in route $r$ at occasion $j\left(h_{i, j}=\mathrm{A}, \mathrm{B}, \mathrm{C}\right.$, or D$)$ or not detected at occasion $j\left(h_{i, j}=0\right)$. The detection history compactly represents each fish's detection and movement history through the telemetry network. For example, the detection history A0ADD00 indicates a fish that was released into the Sacramento River $\left(h_{i, 1}=\mathrm{A}\right)$ and was not detected at $\mathrm{A}_{2}$ but was detected at $\mathrm{A}_{3}\left(h_{i, 2: 3}=0 \mathrm{~A}\right)$, indicating it remained in the Sacramento River at its junction with Sutter and Steamboat sloughs. This fish was then detected entering Georgiana Slough at $\mathrm{D}_{4}$ and once more at $\mathrm{D}_{5}$ before never being detected again ( $h_{i, 4: 7}=\mathrm{DD} 00$ ). Associated with the observed detection history of each individual is the vector of observed cumulative travel times $T_{i}$. For example, if $h_{i}=\mathrm{A} 0 \mathrm{ADD} 00$ then $T_{i}=\left(T_{i, 1}, \mathrm{NA}, T_{i, 3}, T_{i, 4}, T_{i, 5}\right.$, NA, NA) where $T_{i, 1}=0, T_{i, j}$ is the time from release to detection at a telemetry station at sampling occasion $j$, and $T_{i, j}$ is missing (NA) when an individual is not detected. Time-varying covariate values $x_{i, j}$ defined based on arrival date in each reach are missing (NA) whenever an individual is not detected. Thus, for A0ADD00, $\quad \lambda=\left(x_{i, 1}\right.$, NA, $x_{i, 3}, x_{i, 4}, x_{i, 5}$, NA, NA $)$.

Observed reach-specific travel times $t_{i, r, j}$ for individual $i$ in route $r$ at occasion $j$ are obtained by taking the consecutive differences of the cumulative travel times. For A0ADD00, $t_{i}=(\mathrm{NA}, \mathrm{NA}$, $\left.t_{i, \mathrm{~A} 3}, t_{i, \mathrm{D4}}, \mathrm{NA}, \mathrm{NA}\right)$. Note that $t_{i, r, j}$ is observed only when fish are detected at consecutive telemetry stations whereas $T_{i, j}$ is defined whenever a fish is detected.

Adapting the notation of the King et al. (2010), the complete data likelihood augments the observed detection history, $h_{i}$, by imputing the latent (unobserved) states when individuals are not detected:

$$
z_{i, j}= \begin{cases}h_{i, j} & \text { if } h_{i, j} \neq 0  \tag{6}\\ g_{i, j} & \text { if } h_{i, j}=0\end{cases}
$$

where $g_{i, j}$ is the latent state of unobserved individual $i$ at detection occasion $j, z_{i, j}$ is the state of individual $i$ at detection occasion $j$ (whether detected or non-detected), and $z_{i}$ is the complete state history for individual $i$. Although death can never be directly observed in detection history $h_{i}$, death is included as a latent state such that $g_{i, j} \in(A, B, C, D, \dagger)$ where $\dagger$ is the death state.

The complete data likelihood is the product of three conditional likelihoods: 1) a Bernoulli distribution for detection at occasion $j$ given survival to occasion $j$ in state $r, 2$ ) a Bernoulli distribution for survival from occasion $j$ to $j+1$ in state $r$ given survival to occasion $j$, and 3) a generalized Bernoulli distribution (i.e., a multinomial distribution for a single observation) for the probability of moving from state $r$ at occasion $j$ to state $s$ at occasion $j+1$ given survival to occasion $j+1$ :
(7) $L[S, \Psi, P \mid h, g]=\prod_{i=1}^{N} \prod_{j=F_{i} r \in R_{j}}^{J-1}\left[P_{i, r, j+1}^{u_{i, j+l}}\left(1-P_{i, r, j+1}\right)^{v_{i, j+1, r}}\right]\left[S_{i, r, j}^{w_{i, j, r}}\left(1-S_{i, r, j}\right)^{w_{i, j, r, t}}\right] \prod_{s \in R_{j+1}} \Psi_{i, r, s, j, j}^{w_{i, j, s}}$
where
$F_{i}$ is the occasion of release for individual $i$,
$R_{j}$ is the set of states, excluding the death state, available to an individual in state $r$ at occasion $j$
(Fig. 2),
$u_{i, j, r}=I\left(h_{i, j}=r\right)$ and $I(\cdot)$ is an indicator function resolving to one if individual $i$ is detected in state $r$ at occasion $j$ and zero otherwise,
$v_{i, j, r}=I\left(g_{i, j}=r\right)$ is one if individual $i$ is imputed to be in state $r$ at detection occasion $j$ and zero otherwise, $w_{i, j, r, s}=I\left(z_{i, j}=r, z_{i, j+1}=s\right)$ is one if individual $i$ is in state $r$ at detection occasion $j$ and in state $s$ at detection occasion $j+1$ and zero otherwise.

Note that $w_{i, j, r,\rangle}$ is one if individual $i$ dies between $j$ and $j+1, w_{i, j, r, r}=\sum_{s \in R_{j+1}} w_{i, j, r, s}$ is one if the individual survives, and the dot represents any state but the death state.

We modeled reach-specific travel times using a lognormal distribution because travel times of migrating juvenile salmon are typically right-skewed, and the lognormal distribution often fits travel time data well (Muthukumarana et al. 2008). Missing travel times (i.e., $t_{i, r, j}=$ NA) are imputed from a lognormal distribution subject to the constraint

$$
\begin{equation*}
T_{i, j+K}=T_{i, j}+\sum_{k=0}^{K-1} t_{i, r, j+k}^{\text {mis }} \tag{8}
\end{equation*}
$$

where $T_{i, j}$ and $T_{i, j+K}$ are observed cumulative travel times, $t_{i, r, j}^{\mathrm{mis}}$ are missing reach-specific travel times between occasions $j$ and $j+K$, and the $K$ is the number of missing reach-specific travel times between $T_{i, j}$ and $T_{i, j+K}(K=2, \ldots, J-1)$. Since the sum of missing travel times are
constrained to be equal to $T_{i, j+K}-T_{i, j}$, this constraint on imputed travel times imposes a form of left censoring, thereby providing additional information to the parameter estimation.

Given observed and imputed travel times, the complete data likelihood for the travel time data is:

$$
\begin{equation*}
L\left[\mu, \sigma \mid t^{\mathrm{obs}}, t^{\mathrm{mis}}\right] \propto \prod_{i=1}^{N} \prod_{j=F_{i}}^{J-1} \prod_{r \in R_{j}} \frac{1}{\sigma_{i, r, j} t_{i, r, j}} \exp \left(\frac{-\left(\ln \left(t_{i, r, j}\right)-\mu_{i, r, j}\right)^{2}}{2 \sigma_{i, r, j}^{2}}\right) \tag{9}
\end{equation*}
$$

where $t_{i, r, j}$ is the observed $\left(t^{\mathrm{obs}}\right)$ or imputed $\left(t^{\text {mis }}\right)$ travel time for individual $i$ in state $r$ at detection occasion $j$ and $\sigma_{i, r, j}^{2}$ is the variance of the lognormal travel time distribution for individual $i$ in state $r$ at occasion $j$. We estimated $\sigma_{i, r, j}^{2}$ as a constant over all individuals for each reach.

## Other parameter constraints

In addition to constraining parameters as a function of covariates, a number of other constraints were imposed owing to telemetry station outages, multiple release locations, and parameter identifiability issues.

For reach 0, individuals were either released at Sacramento (rkm 172), at rkm 191, or well upstream of these locations (>rkm 191, Table 1). For fish released well upstream of Sacramento, we included in the analysis only those that were detected by telemetry stations in the vicinity of Sacramento or at a telemetry station located near the Feather River at rkm 204 (see "number analyzed" in Table 1). To account for the effect of detection or release upstream of Sacramento on travel times through reach 0, we included coefficients that estimated the
difference in intercepts for fish detected at rkm 204 or released at rkm 191 relative to those detected or released at Sacramento (rkm 172).

We treated the first reach after release as an "acclimation" reach to allow fish to recover from handling and release procedures before drawing inferences about travel time and survival (reach 0 for releases at Sacramento and reach 5 for releases in Georgiana Slough). Therefore, fish released directly into Georgiana Slough (rkm 115, Table 1) were modeled with unique coefficient values in reach 5 relative to fish that entered reach 5 volitionally from upstream locations. Coefficients based only on fish that entered reach 5 volitionally were then used for inference about travel time and survival in reach 5.

Telemetry station $\mathrm{A}_{3}$ was not deployed until January 2007, affecting release 1, and was not deployed between December 2007 and March 2008, affecting releases 4-7. To incorporate the effect of these receiver outages, detection probability was set to zero for fish that were imputed to arrive at site $\mathrm{A}_{3}$ during these time periods. In preliminary analysis, we found coefficients associated with survival in reach 2 were weakly identifiable (i.e., large credible intervals), and we identified undue influence of the prior distribution on $U$, the upper limit of the logistic function for routing into Sutter and Steamboat Slough ( $\Psi_{\text {AB3 }}$ ). Both issues were likely driven by the extended receiver outages at telemetry station $A_{3}$. Therefore, we set all survival coefficients for reach 2 equal to those for reach 1 and estimated common slopes, intercepts, and random-effects parameters. This constraint was supported by previous analyses showing similar survival between reaches 1 and 2 (Perry 2010; Perry et al. 2010).

For routing into Sutter and Steamboat Slough, we included auxiliary data from an independent telemetry study to bolster parameter estimates associated with $\Psi_{\text {AB3 }}$ (California Department of Water Resources 2016, Romine et al. 2017). Of 4,528 acoustically tagged
juvenile late-fall Chinook salmon released at Sacramento between 1 March and 15 April 2014, 3,548 fish were detected at the junction of the Sacramento River and Sutter and Steamboat Slough. We modeled this binary data ( $1=$ Sutter and Steamboat Slough, $0=$ Sacramento River $)$ using a Bernoulli likelihood with probability $\Psi_{\text {AB3 }}$ and jointly estimated the parameters of Eq. 4 for $\Psi_{\text {AB3 }}$ over both data sets.

Last, unique detection probabilities could not be estimated at the entrance to the Delta Cross Channel (telemetry station $\mathrm{C}_{4}$ ) and Georgiana Slough (telemetry station $\mathrm{D}_{4}$ ) owing to a single downstream detection site common to both reaches (telemetry station $\mathrm{D}_{5}$ ). Therefore, a common set of coefficients for detection probability were estimated for sites $\mathrm{D}_{4}$ and $\mathrm{C}_{4}$.

Prior distributions, parameter estimation, and goodness of fit
Prior distributions for parameters associated with routing, survival, and detection were based on the default priors for logistic regression recommended by Gelman et al. (2013). First, all continuous covariates were scaled to have mean zero and standard deviation 0.5 (for discharge, $Q$, mean $=610.1 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}, \mathrm{SD}=407.1 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$; for fork length, $\ell$, mean $=155.1 \mathrm{~mm}, \mathrm{SD}$ $=10.8 \mathrm{~mm})$. Next, slope parameters associated with routing, survival, and detection were drawn from a Student's $t$ distribution with a mean of zero, standard deviation of 2.5, and 7 degrees of freedom. Intercepts associated with routing, survival, and detection were drawn from a Cauchy $(0,10)$ distribution. We used a $\operatorname{Normal}(0,1)$ distribution truncated at zero as the prior distribution for $\xi_{\text {S }}$ and $\xi_{\Psi}$ (Gelman et al. 2013). Last, a $\operatorname{Uniform}(0,1)$ prior was used for $L$ and $U$. For travel time parameters, slopes and intercepts for $\mu$ were drawn from a $\operatorname{Normal}(0,10)$ prior distribution, and $\xi_{\mu}$ and $\sigma$ were drawn from a $\operatorname{Uniform}(0,10)$ prior.

We coded the model in the Markov Chain Monte Carlo (MCMC) software package JAGS (http://mcmc-jags.sourceforge.net/) as called from R (Denwood 2016), which allowed us to simultaneously estimate all model parameters and impute missing data (see Supplement $\mathrm{B}^{1}$ ). JAGS uses Gibbs sampling and Metropolis-Hastings methods to sequentially update each parameter value, conditional on the current value for all other parameters. We ran three MCMC chains in JAGS each for 50,000 iterations that consisted of a 1000-iteration adaptation phase and an additional 30,000 -iteration burn-in phase. The final 20,000 iterations were thinned at a rate of 1 in 20 resulting in 1,000 iterations from each chain that were used to form the joint posterior distribution of the parameters. With these MCMC settings, the model took five days to run (10,000 iterations per day) on desktop computer with a 3.5 GHz processer and 64 GB of RAM.

We inspected trace plots of each MCMC chain and used the $\hat{R}$ statistic to assess convergence of the posterior for each parameter, where $\hat{R}<1.1$ indicates convergence (Gelman et al. 2013). We then performed posterior predictive checks to assess goodness of fit by simulating replicated data from the joint posterior distribution. We used the joint log-likelihood of the capture histories (Eq. 7) and travel times (Eq. 9) as a goodness of fit statistic, which was calculated for both observed and replicated data for each draw in the joint posterior distribution. We calculated the probability that the observed data could have been generated by the model by calculating the proportion of times that the likelihood of the observed data was greater than that for replicated data. Often referred to as a Bayesian $p$-value, a probability $>0.95$ or $<0.05$ is typically taken as evidence of lack of fit (Gelman et al. 2013).

## Results

The $\hat{R}$ statistics indicated that the Markov Chains converged to a stable stationary distribution. Of the 155 estimated parameters, $\hat{R}$ was less than 1.1 for all but one parameter ( $\theta_{0, \mathrm{D} 4,4}$, the intercept for $P$ at telemetry station $\mathrm{D}_{4}$ in year 4 ), and its $\hat{R}$ was 1.115 , just slightly higher than the standard cutoff value. In addition, we found no evidence of lack of fit; $54.4 \%$ of log-likelihood values for the observed data were greater than those for replicated data, indicating that the observed data were just as likely to have been generated by the model compared to replicated data that was known to have been generated by the model.

Daily inflow to the Delta varied widely over the study period, ranging from $193 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$ to $2,180 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$ (Fig. 3), which encompassed the $1^{\text {st }}$ to $95^{\text {th }}$ percentile of daily discharge in the 69year flow record for the December through March migration period. Inflows influenced detection probabilities, travel time, survival, and routing. We found that discharge had a negative effect on detection probabilities at most telemetry stations, but the magnitude of the effect declined from the upper to lower Delta as tidal influence increased (Fig. A3 ${ }^{1}$ ). In general, detection probabilities were greater than 0.8 at most telemetry stations when flows were below $1,000 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$, but decreased at higher flows with the rate of decrease varying among years and telemetry stations (Fig. A4 ${ }^{1}$ ).

Most survival and travel time parameters associated with reach 6 (the Delta Cross Channel) exhibited wide credible intervals because only 6 release groups were released prior to mid-December when the Delta Cross Channel undergoes mandatory closures for fish protection (Fig. 4). Consequently, there was relatively little data from which to estimate the effects of river discharge on travel time and survival for reach 6.

For all other reaches, we found that median travel time was influenced by river flow. Posterior distributions for the effect of flow on travel time $\left(\alpha_{1}\right)$ were negative and credible
intervals excluded zero, indicating that increases in river flow reduced median travel times (Figs. 4 and 5). Credible intervals for the effect of the Delta Cross Channel on median travel time ( $\alpha_{2}$ ) overlapped zero, with the exception of reach 7, indicating little evidence for an effect of an open gate on travel time (Fig. 4). Credible intervals for the standard deviation of the release-group random effects ( $\xi_{\mu}$ ) were well above zero, providing evidence that median travel times varied among release groups after accounting for other effects in the model. At low inflows, median travel times for tidal reaches (reaches 7 and 8) were considerably longer than other reaches (Fig. 5). Furthermore, at low flows, median travel times for reach 8 were about four times that of reach 7.

In contrast to travel time, survival was strongly related to river flow in just three of eight reaches. In the upper two reaches, which exhibit the least tidal influence, the effect of flow ( $\beta_{1}$ ) was positive (Fig. 4) but the relative change in survival was small because survival was $>0.90$ over the range of observed discharge (Fig. 6). However, we estimated strong positive effects of river flow in reaches 3-5 (Fig. 4); these reaches transition from bidirectional to unidirectional flow as river discharge increases (Fig. 3 middle panel). Although discharge affected travel time in the tidal reaches (reaches 7 and 8 ), the posterior distributions of $\beta_{1}$ were centered on zero for these reaches and credible intervals were narrow, providing strong evidence of little relationship between survival and discharge. We also found evidence that operation of the Delta Cross Channel, which removes water from the Sacramento River, was associated with lower survival in reaches of the Sacramento River downstream of the Delta Cross Channel (reaches 4 and 7). For these reaches, posterior medians of $\beta_{2}$ were negative, and $75-90 \%$ of the posterior distribution was less than zero (Fig. 4). Similar to findings with travel time, the posterior distributions for standard deviations of random effects associated with survival were positive, indicating
additional release-to-release variation in survival over and above the effects of covariates included in the model. Last, we also found a positive effect of fork length on survival ( $\beta_{3}$, median $=0.152,90 \%$ credible interval $=0.062-0.243$, Fig. A5 $\left.{ }^{1}\right)$.

Reach-specific flow-survival relationships revealed that survival increased sharply with river flow in transitional reaches but not riverine or tidal reaches (Fig. 6). Survival in riverine reaches (reaches 1 and 2 ) were high regardless of discharge, approaching 1 as flow increased. In transitional reaches, median survival at the lowest flows was about 0.75 for reach 4 (Sacramento River) and 0.5 for reach 3 (Sutter and Steamboat Slough) and reach 5 (Georgiana Slough). In these reaches, survival increased sharply with river flow, approaching 1 as river flow exceeded $1,000 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$, which coincides with the transition from bidirectional to unidirectional flow (Fig. 3 middle panel). In tidal reaches, survival was not related to discharge, but median survival in reach 7 (Sacramento River) was about twice that observed in reach 8 (interior Delta).

We found that routing probabilities (Fig. 7) followed a relationship similar to that between total discharge ( $Q$ ) and the fraction of discharge entering each route (Fig. A2), indicating that the distribution of mean daily flow among channels is a key driver of migration routing (see also Cavallo et al. 2015). As discharge increases, the probability of entering Sutter and Steamboat Slough increased by 12 percentage points from about 0.23 to an estimated upper limit ( $U$ ) of 0.35 (Table 2, Fig. 7). In contrast, as flow increases, the probability of entering Georgiana Slough (when the Delta Cross Channel gate is closed) decreased by 16 percentage points from 0.43 to an estimated lower limit ( $L$ ) of 0.27 (Table 2, Fig. 7). For these routes, routing probabilities approach upper and lower limits at an inflow of about $1,000 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$ (Fig. 7), the point at which transitional reaches switch from bidirectional to unidirectional flows (Fig. 3 middle panel). We found little variation in routing probability among release groups for Sutter
and Steamboat Sloughs and the Delta Cross Channel, but considerable variation for Georgiana Slough, particularly at low discharge (Fig. 7).

We found that operation of the Delta Cross Channel increased the proportion of fish migrating through interior Delta (reach 8) where survival is low. Routing into the Delta Cross Channel decreased as flow increased, although credible intervals were wide (Table 2, Fig. 7). We found evidence that an open Delta Cross gate reduced the probability of entering Georgiana Slough (Table 2, Fig. 7 lower right panel). However, the combined probability of entering Georgiana Slough and Delta Cross Channel, both of which lead fish to the interior Delta (reach 8), was 15 percentage points higher than the probability of entering Georgiana Slough alone when the gates are closed (Fig. 7 lower right panel).

The reach-specific survival relationships with flow dictate the composite survival of juvenile salmon migrating through the Delta via alternative migration routes. At low flows, fish migrating through the Sacramento River exhibit the highest through-Delta survival, followed by Sutter and Steamboat Slough, but as river discharge increases, survival for Sutter and Steamboat Slough approaches that of the Sacramento River, leveling off at a survival of about 0.75 (Fig. 8). Survival of fish migrating through Georgiana Slough also increases with inflow but approaches a maximum of about 0.4 . Since survival in all reaches except 7 and 8 approach 1 as discharge increases, survival in the tidal reaches imposes an upper limit on the overall through-Delta flowsurvival relationship for each route.

Since routing probabilities determine the fraction of the population experiencing a given route-specific survival, both factors contribute to the shape of the relationship between overall survival and discharge. Mean overall survival increases with discharge from about 0.32 to 0.70 and falls in between the route-specific survival relationships (Fig. 8 lower right panel).

However, at low flows, overall survival is pulled more towards the low survival of Georgiana Slough (Fig. 8 lower right panel) because the proportion of fish entering Georgiana Slough is highest at low flows (Fig. 7). By contrast, as the proportion of fish entering Georgiana Slough decreases with increasing flow, overall survival not only increases owing to the flow-survival relationships, but is weighted more towards the higher-survival migration routes owing to the flow-routing relationships.

Route-specific travel time distributions also vary considerably with river flow, and fish traveling through the interior Delta (reach 8) via Georgiana Slough or the Delta Cross Chanel exhibit longer travel times than those that migrate through the north Delta via the Sacramento River and Sutter and Steamboat Slough (Fig. 9). For example, at inflows of $235 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$, the median travel time for Georgiana Slough is 18.0 d , with some travel times as long as 40 d . By comparison, for north Delta routes, median travel times are 12.2-12.6 d, with the tail of the distribution extending to 30 d . In contrast, at inflows of $1,357 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$, expected median travel times are 2.7-3.1 d for north Delta routes compared to 6.4 d for Georgiana Slough, with a $10-\mathrm{d}$ difference between the tails of the distributions.

## Discussion

Understanding spatiotemporal variation in survival of migrating populations is critical for identifying underlying mechanisms driving survival, particularly in a highly dynamic and spatially complex environment such as the Sacramento-San Joaquin River Delta. Although variation in survival of juvenile salmon migrating through the Delta has long been linked to freshwater inflows (Kjelson and Brandes 1989; Newman and Rice 2002; Newman 2003), we lacked understanding of how spatial variation in survival gave rise to this overall relationship.

Our analysis has revealed that the overall flow-survival relationship is driven by three key reaches that transition from unidirectional flow at high inflows to tidally driven bidirectional flow at low inflows. In contrast, riverine reaches exhibited high survival at all levels of inflow and tidal reaches had lower but constant survival with respect to inflow. Thus, the flow-survival relationship captures the gradient that occurs as transitional reaches shift from tidal to riverine environments as inflow increases.

In addition to being the hub of California's water delivery system, the Delta forms a critical nexus between freshwater and ocean environments. Juvenile salmon emigrating from natal tributaries first experience a tidal environment during their migration through the Delta. Juvenile salmon are particularly vulnerable during this transition because they must modify their migration tactics to progress seaward while undergoing physiological changes in preparation for seawater entry. Although some researchers have found high survival rates in estuaries (Clark et al. 2016), others have found that migration through estuaries is associated with high mortality rates relative to riverine or early marine phases (Thorstad et al. 2012 and references therein). For example, in a study of juvenile Atlantic salmon (Salmo salar) in riverine, estuarine, and early ocean environments, Halfyard et al. (2013) found that survival was most impacted in estuarine habitats near the head of tide. Our study is consistent with these findings and highlights how river inflows can interact with tides to influence survival by shifting the location at which the hydrodynamics switch from unidirectional to bidirectional flow.

Many studies (ours included) have correlated travel time and survival to river flow because these relationships provide a direct linkage between a key management variable and the subsequent response of migrating juvenile salmon populations (Connor et al. 2003; Smith et al. 2003; Courter et al. 2016). However, it is important to recognize that river flow affects travel
time and survival through both direct and indirect mechanisms. River flow directly influences migration rates of juvenile salmon by dictating water velocity, which is a function of channel geometry (Zabel and Anderson 1997; Zabel 2002; Tiffan et al. 2009). In turn, migration rates dictate arrival timing at ocean entry, which can influence early ocean survival (Satterthwaite et al. 2014). In contrast, river flow affects survival indirectly through a number of possible mechanisms. River flow can affect the proportion of fish using alternative routes at hydroelectric projects (Coutant and Whitney 2000) and in channel network systems such as the Delta (Cavallo et al. 2015; Perry et al. 2015). If survival differs among routes, which is often the case, then river discharge affects population survival by influencing the proportion of fish using high- or low-survival routes (Perry et al. 2013, 2016). In addition, river flow is often correlated with other environmental variables that influence survival such as turbidity and water temperature (Baker et al. 1995; Connor et al. 2003; Smith et al. 2003), which in turn may influence predation rates (Vogel and Beauchamp 1999; Ferrari et al. 2013).

Reducing travel time and exposure to predators is a key mechanism by which river flow has been hypothesized to affect survival of migrating juvenile salmonids, but establishing this linkage has proven elusive. Although travel time has been consistently linked with discharge via water velocity, some studies in the Snake and Columbia Rivers found no significant relationship between travel time and survival, but a significant relation between migration distance and survival (Bickford and Skalski 2000; Smith et al. 2002). To explain these counterintuitive findings, Anderson et al. (2005) developed a predator-prey model that expressed survival as a function of both travel time and travel distance. Their analysis revealed that the dependence of survival on travel time is dictated by the nature of predator-prey interactions. When prey migrate in a directed fashion through a field of stationary predators, survival is independent of travel time
and depends only on travel distance. In contrast, survival depends on travel time when random movement dominates directed migration or when predators adopt prey searching tactics. These findings illustrate how the relationship between flow and survival is context dependent, arising from a combination of mechanisms that may directly affect survival (e.g., temperature) or indirectly affect survival by modifying predator encounter rates (temperature, turbidity, predator and prey behavior). Although our analysis here focused on estimating the association between discharge and survival, our modeling framework allows exploration of alternative model structures for linking flow to survival via travel time. For example, the XT model can be incorporated into our analytical framework and compared against the current model structure to assess the strength of evidence for the dependence of survival on travel time.

Predation by a host of non-native piscivorous fishes is thought to be the primary proximate cause of juvenile salmon mortality in the Delta (Cavallo et al. 2013; Grossman 2016; Sabal et al. 2016). Variation in survival among reaches observed in our study is consistent with expectations based on predator-prey models. Juvenile salmon migrate downstream through riverine reaches in a directed fashion and survival was high regardless of inflows and variation in travel time. We observed a similar pattern at high flows when transitional reaches exhibit unidirectional flows similar to riverine reaches. As inflow declines and tidal influence moves upstream into transitional reaches, not only does travel time increase but travel distance increases because juvenile salmon may be advected upstream on flood tides (Moser et al. 1991). Simultaneously increasing both travel time and cumulative travel distance will act to increase predator encounter rates. Thus, the flow-survival relationship that we observed in transitional reaches likely arises from the transition from directed downstream movement at high flows to less directed, bidirectional movement during low flows. In tidal reaches, our a priori expectation
was that neither travel time nor survival would be related to inflow because the magnitude of tidal flows (on the order of $\pm 3,500 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$ ) swamps the signal of net inflow (Fig. 3 bottom panel). Although we observed no relation between survival and inflow in tidal reaches, we were surprised to find a strong effect of inflow on travel time, suggesting that survival may be decoupled from travel time in tidal reaches.

We included only inflow, DCC gate position, and fork length as covariates on reachspecific survival, but numerous other factors drive variation in survival. By casting our markrecapture model in a hierarchical Bayesian framework and including a random effect on release group, we were able to quantify the magnitude of this variation but not its source. Seasonal and annual variation in reach-specific predator densities, environmental drivers (e.g., water temperature, turbidity), and spring-neap tidal cycles will act to modulate how travel time and survival respond to changes in inflow, thereby propagating variation among cohorts of juvenile salmon that experience a common set of flow conditions. Specifically, variation in survival among release groups was highest in tidal reaches and in transition reaches during low inflows, further suggesting that tidal cycles play an important role in driving variation in survival. For example, at a given inflow, cohorts migrating through the Delta during neap tides will experience lower-magnitude flood tides and first encounter bidirectional flows further downstream relative to cohorts migrating during spring tides. The high release-to-release variation in our study provides opportunity for future work to quantify how factors other than inflow influence survival in the Delta.

Our Bayesian mark-recapture model makes two important advances in the development of statistical mark-recapture models used to estimate survival of migrating juvenile salmonids. First, modeling individual covariates that vary through time is challenging owing to missing
covariate values for undetected individuals. Consequently, most approaches have ignored within- and among-individual variation by averaging covariates over individuals within release groups (Conner et al. 2003; Smith et al. 2003) or averaging covariates over space and time for each individual (Stitch et al. 2015). In contrast, our model allows for individual covariates that vary through space and time by incorporating an auxiliary model for travel times to impute missing travel time, reach entry times, and covariates at the time of reach entry. Given the considerable capital expense associated with conducting telemetry studies, our framework allows the maximum amount of information to be extracted from telemetry data sets that typically have small sample sizes. Second, our joint travel time and mark-recapture model explicitly considers both migration and demographic processes under a single analytical framework. Thus, our modeling framework opens the door to a number of useful extensions such as modeling survival directly as a function of an individual's travel time (Muthukumarana et al. 2008) or by using an event-time framework (e.g., Sparling et al. 2006; Zabel et al. 2014) where survival can be modeled as a function of temporal variation in covariates during an individual's residence time within a reach.

Our analysis provides insight into how water management decisions that influence inflow and water routing are likely to affect travel time, routing, and survival of migrating juvenile salmonids. First, survival decreases sharply and routing into the interior Delta (where survival is low) increases sharply as Delta inflows decline below approximately $1,000 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$, the point at which transitional reaches shift from bidirectional to unidirectional flow (Figs 7 and 8). In contrast, at inflows greater than $1,000 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$, survival is maximized and changes relatively little with flow while routing into the interior Delta via Georgiana Slough is minimized and insensitive to inflow. These findings indicate that water management actions that reduce inflows to the Delta
will have relatively little effect on survival at high flows, but potentially considerable negative effects at low flows. Furthermore, operation of the Delta Cross Channel not only increases the fraction of the population that enters the interior Delta where survival is low (Fig. 7), but is associated with lower survival for the Sacramento River (Fig. 4). These compounding effects of opening the Delta Cross Channel act to further reduce overall survival relative to inflows alone (Fig. A6 ${ }^{1}$ ). Our findings illustrate how tradeoffs between juvenile salmon survival and water management for human use vary with the amount of flow entering the Delta. Thus, our modeling framework can be used as a management tool to explore the consequences of such tradeoffs and to quantitatively assess the effect of alternative management scenarios on travel time, routing, and survival.

Water flow has been dubbed the "master" variable in the Delta because of its economic importance and its pervasive effect on all components of this complex and dynamic aquatic ecosystem (Mount et al. 2012; Lund et al. 2015). Indeed, our work is beginning to shed light on the multiple ways in which river flows differentially affect survival in different reaches of the Delta and interact with water and fish routing to affect overall survival. In turn, these insights will aid managers in devising strategies to balance consumptive water use with management actions that aim to recover threatened and endangered salmon populations in the Sacramento River.

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## 889

890
Table 1. Description of release groups and data sources.

| Release group | Source | Year | Release <br> dates | Number <br> released | Number analyzed | Release sites (rkm) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | Perry et al. $(2010,2013)$ | 2006 | Dec 5-6 | 64 | 64 | 172 |
| 2 | Perry et al. $(2010,2013)$ | 2007 | Jan 17-18 | 80 | 80 | 172 |
| 3 | Michel et al. (2015) |  | $\begin{aligned} & \text { Jan 15-Feb } \\ & 2 \end{aligned}$ | 200 | 11 | 517 |
| 4 | Perry et al. (2013) |  | Dec 4-7 | 208 | 208 | 115, 172 |
| 5 | Michel et al. (2015) |  | Dec 7 | 150 | 60 | 345, 398, 500 |
| 6 | Perry et al. (2013) | 2008 | Jan 15-18 | 211 | 211 | 115, 172 |
| 7 | Michel et al. (2015) |  | Jan 17 | 154 | 65 | 345, 398, 500 |
| 8 | Perry et al. (2013) |  | Nov 30- <br> Dec 6 | 292 | 292 | 115, 172 |
| 9 | Michel et al. (2015) |  | Dec 13 | 149 | 82 | 345, 398, 500 |
| 10 | Michel et al. | 2009 | Jan 11 | 151 | 63 | 345, 398, 500 |

(2015)

| 11 | Perry et al. | Jan 13-19 | 292 | 292 | 115,172 |
| :--- | :--- | :--- | :--- | :--- | :--- |
|  | $(2013)$ |  |  |  |  |
| 12 | Perry et al. | Dec 2-5 | 239 | 239 | 115,191 |
|  | $(2012)$ |  |  |  |  |

13
Michel et al.
Dec 15153
$63 \quad 345,398,500$
(2015)

14
Perry et al.
Dec 16-19 240
240
115, 191
(2012)

15
Michel et al. 2010 Jan 6
153
42
345, 398, 500 (2015)

16
Michel et al.
Dec 17
120
79
500
(2015)

17
Michel et al.
2011 Jan 5
120
79
500
(2015)
$\begin{array}{lll}\text { All groups } & 2,976 & 2,170\end{array}$

Note: Release sites are indicated by river kilometer (rkm) measured from the distance to the Pacific Ocean. For fish release upstream of the Delta ( $>$ rkm 208), the number analyzed indicates fish that were included in the analysis based on detections at telemetry stations near the entrance to the Delta at rkm 189 or rkm 226.

Table 2. Posterior medians ( $90 \%$ credible intervals) for routing probabilities expressed as a function of time-varying individual covariates. Parameter values without associated credible intervals were set to the given value.

| Parameter | Sutter and Steamboat <br> Slough $\left(\Psi_{\mathrm{AB} 3}\right)$ | Delta Cross <br> Channel $\left(\Psi_{\mathrm{AC} 4}\right)$ | Georgiana <br> Slough $\left(\Psi_{\mathrm{AD} 4 \mid \mathrm{C}^{\prime}}\right)$ |
| :--- | :--- | :--- | :--- |
| $U$ | $0.35(0.31-0.43)$ | 1 | 1 |
| $L$ | 0 | 0 | $0.27(0.20-0.32)$ |
| $\gamma_{0}$ | $1.89(0.91-3.30)$ | $-1.49(-2.40--0.67)$ | $-2.95(-4.57--1.83)$ |
| $\gamma_{1}$ | $2.17(1.10-4.15)$ | $-1.25(-3.47-0.90)$ | $-6.53(-5.46--1.24)$ |
| $\gamma_{2}$ | 0 | 0 | $-0.55(-2.76-0.33)$ |
| $\xi_{\Psi}$ | $0.19(0.04-0.50)$ | $0.31(0.06-0.87)$ | $0.89(0.46-1.58)$ |

Note: $U=$ upper limit of logistic function, $L=$ lower limit of logistic function, $\gamma_{0}=$ intercept, $\gamma_{1}$ $=$ slope for effect of discharge on routing, $\gamma_{2}=$ slope for effect of open Delta Cross Channel gates on routing, $\xi_{\Psi}=$ standard deviation of the release group random effect.

## Figure Captions

Fig. 1. Map of the Sacramento-San Joaquin River Delta showing the location of acoustic telemetry receiving stations (filled black circles) used to detect acoustic tagged juvenile salmon as they migrated through the Delta. Telemetry stations are labeled by migration route (A-D) and sampling occasion (1-7; see Fig. 2). These telemetry stations divide the Delta into eight discrete reaches (shown by numbered shaded regions), with an additional reach upstream of telemetry station $A_{2}$ (Reach 0 ) used as acclimation reach to allow fish to recover from post-release handling. The location of water pumping stations in the southern interior Delta is indicated by the pink diamonds.

Fig. 2. Schematic of the multistate mark-recapture model with parameters indexed by state (migration route) and sampling occasion. Parameters include reach-specific survival probabilities $(S)$, site-specific detection probabilities $(P)$, routing probabilities $(\Psi)$, and $\lambda$, the joint probability of surviving and being detected at telemetry stations downstream of site $\mathrm{A}_{6}$. Release locations are indicated by the $n$th release in route $r$ at occasion $j: n_{\mathrm{Al}}$ (at Sacramento or upstream - see Table 1) and $n_{\text {D } 4}$ in Georgiana Slough.

Fig. 3. Daily inflow into the Sacramento-San Joaquin River Delta (top panel) and tidal influence on discharge at three locations in the Sacramento River during 2010 (middle and bottom panels). The top panel shows mean daily discharge of the Sacramento River at Freeport ( $\mathrm{A}_{2}$ in Fig. 1). In the two lower panels, lines show mean daily discharge and the shaded regions encompass the daily minimum and maximum discharge with values less than zero indicating reverse flows
caused by tidal forcing. The middle panel shows the Sacramento River at Freeport (black line, gray shading) and the Sacramento River just downstream of Georgiana Slough (pink link and shading; $\mathrm{A}_{4}$ in Fig. 1). The bottom panel shows the Sacramento River at Rio Vista ( $\mathrm{A}_{5}$ in Fig. 1).

Fig. 4. Summary of posterior distributions of parameters estimating the effects of river flow and Delta Cross Channel (DCC) gate position on travel time and survival. Points show the median of the posterior distribution, heavy lines show the $25^{\text {th }}$ to $75^{\text {th }}$ percentile, and thin lines show the $5^{\text {th }}$ to $95^{\text {th }}$ percentile. Green bars are density strips with darker regions illustrating higher posterior density. Parameter definitions as follows: $\alpha_{1}=$ slope for effect of discharge on mean of logtravel time, $\alpha_{2}=$ slope for effect of an open DCC gate on mean of log-travel time, $\xi_{\mu}=$ standard deviation of release group random effect on $\mu, \sigma=$ variance parameter of the lognormal travel time distribution, $\beta_{1}=$ slope for effect of discharge on survival, $\beta_{2}=$ slope for effect of an open DCC gate on survival, $\xi_{S}=$ standard deviation of release group random effect on survival.

Fig. 5. Reach-specific relationships between median travel time and inflow to the Delta as measured at the Sacramento River at Freeport (shown for closed Delta Cross Channel gates and plotted at the mean fork length). The heavy magenta line shows the mean relationship and the dotted lines show the random effects estimates for each release group based on medians of the joint posterior distribution. The dark gray region shows $95 \%$ credible intervals about the mean relationship. The light gray region shows the $95 \%$ confidence interval among release groups.

Fig. 6. Reach-specific relationships between survival and inflow to the Delta as measured at the Sacramento River at Freeport (shown for closed Delta Cross Channel gates and plotted at the
mean fork length). The heavy magenta line shows the mean relationship and dotted lines show the random effects estimates for each release group based on medians of the joint posterior distribution. The dark gray region shows $95 \%$ credible intervals about the mean relationship. The light gray region shows the $95 \%$ confidence interval among release groups.

Fig. 7. Relationships between routing probability and inflow to the Delta as measured at the Sacramento River at Freeport ( $\mathrm{A}_{2}$ in Fig. 1). The lower right panel shows the effect of Delta Cross Channel (DCC) gate position on routing probabilities at the junction of the Sacramento River, Delta Cross Channel, and Georgiana Slough ( $\mathrm{A}_{4}, \mathrm{C}_{4}$, and $\mathrm{D}_{4}$ in Fig. 1), plotted at the posterior median of the parameters. Other panels show mean routing relationships (heavy magenta line), random effects estimates for each release group (dotted lines), $95 \%$ credible interval about the mean relationship (dark gray region), and $95 \%$ confidence interval among release groups (light gray region).

Fig. 8. Route-specific survival through the Sacramento-San Joaquin River Delta between Freeport ( $\mathrm{A}_{2}$ in Fig. 1) and Chipps Island ( $\mathrm{A}_{6}$ in Fig. 1). Route-specific survival based on posterior median parameter values was calculated as the product of reach-specific survival for reaches that trace each unique migration route through the Delta (shown for closed Delta Cross Channel gates). The first three panels show the mean relationship for each route with thin gray lines showing the random effects estimates for each release group. The last panel shows overall survival through Delta for all routes (with random effects estimates as thin gray lines) along with route-specific survival relationships. Overall survival was calculated as the average of routespecific survival weighted by routing probabilities (see Eq. 1).

Fig. 9. Route-specific travel time distributions between Freeport ( $\mathrm{A}_{2}$ in Fig. 1) and Chipps Island ( $\mathrm{A}_{6}$ in Fig. 1) at the $5^{\text {th }}$ (top panel) and $95^{\text {th }}$ (bottom panel) percentiles of discharge based on the historical flow record ( $235 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$ and $1357 \mathrm{~m}^{3} \cdot \mathrm{~s}^{-1}$, respectively). Arrows show the median travel times for each route. Travel time distributions were based on posterior medians of parameters for reach-specific travel time distributions assuming closed Delta Cross Channel gates.


Fig. 1. Map of the Sacramento-San Joaquin River Delta showing the location of acoustic telemetry receiving stations (filled black circles) used to detect acoustic tagged juvenile salmon as they migrated through the Delta. Telemetry stations are labeled by migration route (A-D) and sampling occasion (1-7; see Fig. 2). These telemetry stations divide the Delta into eight discrete reaches (shown by numbered shaded regions), with an additional reach upstream of telemetry station $\mathrm{A}_{2}$ (Reach 0) used as acclimation reach to allow fish to recover from post-release handling. The location of water pumping stations in the southern interior Delta is indicated by the pink diamonds.

## Migration route (State)



Fig. 2. Schematic of the multistate mark-recapture model with parameters indexed by state (migration route) and sampling occasion. Parameters include reach-specific survival probabilities $(S)$, site-specific detection probabilities $(P)$, routing probabilities $(\Psi)$, and $\lambda$, the joint probability of surviving and being detected at telemetry stations downstream of site $\mathrm{A}_{6}$. Release locations are indicated by the $n$th release in route $r$ at occasion $j: n_{\mathrm{Al}}$ (at Sacramento or upstream - see Table 1) and $n_{\text {D } 4}$ in Georgiana Slough.


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Fig. 6. Reach-specific relationships between survival and inflow to the Delta as measured at the Sacramento River at Freeport ( $\mathrm{A}_{2}$ in Fig. 1; shown for closed Delta Cross Channel gates and plotted at the mean fork length). The heavy magenta line shows the mean relationship and dotted lines show the random effects estimates for each release group based on medians of the joint posterior distribution. The dark gray region shows $95 \%$ credible intervals about the mean relationship. The light gray region shows the $95 \%$ confidence interval among release groups.


Fig. 7. Relationships between routing probability and inflow to the Delta as measured at the Sacramento River at Freeport ( $\mathrm{A}_{2}$ in Fig. 1). The lower right panel shows the effect of Delta Cross Channel (DCC) gate position on routing probabilities at the junction of the Sacramento River, Delta Cross Channel, and Georgiana Slough ( $\mathrm{A}_{4}, \mathrm{C}_{4}$, and $\mathrm{D}_{4}$ in Fig. 1), plotted at the posterior median of the parameters. Other panels show mean routing relationships (heavy magenta line), random effects estimates for each release group (dotted lines), $95 \%$ credible interval about the mean relationship (dark gray region), and $95 \%$ confidence interval among release groups (light gray region).


Fig. 8. Route-specific survival through the Sacramento-San Joaquin River Delta between Freeport ( $\mathrm{A}_{2}$ in Fig. 1) and Chipps Island ( $\mathrm{A}_{6}$ in Fig. 1). Route-specific survival based on posterior median parameter values was calculated as the product of reach-specific survival for reaches that trace each unique migration route through the Delta (shown for closed Delta Cross Channel gates). The first three panels show the mean relationship for each route with thin gray lines showing the random effects estimates for each release group. The last panel shows overall survival through Delta for all routes (with random effects estimates as thin gray lines) along with route-specific survival relationships. Overall survival was calculated as the average of routespecific survival weighted by routing probabilities (see Eq. 1).


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## Supplement A.

Perry et al. Flow-mediated effects on travel time, survival, and routing of juvenile Chinook salmon in a spatially complex, tidally forced river delta


Fig. A1. Scatter plots showing the relationship between daily discharge of the Sacraemento River at Freeport and discharge entering other channels downstream of Freeport. The grey reference line shows where discharge in a given channel is equal to discharge at Freeport. Subscripted letters in parentheses refer to the location of USGS gauging stations near telemetry stations shown in Fig. 1.


Fig. A2. Fraction of dishcharge entering Sutter and Steamboat Slough (top) and Delta Cross
Channel and Georgiana Slough (bottom) calculated as the proportion of total inflow relative to Sacramento River discharge at Freeport.


Fig. A3. Posterior distributions for $\theta_{1}$, the coefficient measuring the effect of flow on detection probability at each telemetry station. Points show the median of the posterior distribution, heavy lines show the $25^{\text {th }}$ to $75^{\text {th }}$ percentile, and thin lines show the $5^{\text {th }}$ to $95^{\text {th }}$ percentile. Green bars are density strips with darker regions illustrating higher posterior density. See Fig. 1 for locations of telemetry stations.


Fig. A4. Year-specific effects of discharge on detection probability of each telemetry station based on posterior medians of the parameters. Lines are plotted over the range of discharge observed in each year. See Fig. 1 for locations of telemetry stations.


Fig. A5. Reach-specific relationships between survival and fork length (shown for closed Delta Cross Channel gates and plotted at the mean discharge). The heavy magenta line shows the mean relationship and dotted lines show the random effects estimates for each release group based on medians of the joint posterior distribution. The dark gray region shows $95 \%$ credible intervals about the mean relationship. The light gray region shows the $95 \%$ confidence interval among release groups.


Fig. A6. Route-specific flow-survival relationships for Delta Cross Channel (DCC) open and closed. Relationships are plotted at the posterior medians of parameter distributions.


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[^1]:    ${ }^{1}$ With the exception of 2016, the "wet year" phrase is based on resource agency water-year type or classification (CDEC water year classification, DWR, 2017b, see "Notes"). The 2016 water year for the Sacramento and San Joaquin valleys was classified as below normal and dry, respectively. However, 2016 is qualitatively described in this document as a wet year because above-average precipitation was recorded in most of the SFE watershed and spawning and migratory conditions typically associated with wet years occurred (e.g., elevated winter or spring flows, extended periods of bypass inundation).

[^2]:    ${ }^{2}$ This refers to juveniles from brood years 2006 and 2011. At the time of writing, juvenile samples from brood year 2016 have not been analyzed.

[^3]:    ${ }^{\text {a }}$ Some ocean fishery recoveries are received from out-of-state sources
    ${ }^{\mathrm{b}}$ First year of consistent Coded-wire tagging
    ${ }^{c}$ CDFW refers to the California Department of Fish and Wildlife, and USFWS refers to the United States Fish and Wildlife Service.

[^4]:    ${ }^{1}$ Supplementary data are available with the article through the journal Web site at $\mathrm{http}: / /$ nrcresearchpress.com/doi/supp/...

